# Wages, Job Mobility, and Firm Performance: an analysis using matched employer-employee data 

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## Outline of the Talk

- Motivation: many theories predict a relation between wages, turnover, seniority, and firm performance
- Can we find these relations in the data?
- Theoretical models
- Construction of the data set
- Statistical model
- Results


## Motivation

- Job mobility?
- Firms: impact of hiring, promotion, and separations (worker selection, more generally) on performance
- Theoretical background: turnover and matching, accumulation of human capital, efficiency wages, selection, career concerns.
- Data: a large longitudinal linked employer-employee data-set (1976-1996)
- General project: joint examination of wage setting, firm-specific mobility policies, and firm performance in France and the US


## Theoretical Models

- General v. Firm-specific human capital
- A priori low v. high returns to human capital
- Rent-sharing
- Efficiency wages and costly turnover
- Performance and turnover negatively related.
- Matching or screening
- Quality matters, selection matters
- Career concerns, tournaments
- Competition between workers within firms. Competition between firms
- Incentives
- Pay and career tied to individual performance


## Multiple Equilibria or Models

- Most models, either economic or econometric, assume homogeneity of effects between firms
- We posit extreme heterogeneity
- Across firms
- Unobserved to us
- Each firm follows one model across time


## A Model of Workers' Career

- Worker $i$, employed in firm $j$ for $s$ years has productivity:

$$
p_{i j}^{s t}=X_{i j}^{(1) t} \beta+X_{i j}^{(2) s t} \beta_{j}+\phi_{i}+e_{i j}^{0}+\varepsilon_{i j}^{s t}
$$

$-X^{(1)}$ corresponds to variables valued on the labor market (experience, ...)
$-X^{(2)}$ corresponds to variables valued in the firm

- Individual effect, match effect


## Wages

- Rent-sharing ?
- Opportunity wage: $X_{i j}^{(1) t} \beta+\phi_{i}$
- Rent to share: $X_{i j}^{(2) s t} \beta_{j}+e_{i j}+\varepsilon_{i j}^{s t}$
- Hence :

$$
w_{i j}^{s t}=X_{i j}^{(1) t} \beta+\phi_{i}+\gamma_{j}\left(X_{i j}^{(2) s t} \beta_{j}+e_{i j}+\varepsilon_{i j}^{s t}\right)
$$

Other interpretations : discrimination?
From wages alone, one can identify only $\beta_{j} \gamma_{j}$
The firm seeks to maximize the intertemporal sum of rents deriving from employment

## Sequence Within a Job

- Worker is hired : her fixed effect is observed
- Compensation at entry :
- Observable Characteristics valued on the market
- Firm-specific effect (incentives or general compensation policy)
- Individual effect or match effect
- Employment duration (firm and worker decision)
- Depends on firm's policy through the parameters of the law of duration
- Workers can leave if prospects are unappealing
- Wage changes :
- Productivity changes within the firm
- Firm-specific compensation policy


## Employment Duration

- Realized duration is tied to the firm and the workers behaviors
- Firm looks at worker's productivity and potential abilities:
- Expected correlation attendue with the individual and the match specific effect (a priori positive).
- Internal market in the firm?
- Position in the age pyramid of the firm at entry
- Potential attachment to the job :
- Number of previous jobs
- Duration in the previous job


## Wage Changes

- Structure of wage changes :
- Induced by the presence in the firm (duration, endogenous selection, firm-specific coefficients)
- Related to changes in the rent conditional on $X^{(1)}$
- Introduction of individual fixed effects:
- Correlation between levels of pay and changes in pay
- Introduction of the residual from the entry equation :
- Match quality and wage changes


## Statistical Model

- Initial wage
- Employment duration (potentially censored, Firm-specific)
- Employment duration depends on initial wage
- Wage changes (Firm-specific)
- Endogenous duration induces selection effects in the wage changes
- Dynamic effects of the initial wage on future wage changes in the firm


## Initial Wage Equation

$\left(\ln w_{i J(i, t)} \mid T_{i J(i, t)}=0\right)=X_{i J(i, t)}^{(1)} \beta+\theta_{i}+\psi_{\mathrm{J}(i, t)}+\varepsilon_{i J(i, t)}$

- where $T$ is seniority
- where $\theta_{l}$ is the individual effect in the entry equation
- où $\psi_{\mathrm{j}}$ is the firm effect in the entry equation
- Other variables: experience, year indicators, region, full-time indicator (all in full interaction with sex)
- Earnings :log of annualized earnings


## Job Duration Equation (firm-specific)

## $\ln T_{i j}=Z_{i j} \alpha_{j}+\eta_{i j}$

- where $\ln T$ is total duration of $i$ at employer $j=$ $J(i, t)$.
- other variables:
- experience, sex, full-time indicator
- position in the firm-specific age distribution (at entry, quartiles),
- $\theta$ et $\varepsilon$,
- Number of previous jobs, duration of the previous job
- Duration depends on initial wage


## Wage-Change (firm-specific)

$$
\ln w_{i t}-\ln w_{i t-1}=\Delta X_{i t}^{(1)} \beta+\Delta X_{i t}^{(2)} \beta_{J(i, t)}+v_{i t J(i, t)}
$$

- $X^{(1)}$ corresponds to the variables also present in the initial wage equation (experience in particular)
- $X^{(2)}$ corresponds to the variables valued in the firm : seniority with returns depending on
- $\theta$ and $\varepsilon$ interacted with seniority indicators
- Sex, education
- Full-time to part-time change in status
- Time Dummies
- Duration (selection effect) and initial wage are endogenous in the wage change equation


## Likelihood

- Contrib. of job (i,j, $t_{i j}$ ) to the likelihood:

$$
\begin{aligned}
L_{i j}^{t} & =f\left(w_{i j}^{0}, \Delta w_{i j}^{1}, \ldots, \Delta w_{i j}^{T_{i j}}, T_{i j} / I_{t-1}\right) \\
& =f^{0}\left(w_{i j}^{0} / I_{t-1}, j\right) * f_{j}^{T}\left(T_{i j}>0 / w_{i j}^{0}, I_{t-1}\right) * f_{j}^{w}\left(\Delta w_{i j}^{1} / T_{i j}>0, w_{i j}^{0}, I_{t-1}\right) * \ldots \\
& * f_{j}^{T}\left(T_{i j}=t_{i j} / T_{i j}>t_{i j}-2, w_{i j}^{0}, I_{t-1}\right) * f_{j}^{w}\left(\Delta w_{i j}^{t_{i j}} / T_{i j}>t_{i j}-1, w_{i j}^{0}, I_{t-1}\right)
\end{aligned}
$$

- $I_{t-1}$ contains all past information
- $J$ corresponds to the firm (model expressed firm by firm)
- Contribution of individual $i$ to the likelihood:

$$
L_{i}=\prod_{k=1}^{N_{i}} L_{i j(k)}
$$

## Likelihood (continued)

- Rewriting of the likelihood:

$$
L_{i j}^{t}=f^{w}\left(w_{i j}^{0} / I_{t-1}\right) * f_{j}^{T}\left(T_{i j}=t_{i j} / w_{i j}^{0}, I_{t-1}\right) * \prod_{1}^{t_{i j}} f_{j}^{w}\left(\Delta w_{i j}^{k} / T_{i j}>k-1, w_{i j}^{0}, I_{t-1}\right)
$$

- which gives:

$$
\begin{aligned}
& f^{w}\left(w_{i j}^{0}\right)=\varphi\left(\frac{\ln \left(w_{i j}^{0}\right)-X_{i t} \beta-\theta_{i}-\psi_{j}}{\sigma_{w}}\right) \\
& f_{j}^{T}\left(T_{i j} / \varepsilon_{i j}, Z_{i j}\right)=\varphi\left(\frac{\ln \left(T_{i j}\right)-Z_{i j} \alpha_{j}-\rho_{j}^{1} \varepsilon_{i j}}{\sigma_{j}^{T}}\right)^{d_{i j}=0} \bar{\Phi}\left(\frac{\ln \left(T_{i j}\right)-Z_{i j} \alpha_{j}-\rho_{j}^{1} \varepsilon_{i j}}{\sigma_{j}^{T}}\right)^{d_{i j}=1} \\
& f_{j}^{\Delta w}\left(\Delta w_{i j}^{1}, \ldots, \Delta w_{i j}^{T_{i j}} / T_{i j}, \varepsilon_{i j}, X^{(1)}, X^{(2)}\right)=\prod_{s=1}^{E\left(T_{i j}\right)} f\left(\Delta w_{i j}^{s}-\Delta X^{(1)} \beta-\Delta X^{(2)} \beta_{j} / s<T_{i j}, \Delta X^{(1)}, \Delta X^{(2)}, \varepsilon_{i j}\right)
\end{aligned}
$$

## Parameter Estimation

- Maximum likelihood?
- Sequential Estimation procedure
- Entry wage equation : Identification of $\beta$ and of individual and firm effects, OLS
- Job duration equation : Duration model with censoring, firm by firm, includes the individual residual from the entry wage equation
- Wage change equation : OLS firm by firm
- Wages corrected for $X^{(1)}$
- Introduction of the entry wage residual
- Introduction of the Mills ratio induced by the duration equation


## Construction of individual data

- Years 1976-1996 except 1981, 1983 and 1990
- 1 observation corresponds to a NNI-SIREN-YEAR (person-firmyear)
- Construction des job spells

- Imputation for the NNI wrongly coded and the missing years



## Job Seniority

- Observed for all spells that start after 1976
- Statistical Imputation for all spells starting in 1976 using ESS 1978



## Final Data Structure

- 13,770,082 observations (of which 4,886,669 are for entry jobs)
- 1,682,080 individuals
- 515,557 firms (with 6498 "firms" for which the full model was estimated including 189 "3-digit industries*Size" agregate firms)
- About 6,000 firms in the performance equations


## Construction of the FirmLevel Data

```
Individual data-set
observation:
NNI-SIREN-YEAR
Firm-by-Firm Estimation
Firm-Specific Parameters from each
    equation (firms with 200 obs. +)
```

Firm data-set (BIC-BRN) observation: SIREN-YEAR
Performance
Aggregation over years for each sub-period

Merging

Final data-set :
observation: SIREN or SIREN-SUBPERIOD

| Average Coefficients for the Starting Wage Equation Selected Variables |  |  |
| :---: | :---: | :---: |
| Variable | Men | Women |
| Labor force experience | 0,0973 | 0,0742 |
| Experience^2/100 | -0,5447 | -0,4269 |
| Experience^3/1000 | 0,1397 | 0,1192 |
| Experience^4/10000 | -0,0138 | -0,0122 |
| Paris | 0,0870 | 0,1024 |
| Full time job | 0,8221 | 0,7834 |
| Number | 2920340 | 1966329 |
| Note: Estimated using exact least squares (congugate gradient algorithm). The equation also includes year, person, and firm effects. |  |  |

## All Entrants, Descriptive Statistics

| Variable | Means | StDev |
| :--- | :---: | :---: |
|  |  |  |
| Ile de France | 0.3134 |  |
| Job duration | 1.0804 | 3.5061 |
| Age at the end of schooling | 18.3592 | 1.7543 |
| male | 0.5976 |  |
| Year | 1985.4500 | 6.6716 |
| Log earnings | 3.3135 | 1.0661 |
| Age at the end of schooling | 30.9861 | 10.7513 |
| Number of previous jobs | 3.5570 | 16.3801 |
| Duration of the previous job | 1.8492 | 4.2156 |
| Mobility | 0.5259 | 0.4993 |
| Low-education | 0.5507 |  |
| Technical education | 0.2865 |  |
| High-education | 0.1628 |  |


| Standardized variables |  |  |
| :--- | :--- | :--- |
| Log wage | 0.0000 | 1.0657 |
| Xbeta | 0.0000 | 0.4522 |
| Person effect | 0.0000 | 0.5680 |
| Firm effect | 0.0000 | 0.4910 |
| Residual | 0.0000 | 0.6158 |


| Correlation of Components of the Starting Wage Rate |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| $\mathbf{N = 4 , 8 8 6 , 6 6 9}$In w | $x \beta$ | $\theta$ | $\psi$ | Residual |  |
| Log annualized net salary (In w) | 1,000 | 0,473 | 0,486 | 0,449 | 0,578 |
| Predicted by time-varying effects (x $\times$ ) | 0,473 | 1,000 | 0,002 | 0,103 | 0,000 |
| Person effect ( $\theta$ ) | 0,486 | 0,002 | 1,000 | $-0,105$ | 0,000 |
| Firm effect $(\psi)$ | 0,449 | 0,103 | $-0,105$ | 1,000 | 0,000 |
| Residual | 0,578 | 0,000 | 0,000 | 0,000 | 1,000 |


|  | Distribution of job-duration equation coefficients |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Mean | StDev | 1pctle | 5pctle | 1st quartile | Median | 3rd quartile | 95 pctle | 99pctle |
| Constant | -0.19 | 1.00 | -2.31 | -1.71 | -0.79 | -0.14 | 0.27 | 1.55 | 2.29 |
| Year<=1980 | 0.30 | 0.76 | -1.38 | -0.45 | 0.10 | 0.22 | 0.48 | 1.38 | 2.73 |
| 1980<Year<=1989 | -0.24 | 0.63 | -1.53 | -0.69 | -0.41 | -0.29 | -0.11 | 0.59 | 1.24 |
| Position in age distribution at entry: |  |  |  |  |  |  |  |  |  |
| Age<=25centile | 1.48 | 1.04 | -0.92 | -0.13 | 0.91 | 1.44 | 2.19 | 3.20 | 3.86 |
| 25centile<Age<=50centile | 1.15 | 0.82 | -0.76 | -0.11 | 0.70 | 1.13 | 1.64 | 2.36 | 3.01 |
| 50centile<Age<=75centile | 0.80 | 0.56 | -0.54 | -0.08 | 0.57 | 0.83 | 1.15 | 1.55 | 2.07 |
| Full-time | 0.43 | 0.67 | -1.08 | -0.33 | 0.16 | 0.43 | 0.64 | 1.32 | 2.19 |
| Male | -0.17 | 1.15 | -0.81 | -0.47 | -0.31 | -0.19 | -0.09 | 0.30 | 0.70 |
| Experience | 0.12 | 0.08 | -0.06 | -0.01 | 0.07 | 0.12 | 0.16 | 0.24 | 0.35 |
| Experience ${ }^{12}$ | -0.15 | 0.16 | -0.66 | -0.41 | -0.20 | -0.14 | -0.08 | 0.03 | 0.17 |
| Low-Education | -0.46 | 0.83 | -2.28 | -1.10 | -0.63 | -0.46 | -0.33 | 0.27 | 1.32 |
| High-Education | -0.26 | 1.06 | -2.64 | -1.30 | -0.55 | -0.21 | -0.01 | 0.62 | 1.90 |
| Duration of previous job | 0.28 | 0.28 | 0.04 | 0.14 | 0.23 | 0.26 | 0.30 | 0.45 | 0.81 |
| Number of previous jobs | -0.05 | 0.12 | -0.42 | -0.15 | -0.05 | -0.04 | -0.03 | -0.01 | 0.01 |
| Individual effect | -0.06 | 0.31 | -0.86 | -0.47 | -0.19 | -0.07 | 0.12 | 0.33 | 0.82 |
| Residual from the entry wage equation | -0.22 | 0.32 | -0.83 | -0.59 | -0.35 | -0.23 | -0.13 | 0.23 | 0.83 |


| Distribution of the Students of the Duration Equation Coefficients |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Mean | StDev | 1 pctle | 5pctle | 1st quartile | Median | 3rd quartile | 95 pctle | 99pctle |
| Constant | -42.56 | 127.89 | -314.65 | -314.65 | -63.27 | -12.17 | 9.00 | 42.95 | 118.88 |
| Year<=1980 | 8.30 | 12.77 | -23.50 | -7.67 | 1.27 | 8.02 | 13.23 | 24.40 | 45.92 |
| 1980<Year<=1989 | -12.64 | 40.62 | -47.57 | -44.82 | -24.07 | -9.56 | -1.28 | 4.37 | 58.10 |
| Position in age distribution at entry: |  |  |  |  |  |  |  |  |  |
| Age<=25centile | 30.68 | 30.81 | -1.99 | -0.47 | 4.68 | 20.78 | 52.33 | 113.68 | 113.68 |
| 25centile<Age<=50centile | 29.31 | 31.37 | -2.10 | -0.65 | 3.72 | 19.71 | 50.78 | 121.80 | 121.80 |
| 50centile<Age<=75centile | 26.40 | 28.33 | -2.54 | -0.87 | 3.46 | 18.13 | 41.43 | 110.20 | 110.20 |
| Full-time | 18.34 | 22.04 | -19.72 | -2.00 | 2.88 | 11.74 | 31.17 | 47.53 | 92.59 |
| Male | -11.00 | 20.24 | -41.71 | -41.71 | -19.24 | -6.83 | -0.84 | 2.25 | 6.29 |
| Experience | 29.02 | 27.44 | -2.42 | -0.20 | 5.26 | 22.55 | 49.11 | 93.83 | 93.83 |
| Experience ${ }^{12}$ | -14.09 | 19.19 | -40.29 | -40.29 | -25.75 | -11.88 | -3.09 | 0.98 | 2.14 |
| Low-Education | -9.58 | 10.70 | -30.16 | -30.16 | -13.15 | -6.77 | -2.05 | 0.51 | 1.65 |
| High-Education | -3.61 | 34.44 | -26.34 | -26.34 | -5.13 | -2.22 | -0.14 | 2.14 | 16.73 |
| Duration of previous job | 86.28 | 83.57 | 1.43 | 3.78 | 19.95 | 63.80 | 129.73 | 252.20 | 252.20 |
| Number of previous jobs | -29.68 | 34.73 | -108.63 | -108.63 | -39.66 | -17.04 | -5.22 | -1.02 | 0.50 |
| Individual effect | -5.27 | 34.55 | -48.48 | -39.40 | -10.84 | -1.23 | 2.36 | 18.60 | 18.60 |
| Residual from the entry wage equation | -18.70 | 44.77 | -70.31 | -70.31 | -25.82 | -11.53 | -2.27 | 1.37 | 5.01 |

## Results of the Job-duration equation

- Main Results :
- Old hires stay less
- Part-time hires stay less
- Men stay longer than women
- Low and high-education workers stay less in most firms
- Experienced workers stay longer periods in most firms
- The number of previous jobs and previous seniority are very important determinants of duration
- For all variables, coefficients are statistically different from 0 for $75 \%$ of all firms

|  | Distribution of the wage change equation coefficients |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Mean | StDev | 1 pctle | 5pctle | 1st quartile | Median | 3rd quartile | 95 pctle | 99pctle |
| Constant | -0.006 | 2.26 | -0.5884 | -0.1431 | -0.03372 | -0.01439 | 0.00239 | 0.14941 | 0.5599 |
| Year<=1980 | 0.009 | 0.11 | -0.2839 | -0.0714 | -0.00059 | 0.00698 | 0.01549 | 0.10244 | 0.341 |
| 1980<Year<=1989 | 0.006 | 0.11 | -0.3382 | -0.0639 | -0.00212 | 0.0073 | 0.0197 | 0.07301 | 0.3201 |
| Seniority<=2 | 0.008 | 2.26 | -0.3477 | -0.1022 | -0.00565 | 0.00914 | 0.03135 | 0.12844 | 0.4115 |
| 2 < Seniority <= 5 | -0.014 | 1.69 | -0.3654 | -0.108 | -0.02131 | -0.00891 | 0.00131 | 0.06782 | 0.2838 |
| 5 < Seniority <=10 | 0 | 0.98 | -0.282 | -0.0827 | -0.01126 | -0.00393 | 0.00323 | 0.05458 | 0.2361 |
| Change Part-Time to Full-Time | 0.578 | 0.37 | -0.0489 | 0.20137 | 0.44601 | 0.51536 | 0.64416 | 1.24671 | 1.9601 |
| Individual effect**(sen<=2) | -0.195 | 0.15 | -0.6423 | -0.4456 | -0.24773 | -0.15691 | -0.12349 | -0.04902 | 0.112 |
| Individual effect* $2<$ Sen $<=5$ ) | -0.043 | 1.73 | -0.4392 | -0.2108 | -0.05483 | -0.02807 | -0.0154 | 0.07553 | 0.2898 |
| Individual effect* $5<$ Sen<=10) | -0.027 | 1.99 | -0.5098 | -0.1856 | -0.04497 | -0.02665 | -0.01156 | 0.0668 | 0.3193 |
| Individual effect*(10<Sen) | -0.031 | 2.34 | -0.4638 | -0.1273 | -0.03488 | -0.02212 | -0.00846 | 0.06487 | 0.3171 |
| Male | 0.011 | 0.12 | -0.1486 | -0.04 | 0.00057 | 0.01221 | 0.01993 | 0.06442 | 0.1642 |
| Mills Ratio (transformed) | 0.006 | 1.42 | -0.1204 | -0.0338 | 0.00112 | 0.00825 | 0.01896 | 0.07414 | 0.2224 |
| Low-Education | -0.012 | 0.16 | -0.423 | -0.1282 | -0.02637 | -0.00906 | 0.00369 | 0.0905 | 0.3917 |
| High-Education | 0.059 | 0.38 | -0.5114 | -0.0894 | 0.02541 | 0.05272 | 0.08292 | 0.229 | 0.6736 |
| Initial residual*(sen<=2) | -0.311 | 0.17 | -0.8439 | -0.6053 | -0.38324 | -0.28533 | -0.23605 | -0.10022 | 0.0668 |
| Initial residual*(2<Sen<=5) | -0.047 | 0.71 | -0.5756 | -0.2801 | -0.06452 | -0.02382 | 0.00223 | 0.09678 | 0.308 |
| Initial residual* $(5<$ Sen $<=10$ ) | 57.091 | 11056.34 | -0.5899 | -0.2579 | -0.05218 | -0.00919 | 0.00892 | 0.07126 | 0.3518 |
| Initial residual*(10<Sen) | 140.771 | 20415.61 | -0.624 | -0.205 | -0.04138 | 0 | 0.03394 | 0.14149 | 0.4261 |
| Standard Deviation of the residuals | 0.527 | 0.21 | 0.1969 | 0.26859 | 0.41462 | 0.48394 | 0.58166 | 0.90553 | 1.444 |

Notes : Estimated using 6,598 firms (including agregates), returns to experience are substracted from wages

| Distribution of the Students of the wage change equation coefficients |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Mean | StDev | 1 pctle | 5pctle | 1st quartile | Median | 3rd quartile | 95 pctle | 99pctle |
| Constant | -0.9807 | 2.0479 | -5.9445 | -4.3054 | -1.9948 | -0.7582 | 0.0874 | 1.615 | 5.339 |
| Year<=1980 | 1.3357 | 3.3311 | -9.1621 | -1.9267 | -0.0717 | 0.8862 | 2.2204 | 10.196 | 10.196 |
| 1980<Year<=1989 | 2.0798 | 6.9921 | -11.791 | -3.0929 | -0.3361 | 0.9827 | 2.7318 | 23.203 | 23.203 |
| Seniority<=2 | 1.3219 | 2.8989 | -5.405 | -3.1079 | -0.2702 | 1.1407 | 2.2877 | 6.714 | 11.134 |
| $2<$ Seniority <= 5 | -1.2126 | 1.9752 | -8.6364 | -4.9146 | -2.0998 | -0.9958 | 0.055 | 1.492 | 2.946 |
| $5<$ Seniority <=10 | -0.5058 | 1.4398 | -4.4459 | -2.6147 | -1.2587 | -0.3005 | 0.2789 | 1.354 | 2.953 |
| Change Part-Time to Full-Time | 58.5302 | 50.432 | -0.4382 | 2.439 | 11.9887 | 48.1005 | 85.8519 | 158.624 | 158.62 |
| Individual effect*(sen<=2) | -19.5487 | 17.6551 | -65.213 | -59.113 | -32.3868 | -15.017 | -4.4619 | -0.517 | 0.784 |
| Individual effect*(2<Sen<=5) | -3.2051 | 3.2881 | -10.175 | -10.175 | -5.491 | -2.3799 | -0.644 | 0.838 | 1.927 |
| Individual effect*(5<Sen<=10) | -2.487 | 2.5006 | -7.1217 | -6.4827 | -4.4417 | -1.941 | -0.467 | 0.81 | 1.597 |
| Individual effect* $10<$ Sen $)$ | -2.5539 | 2.8666 | -10.212 | -9.9504 | -3.8744 | -1.7926 | -0.431 | 0.828 | 2.002 |
| Male | 1.64 | 2.8517 | -4.4001 | -4.4001 | 0.0139 | 1.5174 | 3.2666 | 6.095 | 8.797 |
| Mills Ratio (transformed) | 0.0063 | 1.4189 | -0.1204 | -0.0338 | 0.0011 | 0.0083 | 0.019 | 0.074 | 0.222 |
| Low-Education | -0.8393 | 1.366 | -4.0692 | -3.19 | -1.6885 | -0.6331 | 0.0705 | 1.039 | 1.954 |
| High-Education | 2.7451 | 2.955 | -2.2255 | -0.6928 | 0.5632 | 1.9874 | 4.0881 | 9.567 | 9.567 |
| Initial residual*(sen<=2) | -27.9567 | 25.7496 | -96.057 | -96.057 | -40.2169 | -21.2785 | -6.6026 | -1.115 | 0.515 |
| Initial residual* $2<$ Sen<=5) | -1.698 | 4.0379 | -12.987 | -6.3549 | -3.4212 | -1.4644 | 0.0917 | 3.838 | 8.446 |
| Initial residual* ${ }^{*}$ <Sen<=10) | -0.6522 | 2.7152 | -7.6591 | -4.793 | -1.9161 | -0.3172 | 0.7653 | 3.15 | 5.034 |
| Initial residual* $10<$ Sen $)$ | 0.6641 | 3.1899 | -7.1074 | -3.3688 | -1.2508 | 0 | 2.9227 | 6.923 | 10.259 |

Notes: Estimated using 6,598 firms (including agregates), returns to experience are substracted from wages

## Results of the Wage Change Equation

- Most coefficients have a symmetric distribution (around 0) but
- Going to Full-time work always increase wage
- High hiring wages always induce lower wage changes
- Mills ratio is almost never significantly different from 0
- For all variables, $25 \%$ of the coefficients are significantly different from 0 (except for entry wage*seniority, 75\%)


|  |  | Value-Added |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Parameter | Stderr | Parameter | Stderr |
|  | Intercept | 4.80 | 0.06 | 4.84 | 0.06 |
|  | Capital | 0.19 | 0.00 | 0.19 | 0.00 |
|  | Employment | 0.73 | 0.01 | 0.72 | 0.01 |
|  | ANC1 |  |  | 0.02 | 0.01 |
| 을 | ANC2 |  |  | -0.02 | 0.01 |
| $\frac{\square}{\square}$ | ANC3 |  |  | 0.01 | 0.01 |
| d | INDAN1 |  |  | 0.02 | 0.03 |
| $\stackrel{8}{\square}$ | INDAN2 |  |  | -0.02 | 0.01 |
| ¢ | INDAN3 |  |  | 0.00 | 0.00 |
| \% | INDAN4 |  |  | 0.00 | 0.00 |
| ${ }_{<}^{\text {® }}$ | PROGB |  |  | 0.02 | 0.01 |
| 3 | PROGH |  |  | -0.03 | 0.01 |
|  | DCOMP |  |  | 0.01 | 0.01 |
| ¢ | DINT |  |  | 0.01 | 0.01 |
| - | DMALE |  |  | 0.02 | 0.00 |
| $\stackrel{\rightharpoonup}{0}$ | DAGE1 |  |  | -0.03 | 0.01 |
| ¢ | DAGE2 |  |  | 0.03 | 0.01 |
| - $\overline{0}$ | DAGE3 |  |  | 0.01 | 0.01 |
| 0 | DANCPRE |  |  | 0.11 | 0.02 |
| 응 | DNEMPPRE |  |  | 0.18 | 0.03 |
| $\bigcirc$ | DPERSFE |  |  | -0.02 | 0.01 |
|  | DRESID |  |  | 0.04 | 0.01 |
|  | PERS25 | 0.32 | 0.08 | 0.27 | 0.08 |
|  | PERS50 | 0.30 | 0.12 | 0.28 | 0.12 |
|  | PERS75 | 0.23 | 0.07 | 0.26 | 0.07 |
| $\stackrel{0}{亏}$ | XB25 | 0.21 | 0.06 | 0.22 | 0.06 |
| $\stackrel{\square}{0}$ | XB50 | -0.10 | 0.09 | -0.08 | 0.09 |
| ¢ | XB75 | -0.04 | 0.07 | -0.04 | 0.07 |
|  | RES25 | -0.40 | 0.05 | -0.42 | 0.05 |
|  | RES50 | 0.52 | 0.12 | 0.52 | 0.12 |
|  | RES75 | 0.03 | 0.06 | 0.00 | 0.06 |

## Conclusions

- Employment Policies are very heterogeneous
- Returns to Seniority are very heterogeneous
- Persons with large person effects have lower returns to seniority (in particular, in the first years).
- Persons with unusually high starting wage rates on a particular job have lower returns to seniority (same remark).
- Turnover and employment policies matter for productivity more than compensation policies

