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## ABSTRACT

## Superstition, Conspicuous Spending, and Housing Markets: Evidence from Singapore*

For most people, buying a home is their single largest financial commitment. Previous research shows that Chinese buyers pay less for homes with unlucky addresses and more for homes with lucky addresses. Using Singapore data on housing transactions combined with a plethora of individual buyer characteristics including ethnicity, age, nationality, education, and employment, we study the source of these preferences. We find evidence that buyers are heterogeneous. Consistent with superstition, older people, those who suffered from more traffic accidents, and people buying new apartments have stronger preference for lucky addresses, while people with Western names and senior public-sector employees have weaker preference. Consistent with conspicuous spending, people with Western names, senior public-sector employees, and people buying in luxury districts have weaker preference for lucky addresses.

JEL Classification: D1, R3, Z1
Keywords: superstition, conspicuous spending, real estate, prices, behavioral economics

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## 1. Introduction

For most people, buying a home is their single largest financial commitment. However, it presents many uncertainties -- notably, the quality of the construction and maintenance, whether the home will fit the family's future needs, the development of the neighborhood, and the evolution of interest rates and housing prices. Research in psychology suggests that individuals rely on superstition as a way to cope with misfortune and uncertainty, and to rationalize a complex world (Vyse 1997; Ang et al. 2014; Zhang et al. 2014). Given the large stakes and many uncertainties, people may well rely on superstitious cues when buying a home.

Buying a home is also very likely an individual's largest visible expenditure. Behavioral economics suggests that individuals engage in conspicuous spending to signal their future wealth or status (Feltovich et al. 2002; Charles et al. 2009; Heffetz 2011). Where to live, the scale and design of a home - all are very visible, even conspicuous, and possible ways to signal wealth or status to potential business partners and spouses (Wei et al. 2012).

Prior research has found evidence of superstition among Chinese in buying homes, particularly, a preference for homes numbered " 8 " and dislike of homes numbered " 4 " (Shum et al. 2014; Fortin et al. 2014). However, these preferences could be expressions of conspicuous spending to signal future wealth or status, rather than or in addition to superstition. A house number has no intrinsic value. So, people might pay more for number " 8 " just to show that they can afford it. Just as they pay more to live in a more expensive neighborhood or build a flashier mansion.

Charles et al. (2009) and Heffetz (2011) emphasize the need to understand conspicuous spending in housing, yet qualify that such research is challenging due to confounding factors and measurement issues. Consider, for instance, the decision of where to live. Are people choosing to live in a tony neighborhood for the better amenities, to benefit from social externalities, or simply to show off that they can afford it?

We study the effects of superstition and conspicuous spending on housing markets in the context of Singapore. Singapore is very well suited for researching the issue for several reasons. First, the country is multi-racial, with a majority of Chinese and a minority of other ethnicities. With individual data on buyers, we can analyze demand and prices by nationality, ethnicity, age, education, occupation, and other individual characteristics. Second, most homes are high-rise, with all floors being similar if not identical, and so, we can analyze prices and buyer preferences while abstracting from substantial heterogeneity related to locations. Third, the government strictly regulates the numbering of high-rise buildings according to a fixed format, which cannot be manipulated by developers.

Specifically, under the competing hypotheses of superstition and conspicuous spending, we develop predictions for the effects of various demographic and economic
factors and uncertainty on buyer preference for lucky addresses. Next, we construct a novel dataset based on the legal filings of over 54,000 transactions of new high-rise apartments in the Singapore real estate market between the years 2000 and 2009, matched to data on ethnicity, dialect group, nationality, age, university education, public-sector employment, and prior involvement in traffic accidents. Then, we run an empirical contest between the alternative hypotheses of superstition and conspicuous spending.

We find evidence consistent with both superstition and conspicuous spending, suggesting that buyers are heterogeneous. The evidence of superstition comprises higher demand for lucky addresses from older buyers, those who suffered from more traffic accidents, and buyers of new apartments, which are consistent with older people and those encountering bad luck and facing more uncertainty being more superstitious. Further superstition is lower demand from buyers with Western names and senior public-sector employees, which is consistent with such people being less superstitious. The evidence of conspicuous spending is lower demand for lucky addresses in luxury districts, consistent with luxury areas being a substitute way of signaling wealth and status, as well as lower demand from buyers with Western names and senior public-sector employees, consistent with such people being more established and having less need to signal.

On balance, we conclude that superstition explains somewhat more of Chinese buyers' preference for lucky addresses. Yet, it is important to appreciate that conspicuous spending plays a role, particularly in accounting for the relative magnitude of the discount for unlucky addresses as compared with the premium for lucky addresses. In the absence of conspicuous spending, the theory of loss aversion (Kahneman and Tversky 1979) combined with superstition predict that the discount for unlucky addresses should be larger than the premium for lucky ones.

Our findings contribute to research in finance emphasizing behavioral biases among investors. Prior research has identified how lack of self-control affects savings, reference pricing affects the sale of homes, loss aversion affects the choice between investment in equities and bonds and leads investors to sell winners and hold on to losers, over-confidence precipitates excessive trading, and extrapolation induces investors to over-invest according to past share performance (DellaVigna 2009).

More particularly, weather affects personal mood, and in turn affects stock returns (Hirshleifer and Shumway 2003), risk attitudes in capital markets (Kliger and Levy 2003), and household spending (Agarwal et al. 2012). Solar and lunar eclipses are associated with lower trading volume and stock returns in U.S. and Asian equity markets, with larger effects when stock prices were more volatile and the eclipses were more publicized (Lepori 2009). In China, newly listed firms with lucky listing codes trade at a premium, which dissipates over three years (Hirshleifer et al. 2012). Consumers are subject to "hot hand" fallacies
(Guryan and Kearney 2008; Powdthavee and Riyanto 2015).
Our contribution is to explicate the role of superstition and conspicuous spending in the market for housing, which is the largest asset for most households. Although our study is set in Singapore, the findings are suggestive of consumer behavior in other markets. Older, less Westernized buyers and those who have encountered more bad luck are more prone to superstitious bias. Buyers employed in prestigious and well-paid occupations do not engage in conspicuous spending, and high-class locations and lucky addresses are substitute ways to signal status and wealth.

The rest of the paper is structured as follows. Section 2 explains Chinese numerology and Section 3 introduces Singapore's housing market. Sections 4-6 present the empirical methodology, data, and estimates, while Section 7 concludes.

## 2. Chinese Numerology

Chinese is a language with a common written script, spoken in different ways by region. In some Chinese dialects, the number " 8 " is pronounced like the word "prosperity", while the number " 4 " is pronounced like the word "death". Consequently, " 8 " is considered to be lucky, and " 4 " to be unlucky. As Hirshleifer et al. (2012) observe, Chinese numerological symbolism, which derives from Confucian and Taoist beliefs, deviates from scientific notions of causality in two ways. The similarity in the pronunciation of a number to another word has causal import. Further, association with the number (and so, indirectly with the word) will affect the likelihood of experiencing favorable outcomes.

The effect of Chinese numerological superstition has been quite extensively studied. The government of Hong Kong sells new license numbers by public auction. In auctions between 1997 and 2009, the prices of license numbers including the lucky " 8 " were systematically higher, while prices of license numbers including the unlucky " 4 " were lower ( Ng et al. 2010).

However, the premium for " 8 " could also be interpreted as conspicuous spending to signal future wealth or status. The interpretation as conspicuous spending is consistent with the premium for " 8 " falling after 2006, when the government introduced vanity license numbers, and so, providing another way by which car owners could signal their wealth and status. Moreover, the prices of license numbers including the digit " 5 " were systematically lower, which cannot be interpreted in terms of superstition.

Most research into Chinese numerological superstition has focused on real estate. Fortin et al. (2014) analyzed the sales of about 117,000 of single-family homes in the Greater Vancouver area between 2000 and 2005. In neighborhoods with relatively more Chinese residents and in repeated transactions, the sales prices of houses with street address numbers
ending in " 4 " were 2.2 percent lower and those ending in " 8 " were 2.5 percent higher than houses with other addresses. However, these findings might possibly be confounded by unobservable differences between houses, or explained by conspicuous spending. ${ }^{1}$

Shum et al. (2014) studied a random sample of high-rise residential sales in Chengdu, a large city in southwest China, between 2004 and 2006. The prices of apartments on floors with numbers ending in " 8 " were 8.7 percent higher when sold as second-hand but not significantly higher when sold as new. Buyers who had more 8 's in their mobile telephone number were more likely to buy apartments on the lucky 8th floor. However, more educated people were not less likely and older people were not more likely to buy 8th floor homes, which is not consistent with superstition. Indeed, the estimates seem more consistent with conspicuous spending. A mobile number with more 8 's is very conspicuous.

By contrast with the studies that focused on housing, Hirshleifer et al. (2012) investigated the effect of superstition on equity markets -- in the context of initial public offerings in China between 1991 and 2005. On the Shanghai and Shenzhen stock exchanges, listed companies are identified by a numerical code, which is the equivalent of the U.S. ticker. For instance, the listing code of the Industrial and Commercial Bank of China is 601398.SS. Consistent with superstition, the proportion of issuers with lucky listing codes (that included at least one lucky digit and no unlucky digit) was abnormally high and the proportion of issuers with unlucky listing codes was abnormally low. Moreover, newly listed shares with lucky listing codes initially traded at a premium to shares with unlucky listing codes, but the premium dissipated within three years. ${ }^{2}$

Besides numerological superstition, Chinese people also attach real meaning to the zodiac calendar, which comprises a 12 -year cycle, each named after an animal. The dragon year is considered to be particularly auspicious and is associated with surges of new births among Chinese people. In a study of the U.S. population, Johnson and Nye (2011) found that Asians born in dragon years are more educated than those born in other years, which suggests that Asian parents invest more in the education of dragon children.

## 3. Singapore Housing Market

Singapore is an island nation with land area of about 716 square kilometers (projected to increase to 766 square kilometers by the year 2030). As of June 2012, the population of the

[^1]country was 5.31 million, including 3.29 million citizens and 533,000 permanent residents. The resident population is ethnically diverse, comprising 74 percent Chinese and 26 percent other races. ${ }^{3}$

Singapore's ethnic diversity provides a natural setting in which to investigate the effect of superstition among Chinese people in their decision to buy a home. The nonChinese population provides a ready control group to test whether the effect of lucky and unlucky numbers on housing demand and prices is due to conspicuous consumption or superstition.

Singapore's home ownership rate of over 90 percent is among the world's highest. The residential market comprises two segments. About 77 percent of homes were built by the government's Housing and Development Board, mostly for direct sale to eligible citizens at subsidized prices, with a small proportion for rental. Public housing is primarily owneroccupied. The government closely regulates the sale and resale of public housing and limits information on individual prices. ${ }^{4}$

The remaining 23 percent of apartments are in the private housing market. The private housing market is laissez-faire, except that foreigners are not allowed to buy low-rise apartments such as single-family homes. Information on sales of private housing, including the price and physical characteristics of each transaction, is readily available. Accordingly, like previous studies of superstition in the Singapore housing market, we focus on private housing, where preferences for lucky and unlucky numbers are more likely to be fully revealed.

The Singapore government strictly regulates street addresses and the numbering of high-rise apartments, and so, facilitates study of the effect of superstition on housing prices. The Inland Revenue Authority of Singapore (IRAS), which is responsible for collecting taxes on real estate and real estate transactions, has stipulated a specific format for the numbering of all high-rise buildings. The format comprises a block number, and then an apartment number comprising a sequence of two digits or characters designating the floor, a hyphen, and two or three digits or characters designating the location on the floor. For instance, apartment number 05-08 is the 8th unit on the 5th floor.

The IRAS enforces compliance with the numbering plan. With regard to the numbering of floors, the IRAS states unequivocally, " 01 ' is normally assigned to the first level of a building where the main entrance is located, followed by ' 02 ' for the second level

[^2]and so on. ... Omission of any numbers is not allowed" (IRAS 2013: Paragraph 3B(1)). The IRAS is equally firm with regard to the numbering of apartments on each floor: "Units are generally numbered sequentially in a clock-wise direction starting from " 01 ". An apartment number will be provided for each distinct unit shown on the floor plan. Omission and reservation of any numbers are not allowed for a residential development" (IRAS 2013: Paragraph 3B(3)). In addition, the IRAS assigns a unique 6-digit postal code to every building. ${ }^{5}$

## 4. Empirical Strategy

Previous research into the demand for residential housing among Chinese buyers identified preferences over the house number (Fortin et al. 2014) as well as the floor number (Shum et al. 2014). To encompass both aspects, we define the ordinal variable, lucky address number, as follows:
$2=$ very lucky if the last digits of both floor and unit numbers is " 8 ";
$1=$ lucky if the last digit of either floor or unit number is " 8 " and neither last digit is "4";
$0=$ neutral if not very lucky, lucky, unlucky, or very unlucky;
$-1=$ unlucky if the last digit of either floor or unit number is " 4 " and neither last digit is " 8 ";
$-2=$ very unlucky if the last digits of both floor and unit numbers is " 4 ".
The above ordinal classification represents Chinese dislike of the number " 4 " and preference for the number " 8 " in a fairly flexible way. It explicitly allows complementarity in lucky/unlucky addresses, so, for instance, the premium for very lucky addresses need not be double the premium for lucky addresses. The only ambiguity is whether to classify mixed addresses like 04-08 and 08-04 in the same category as addresses with neither unlucky nor lucky numbers such as 03-07. In the main estimates, we classify mixed addresses as neutral, and in a robustness check, exclude them.

To investigate the effect of superstition and conspicuous spending on the demand for apartments, we estimate an ordered logit model of the buyer's choice by address.
Specifically, we suppose that buyer $i$ receives the latent utility,

$$
\begin{equation*}
y_{i j y q}^{*}=\alpha+S_{i j} \beta+X_{j} \gamma+\mu_{y}+v_{q}+\eta_{i j y q}, \tag{1}
\end{equation*}
$$

[^3]from buying apartment $j$ in quarter $q$ of year $y$. The $S_{i}$ are buyer $i$ 's demographic, economic, and other characteristics, $X_{j}$ are legal and physical attributes of the apartment including legal tenure, location, size, and floor level, $\mu_{y}$ are transaction year effects to control for changes over time in the macro-economy and interest rates, that affect all residential real estate transactions in the same way, $v_{q}$ are transaction quarter effects to control for potential seasonality, and $\eta_{i}$ is a random variable with a logistic distribution, while the $\alpha, \beta, \gamma$ are coefficients. The probability that buyer $i$ will buy apartment $j$ is
\[

$$
\begin{equation*}
P\left(y_{i y q}=j\right)=P\left(\kappa_{j-1} \leq y_{i y q}^{*} \leq \kappa_{j}\right)=F\left(y_{i t}^{*}-\kappa_{j}\right)-F\left(y_{i y q}^{*}-\kappa_{j-1}\right), \tag{2}
\end{equation*}
$$

\]

where $\kappa$ is the cut-off value and $j(=-2,-1,0,1,2)$ is the address, and $F$ is the logistic cumulative function where

$$
\begin{equation*}
F(z)=\frac{e^{z}}{1-e^{z}} . \tag{3}
\end{equation*}
$$

We assume that the buyer's characteristics are exogenous. Yet her decision to buy an apartment with a particular address could be correlated with legal and physical attributes of the apartment. If buyers with different characteristics differ in their preference for these legal and physical attributes, then, without proper control for these attributes, the estimated effects of the characteristics might be biased. For instance, if Chinese have a stronger preference for high floor apartments than non-Chinese, then Chinese buyers will be more likely to purchase apartments on the 8th floor than on the 4th floor even they are not superstitious. Accordingly, we include controls, $X_{j}$, for these attributes.

Our identifying assumption is that, in the absence of superstitious beliefs and conspicuous spending, buyer demand does not vary with address after controlling for legal and physical attributes. To justify the assumption, we note that the design of Singapore highrise residential buildings is fairly uniform, with the design of each floor being the same. So, the " 04 " unit is the same on every floor of the building, and similarly, the " 08 " unit is the same on every floor. Further, the government's strict regulation and clockwise numbering rule prevent real estate developers from selectively numbering more desirable apartments (for instance, closer to or further from the elevator) and home-owners from renumbering their apartments. Accordingly, the controls for legal and physical attributes substantially eliminate heterogeneity among apartments.

## 5. Data

Our study combines five sources of data. The first is a proprietary dataset of legal filings of private housing transactions with the Registry of Land Titles between January 1995 and December 2012. The records include the address of the apartment including the name of the building, and floor and unit number, attributes of the apartment, and details of the transaction.

The apartment attributes are the legal tenure (freehold or leasehold), type of development (apartment, condominium, or landed), and size in square meters. ${ }^{6}$ The transaction details are the date, names and national registration numbers of buyers and sellers, type of sale (new sale, subsale, or resale -- as explained below), price, and whether the purchase was financed by a mortgage loan.

A "new sale" is the sale of the apartment by the real estate developer, which may occur before or after the government issues the "temporary occupation permit", which allows the apartment to be inhabited. A "subsale" is the sale of an apartment before the government issues the temporary occupation permit, by a party who is not the developer. A "resale" is the sale of an apartment after the issuance of the temporary occupation permit, by a party who is not the developer.

The second source of data is a system that we built to classify any person by name into one of four ethnicities - Chinese, Malay, Indian, and others. The classification is based on the value of a linear combination of the characteristics of the data. We used a proprietary dataset of personal information to train the ethnic classifier system. When validated against a data set of 1,327 persons with confirmed ethnicity, the system achieved accuracy of over 99 percent in recognizing Chinese names and 82.7 percent in recognizing non-Chinese names (over 99 percent of persons classified as Chinese were actually Chinese, and 82.7 percent of those classified as non-Chinese were actually not Chinese). We used the classification system to identify the ethnicity (Chinese or non-Chinese) of buyers named in the residential property transactions. To supplement the race classifier, we also compiled family names of the major Chinese dialect groups in Singapore - Hokkien, Teochew, and Cantonese -- and classified Chinese buyers by dialect group.

The third source is a simple classifier that we built to infer a Singaporean's year of birth from their national registration number. The government issued national registration numbers in sequence by year of birth to people born in the year 1955 and after, and so, the classifier only works only for these people. For people born earlier, the national registration number was not issued in sequence, so, it is not possible to infer their age from the national registration number.

The fourth source is graduation lists of Singapore's two major tertiary institutions the National University of Singapore and Nanyang Technological University, and the Singapore Government Directory, 2007 and 2009 editions. For these sources, we kept only

[^4]unique names and then matched by name to buyers in the dataset of private housing transactions to identify buyers with university education or were employed in the public sector.

The fifth source is a proprietary dataset of court records of civil actions against personal defendants involved in traffic accidents between 1998 and 2012. For each case, the record includes the filing date and the names and national registration numbers of the parties. We link the court records to residential property transactions through the buyer's national registration number.

Note that many purchases of residential properties involve more than one buyer, such as husband and wife buying jointly. In classifying buyers by race, nationality, dialect group, education, and employment in the public sector, we classify all buyers involved in the purchase, and then construct indicator variables that equal 1 if at least one buyer belongs to the category and 0 otherwise. For instance, an apartment is defined to be bought by Chinese if at least one buyer is Chinese.

Like Shum et al. (2014), we focus on high-rise residential developments. The plans of the various floors in high-rise buildings are quite similar, if not identical. Possible differences from a common floor plan include first floor apartments which might include a terrace and top floor apartments which might include a terrace or comprise two floors. By contrast, single-family homes are heterogeneous and may differ in unobservable ways that correlate with the house number (Fortin et. al. 2014).

Unlike previous studies, we consider only the primary market, i.e., new sales. Apart from a few transactions with missing data or errors, the data on new sales are quite complete. Importantly, all new apartments in a building are sold by the same seller, and so, withindevelopment estimates would not be confounded by unobserved differences. In addition, the quality of new apartments is uniform within a building. By contrast, studies of the secondary market are subject to possible selection bias and differences among sellers. Moreover, resale apartments might differ in ways that we cannot observe, such as renovation and maintenance, and which might be correlated with unit or floor number.

Referring to Table 1, the distribution of sales by luckiness of address differs between the primary and secondary markets. Specifically, the Kolmogorov-Smirnov test rejects the null hypotheses of that the distributions of lucky addresses are the same in the primary and secondary markets at the 1 percent level. Subsales and resales involve relatively more unlucky addresses and fewer lucky addresses. ${ }^{7}$
[Insert Table 1 here]

[^5]The Urban Redevelopment Authority, a government agency, publishes limited data on all private housing transactions in the online database, REALIS. Importantly, REALIS does not publish the personal information of buyers and sellers. Figure 1 compares the numbers of new sales of private housing in the REALIS database and our proprietary dataset of legal filings. Evidently, the legal filing dataset provides good coverage from 2000 onward and up to 2009. Accordingly, we focus our study on transactions in the period, 2000-09. Transactions are cyclical, peaking in 2002, and falling sharply in 2003, the year of the SARS (Severe Acute Respiratory Syndrome) crisis. Transactions recovered gradually to a new peak in 2007, before falling sharply in 2008, on the advent of the global financial crisis.
[Insert Figure 1 here]
Among the legal filings, we exclude transactions with incomplete or obviously wrong information, and those with corporate buyers, and limit to new sales of high-rise apartments in the period 2000-09. The dataset comprises 54,109 new sale transactions. In additional analyses, we also consider transactions in the secondary market, i.e., subsales and resales.

## [Insert Table 2 here]

Table 2 presents summary statistics of the data. The average real price is $\mathbf{S} \$ 10,601$ per square meter $(2014=100)$ or about US $\$ 581$ per square foot (converted at the average exchange rate of one Singapore dollar to 59 U.S. cents in 2014). The size of the average apartment is 118.7 square meters or 1277.7 square feet, and so, the average price of an apartment is US $\$ 742,344$.

For a first look at the data, Figure 2 plots the estimated coefficients of the floor levels in a regression of the logarithm of unit price per square meter on indicators of floor number and last digit of unit number, apartment size, and building fixed effects. The dotted and solid lines indicate floors with last digit ending in 4 and 8 , respectively. Evidently, the price increases with the floor level with a noticeable dip at 47 . The relationship between floor and price is weaker in buildings with 45 or more floors, of which there are only nine.
[Insert Figure 2 here]
Figure 3 plots the estimated coefficients of the last digit of the unit number in a regression of the logarithm of price per square meter on indicators of the last digit of the unit number. Clearly, the prices of units with numbers ending in " 4 " are lower, while the prices of units with numbers ending in " 8 " are higher.
[Insert Figure 3 here]
With regard to nationality and ethnicity, 80.9 percent of the buyers are Singaporeans (citizens or permanent residents), and just over 91 percent are Chinese. The ethnic classifier system is accurate in classifying the ethnicity of Singaporeans, but less accurate for foreigners. Among the foreign buyers, there are Indonesian Chinese with Malay-like names
which the ethnic classifier might wrongly classify as non-Chinese.

## 6. Estimates

### 6.1. Numerological Preferences

Following up on Figures 2 and 3, we present regression analyses of transaction prices and buyer choice to show that, in the Singapore housing market, ethnic Chinese buyers prefer apartments with the digit " 8 " and dislike apartments with the digit " 4 ". These estimates present a baseline to confirm the preferences of Chinese buyers, as shown in previous research (Ng et al. 2010; Shum et al. 2014; Fortin et al. 2014). However, we stress that these baseline results do not explain the reason for such preferences, whether due to superstition or conspicuous spending.

Table 3 presents ordinary least squares (OLS) regressions of the price of housing (specified as the logarithm of price per square meter). Table 3, columns (a)-(c), report estimates that control for size, floor, legal tenure (freehold) and development type (condominium), as well as progressively finer controls for location, beginning with no control for location, then 2-digit postal district fixed effects, and then 4-digit postal district fixed effects. Our data comprises transactions in 69 out of a total of 802 -digit postal districts in Singapore. These 2-digit postal districts are further divided into 721 4-digit postal districts. Each 4-digit district consists of between 1 and 37 apartment buildings, with an average of less than 6 . Hence, 4 -digit district fixed effects control much more stringently for differences between neighborhoods than do 2-digit fixed effects.

Regardless of whether we control for postal districts, apartments with very unlucky addresses are always the cheapest and those with very lucky addresses the most expensive ones. Controlling for 2-digit postal districts reduces the estimated discount for very unlucky addresses by 2.6 percentage points from 4 to 1.4 percent and the premium of very lucky addresses by 1 percentage point from 4 to 3 percent. This is likely due to the variation in the proportion of apartments with lucky addresses across postal districts.

In Singapore, the maximum height of each building, as stipulated by the government master plan, varies across districts. Buildings with fewer than eight floors do not have any very lucky addresses. Hence, in districts with relatively more short buildings, the supply of very lucky numbers would be lower. If the average housing price also varies across districts, then failing to control for geographic location biases the estimated impact of lucky addresses on housing price. Except for the premium for very lucky addresses, the estimated relationship between lucky number and price hardly changes if we control for 4-digit rather than 2-digit postal districts, which suggests that controlling for 4-digit postal district resolves most of the bias due to across district variation in the proportion of apartments with lucky addresses. Indeed, the coefficients of lucky addresses hardly change if we control more stringently for 6-
digit postal codes, which are unique to buildings. To optimize the number of fixed-effects in the regressions, we decided to control for 4-digit postal districts in all following analyses.
[Insert Table 3 here]
The estimates including postal district fixed effects identify the effects of the buyer's ethnicity by variation within districts. But the impacts of lucky addresses on housing prices are restricted to be the same for both Chinese and non-Chinese. One might argue that nonChinese might not want to pay a premium for an address that Chinese consider to be lucky. Table 3, column (d), relaxes this restriction by adding interactions between ethnicity and addresses. To make our results comparable across columns, we use Chinese as the reference group. As a results, the coefficients on lucky addresses have the same economic meanings across the columns.

Evidently, Chinese buyers pay higher prices for luckier addresses. The regressions exclude the indicator variable for neutral addresses, which is the benchmark. Referring to Table 3, column (d), consistent with Chinese numerological preferences, Chinese buyers pay increasing premia for the lucky addresses, with the premium for the very lucky addresses (the last digit of floor and unit numbers is " 8 ") being 2.4 percent, and receive increasing discounts for unlucky addresses, with the discount for the very unlucky addresses (the last digit of floor and unit numbers is " 4 ") being 1.2 percent. Except for the coefficients on the interaction between non-Chinese and very unlucky addresses, the coefficients on the coefficients for the other three categories of addresses are all insignificant, suggesting that non-Chinese buyers are price takers with respect to luckiness of address although they do not have any systematic numerological preference. ${ }^{8}$

While non-Chinese buyers might have to pay similar premium or enjoy the similar discount for lucky or unlucky addresses, they do pay 1.3 to 3.6 percent more in general. This is likely due to the fact that most non-Chinese buyers are not Singaporean. As foreigners, they might be at an information disadvantage relative to locals. Excluding all foreign buyers, the price premium paid by non-Chinese falls to 0.7 percent and the difference is only marginally significant. For brevity, we report these results in the Supplement.

To check whether the results are stable over time or exist only in the primary market, Table 3, column (e), extends the sample to the period of 1995-2013, and column (f) examines

[^6]the relationship in the secondary market (subsales and resales). These estimates confirm that the findings with respect to Chinese buyers and pricing of unlucky addresses are robust to including earlier and later transactions as well as in the secondary market.

The pattern of prices for unlucky addresses relates to previous research into reference dependence in residential housing markets (Genesove and Mayer 2001). To the extent that people pay less for unlucky addresses, they will ask for lower prices when they subsequently sell their homes. Indeed, we find similar discounts for unlucky addresses in the secondary and primary markets.

Referring to Table 3, column (c), relative to a neutral address, the discounts for very unlucky and unlucky addresses are 1.5 and 0.7 percent, while, the premia for very lucky and lucky addresses are 2.2 and 0.5 percent. The discounts are similar to or smaller than the premia. By contrast, the theory of reference dependent preferences and loss aversion (Kahneman and Tversky 1979) predicts that individuals would be less willing to accept bad things than pay for good things.

In our context, since the majority of apartments have neutral addresses, these apartments are likely to be used as the reference when buyers make their purchasing decisions. The discount for an unlucky address could be viewed as the price that a buyer is willing to accept for potential bad luck in the future. In contrast, the premium from a lucky address is the price that a buyer is willing to pay for potential good luck in the future. Due to risk and/or loss aversion, the discounts for unlucky addresses should be larger than the premia for lucky addresses if the expected gain from purchasing a lucky address is the same as the expected loss from buying an unlucky one. Yet, our estimates show otherwise, which seems inconsistent with the superstitious belief that unlucky addresses bringing bad luck and lucky addresses bringing good luck.

The alternative explanation of conspicuous consumption seems more plausible. This theory emphasizes people spending more to signal status and wealth. It is silent on negative signaling (which is presumably the outcome of buying a home with an unlucky address). Hence, the empirical pattern of discounts for unlucky addresses being similar to or smaller than the premia for lucky addresses seems to fit the hypothesis of conspicuous consumption better than superstition.

Having studied the pricing of luckiness of address, we next investigate buyer's choice of apartment by luckiness of address. Table 4 reports ordered logit regressions of choices ordered categorically from the very unlucky addresses (last digit of floor and unit numbers is " 4 ") to the very lucky addresses (last digit of floor and unit numbers is " 8 ").

$$
\text { [Insert Table } 4 \text { here] }
$$

Table 4, column (a), presents an estimate for all buyers in new sales including
controls for legal and physical attributes and year and quarter of purchase. The coefficient of Chinese, 0.183 (s.e. 0.033), is positive and significant. This implies that the odds of a Chinese person buying an apartment with a very lucky address as compared with other addresses are $e^{0.183}=1.201$ times, or 20.1 percent higher than a non-Chinese person. Table 4, column (b), includes 2-digit postal district fixed effects to control for differences in location. The coefficient of Chinese, 0.220 (s.e. 0.032), is positive, significant, and larger than the estimate without district fixed effects. This suggests that the preference for lucky addresses is stronger within districts than across districts, perhaps because of differences in the supply of lucky addresses across districts. One reason for the difference in supply is geographical differences in height regulation. In areas where buildings are limited to 4 or 6 floors, there are no very lucky addresses. Consequently, the demand for lucky addresses is compressed, which would likely attenuate the estimated coefficient of Chinese buyer.

Table 4, column (c), includes 4-digit postal district fixed effects, which control more stringently for differences in area. The coefficient of Chinese buyer, 0.223 (s.e. 0.034), is quite similar to the estimate with 2-digit district fixed effects, suggesting that the 2-digit fixed effects already substantially control for differences in neighborhood. Nevertheless, to be conservative, we prefer the estimate with 4-digit district effects as it controls more stringently for locational differences. ${ }^{9}$

Next, Table 4, column (d), reports an estimate that investigates differences in the preferences over lucky addresses between Singaporean and non-Singaporean Chinese buyers. The benchmark is the non-Singaporean non-Chinese buyers. The coefficient of Chinese buyer is positive and significant even at the 1 percent level. However, the coefficient on Singaporean Chinese buyers is positive but insignificant, suggesting the preference for lucky addresses is much stronger for Singaporean Chinese than foreign Chinese. Apparently, nonSingaporean Chinese buyers have somewhat weaker preferences over lucky addresses. This difference should be interpreted with caution owing to the difficulty of classifying Indonesian Chinese buyers.

Then, Table 4, column (e), reports an estimate that considers differences in choice between luxury ( 2 -digit postal districts $1,9,10$, and 11) and other areas. The coefficient of Chinese buyers increases to 0.274 , which is considerably larger than the preferred estimate. The coefficient of the interaction between Chinese and luxury areas is $-0.135(p=0.068)$. Apparently, the preference among Chinese buyers for lucky addresses is limited to non-

[^7]luxury areas. Buying a home in a luxury area is conspicuous in itself. It is certainly a credible signal as homes in luxury areas cost more than those in ordinary neighborhoods. The weaker preference for lucky addresses in luxury areas may be the result of buyers using the choice of residential neighborhood to substitute for lucky addresses to signal wealth and status. To this extent, the estimate for luxury areas is consistent with the hypothesis of conspicuous spending.

Finally, Table 4, columns (f)-(i), reports tests of the robustness of the basic finding that Chinese buyers prefer lucky addresses. These variously exclude mixed addresses such as 04-08, exclude buildings with fewer than eight floors (which have fewer lucky addresses and no very lucky addresses), include earlier transactions back to 1995 and later transactions forward to 2012, and focus on the secondary market. Across all of these variations, the coefficient of Chinese buyer is positive and significant.

Referring to Table 4, column (g), the coefficient of Chinese buyer in the subsample of buildings with more than eight floors is larger than the preferred estimate. In buildings with fewer than eight floors, there are fewer lucky addresses and no very lucky addresses. This compresses buyers' choices, which is the dependent variable, and so attenuates the estimated effect of Chinese buyers. Hence, the estimated effect in buildings with more than eight floors is larger.

Table 4, column (i), reports an estimate of buyer choice in the secondary market. The coefficient of Chinese buyer, 0.057 (s.e. 0.025 ), is positive and significant, but substantially smaller than the preferred estimate (Table 4, column (c)). The difference could be due to the different distribution of lucky and unlucky addresses in the secondary market. In the new sale market, buyers can get their preferred address as long as they pay for it. By contrast, in the secondary market, buyers must choose among apartments that are offered for sale. Buyer preference for lucky address may be dominated by preference for location. Hence, even Chinese buyers with relatively strong preference for lucky addresses might buy apartments with unlucky addresses to get their preferred location. Similarly, non-Chinese buyers might pay the market premium for apartments with lucky addresses although they would prefer apartments without lucky addresses.

Referring to Table 1, the proportion of apartments with an unlucky or very unlucky address is 19.9 percent in the primary market and 22.4 percent in the secondary market, and the difference is statistically significant at the 1 percent level. So, it is difficult to avoid these apartments in the secondary market. By contrast, the proportion of apartments with a lucky or a very lucky address is 13.5 percent in the primary market and 11.9 percent in the secondary market, and the difference is statistically significant at the 1 percent level. Hence, it is more difficult for buyers to purchase an apartment with a lucky address in the secondary market.

### 6.2. Superstition or Conspicuous Spending

Having confirmed the Chinese numerological preference, we now turn to investigate whether the preference is due to superstition or conspicuous spending. Under the alternative hypotheses of superstition and conspicuous spending, we analyze how demographic and microeconomic factors, and luck and uncertainty would affect buyer's choice of address. Then, we run an empirical contest between the two hypotheses, to infer which one better fits the data. Owing to possible measurement error in classifying non-Singaporean Chinese (historically, Indonesian Chinese were required to adopt Indonesian names), we limit this analysis to Singaporean Chinese buyers.

The first basket of factors is demographic - the buyer's dialect group, whether the buyer has a Western name, the buyer's age, and income. Singapore's ethnic Chinese population mostly comprises the descendants of people who emigrated from South China in the late nineteenth and the first half of the twentieth century. The three major dialect groups are the Hokkiens ( 40 percent), followed by Teochews ( 20 percent) and Cantonese (15 percent).

The buyer's dialect possibly affects their superstition about home addresses. The pronunciation of " 4 " and " 8 " differs between the dialects. In all three major dialects, " 4 " sounds like "death". However, " 8 " sounds like "luck" only in Cantonese, but not in Hokkien or Teochew. To this extent, the Cantonese should have a stronger preference for lucky numbers than the Hokkiens and Teochews.

Research on consumer spending in the United States and South Africa has found significant variation in conspicuous spending by race, but has not identified any variation within races (Charles et al. 2009; Heffetz 2011; Kaus 2013). Accordingly, here, we posit that the demand for conspicuous spending does not vary by dialect group.
[Insert Table 5 here]
Table 5, column (a), reports the estimates, with other dialects (not Hokkien, Teochew, or Cantonese) as the omitted dialect group. The coefficients of Cantonese, Hokkien and Teochew buyers are all positive but not significant at the conventional levels. Evidently, the demand for lucky numbers is the weakest among the Cantonese, which is not consistent with the similarity in pronunciation being closest in their dialect. However, we should caution that the classification by dialect is limited by many family names belonging to multiple dialects. We classify a person in a dialect only if their family name is unique to that dialect. If there are several buyers, the indicator for Cantonese is 1 as long as one of them has a family name that is unique to that group. Consequently, many buyers in the omitted group might actually be Hokkiens, Teochews, or Cantonese.

Another demographic factor is the buyer having a Western given name like "Susan"
or "John". ${ }^{10}$ Such a name may indicate more Westernized thinking, and so, less belief in Chinese numerological superstitions (and, possibly, more belief in Western religion). In the U.S. population, an individual's given name is strongly related to their parents' economic status (Olivetti and Paserman 2014). In Singapore, a former British colony, English-educated people earn relatively more and may tend to give Western names to their children. To the extent that individual's status has been established by their parents, they would have less need to signal their status. This reasoning suggests that, in Singapore, people with Western names might spend less on conspicuous spending.

We emphasize that the names of buyers and sellers recorded in the legal filings of residential transactions are the names of the persons as officially registered in the national registration system. The transactions do not include nicknames. As reported in Table 5, column (b), the coefficient of Western name, -0.049 (s.e. 0.022), is negative and significant, which suggests that the odds of buying an apartment with a very lucky address as compared with other addresses is $e^{-0.049}=0.952$ times, or 4.8 percent smaller for a buyer with a Western given name than buyers without a Western given name. This result is consistent with such buyers being more Westernized and less believing of Chinese numerological superstition, and also consistent with their coming from an established background and engaging less in conspicuous spending.

Yet another demographic factor is the buyer's age. Both superstition and the demand for conspicuous spending vary with age. Johnson and Nye (2011) find that the parents of children born in the dragon year are older, which suggests that older people are more superstitious. Separately, older people are more religious (Deaton 2011), which we interpret as additional, indirect evidence that older people might be more superstitious. With regard to conspicuous spending, younger people need to establish themselves, and so, their demand for signaling is relatively higher. By contrast, older people have already established themselves, and so, their demand for signaling is relatively lower (Charles et al. 2009). Thus, older people should engage less in conspicuous spending on lucky addresses.

Buyer age is especially important as it is one of only two factors for which the hypotheses of superstition and conspicuous spending generate contrasting predictions. We divide buyers into three age brackets according to the mean age of the buyers: 34 and below, between 35 and 45 , and 46 and above, which are the bottom quartile, middle two quartiles, and top quartiles respectively. Table 5, column (c), reports an estimate with the middle age group as the omitted category. The coefficient of the mean buyer age being 46 and above, 0.124 (s.e. 0.028 ), is positive and significant at the 1 percent level, while the coefficient of the mean buyer age being 34 and below is negative and insignificant.

Evidently, older buyers evince a stronger preference for lucky numbers than middle-

[^8]aged and younger buyers. This pattern is consistent with superstition but not conspicuous spending.

The second basket of factors is microeconomic - education, employment, and wealth. People with more education should be less superstitious (Mocan and Pogorelova 2014). However, Johnson and Nye (2011) find some but not conclusive evidence that Asian mothers giving birth in a dragon year are more educated. By contrast, with regard to conspicuous spending, university graduates already have their degree as a strong objective indicator of status, and so, they should benefit less from conspicuous spending.

Among university graduates, those with degrees in science and technology should be less superstitious than graduates in other disciplines. By contrast, we cannot think of any difference between graduates in science and technology vis-à-vis other disciplines in their demand for conspicuous spending.

The classification of university graduates is subject to considerable false negatives and false positives as it is based on matching the graduation books of Singapore's two major universities by name to legal filings of residential transactions. The reasons for false negatives are that, in matching the two data sources, we limit to unique names in the graduation lists, and so, omitting people with common names like "Lee Li Lian", and our procedure overlooks graduates of foreign universities. The reason for false positives is that a name that is unique in the graduation list need not be unique in the population.

With these provisos in mind, Table 5, columns (d)-(f), report ordered logit regressions focusing on buyers with university degrees, buyers with university degrees in science and technology, and both types of buyers respectively. Subject to the proviso that the effect of education is not precisely estimated, the effect of university education seems to be concentrated among graduates with degrees in science and technology. The imprecision of the estimates may well be due to false negatives or false positives in identifying buyers with university degrees. With great caution, we interpret the estimates as slight evidence in favor of superstition.

Another microeconomic factor is the individual's profession. Using the Singapore Government directory, we identify senior level public sector employees. The omission of lower level employees is not serious as they cannot afford private housing. In Singapore, employment in the public sector is prestigious and well-remunerated. The public sector selects senior employees through rigorous processes, which possibly screen out individuals who are superstitious. Certainly, in Singapore, senior public-sector employees have wellestablished status and enjoy high incomes, and so, they have no need to engage in conspicuous spending to signal either status or wealth.

The classification of senior public-sector employees is subject to considerable false negatives and positives as it is based on matching the Singapore Government directory by
name to legal filings of residential transactions. To match the two data sources, we limit to unique names in the Government directory.

As Table 5, column (g), reports, the coefficient of the buyer being a senior publicsector employee, -0.139 (s.e. 0.039), is negative and significant, suggesting that the odds of a senior level public-sector employee purchasing a very lucky address are $e^{-0.139}=0.870$ times, or 13 percent less than that of other buyers. This result is consistent with public-sector employees both being less superstitious and engaging less in conspicuous spending. However, an alternative interpretation of the empirical result is that the Singapore government penalizes corruption severely, and so, public-sector employees refrain from conspicuous consumption to avoid attention from the anti-corruption agency.

Yet another microeconomic factor is wealth. Johnson and Nye (2011) find some but not conclusive evidence that Asian mothers giving birth in a dragon year have higher incomes, which suggests that wealthier people are more superstitious. Feltovich et al. (2002) predict unequivocally that people in the highest status do not signal. Accordingly, wealthier people would not engage in conspicuous spending. Buyer wealth is especially important as it is one of only two factors for which the hypotheses of superstition and conspicuous spending generate contrasting predictions.

Lacking direct information on buyer wealth, we proxy buyer wealth in two ways. One is whether they pay for the purchase in cash or through a mortgage loan. Presumably, people who can afford to pay for a home in cash are wealthier than other buyers. The other proxy is choice of residential area. People who buy in more expensive areas are probably wealthier.

Table 5, column (h), reports an estimate including an indicator of a purchase financed by a mortgage. The coefficient of mortgage, -0.047 (s.e. 0.030 ), which is not consistent with superstition. The negative estimate is consistent with conspicuous spending, but it is imprecise.

Separately, as Table 4, column (e), reports, the demand for lucky addresses in luxury districts among Chinese buyers is weaker. The coefficient of the interaction between luxury districts and Chinese buyer, -0.135 (s.e. 0.074), is negative and marginally significant. This result is not consistent with wealthier people being more superstitious, but is consistent with wealthier people engaging less in conspicuous spending. The result can also be interpreted in another way consistent with the hypothesis of conspicuous spending. Buying in a luxury district is already a strong signal, relative to which the marginal benefit of a lucky address might be small and so people who buy homes in luxury areas signal status and wealth by location rather than lucky address.

To check possible confounds among the demographic and economic factors, Table 5, column (i), reports an estimate including all of the factors, except university degree, and

Table 5, column (j), reports an estimate including all of the factors. Three findings are robust to the inclusion of other covariates: the negative effect of a Western name, the positive effect of age, and the negative effect of senior public-sector employment. In particular, the negative effect of a university science degree is not robust, possibly because university graduates tend to be younger than average buyers, and so, the university degree partly captures the negative effect of youth.

The third basket of factors is luck and uncertainty. Research in psychology suggests that individuals rely on superstition as a way to cope with misfortune and uncertainty (Vyse 1997; Ang et al. 2014; Zhang et al. 2014). This suggests that people who experience negative shocks or bad luck should more strongly prefer lucky addresses. By contrast, we do not know of any reason why past bad luck should affect conspicuous spending.

We characterize bad luck by the number of civil legal actions for negligence in traffic accidents prior to buying the apartment, distinguishing between the number of legal actions as plaintiff and the number as defendant. The plaintiff, being the victim, is the more unlucky. The defendant is also unlucky but to a lesser degree, considering that they (were alleged to have) caused the accident.

Table 6, column (a), reports an estimate including the number of traffic accidents leading to legal actions involving the buyer as plaintiff (if the purchase involves more than one buyer, the variable is the simple average of the number of actions involving each of the buyers). The coefficient, 0.102 (s.e. 0.042 ), is positive and significant. Table 6, column (b), reports an estimate including the mean number of traffic accidents leading to legal actions involving the buyer as defendant. The coefficient is positive, smaller than the coefficient of the buyer as plaintiff, but not precisely estimated. These results are consistent with people relying on superstition to cope with bad luck, and particularly the victim of a traffic accident being unluckier than the injurer.

## [Insert Table 6 here]

Another characterization of bad luck or uncertainty is SARS, which we characterize by the months of March to June, 2003. The coefficient of SARS is positive but not significant. However, the interpretation of this coefficient is possibly confounded by the sharp concurrent macro-economic downturn, causing an income effect on the demand for lucky addresses.

Yet another source of uncertainty is the quality of development. In Singapore, most new high-rise apartments are sold before construction, and so, the buyer must rely on the plan and a demonstration mock-up. It was only in July 2015 that the government issued regulations stipulating that the demonstration apartment or model accurately represent the actual apartment for sale and that location plans be drawn to scale and truthfully represent the neighborhood. Buyers of new apartments face more uncertainty about construction and
appearance than buyers of resale apartments. If they are superstitious, they should more strongly prefer lucky addresses and avoid unlucky addresses in the primary market. By contrast, the hypothesis of conspicuous consumption does not have different implications for the primary vis-à-vis secondary market.

Referring to Table 4, columns (c) and (i), the coefficient of Chinese buyers is 0.223 (s.e. 0.034) in the primary market (new sales) and 0.057 (0.025) in the secondary market (subsales and resales). Indeed, Chinese buyers evince stronger preference for lucky addresses in the primary market than the secondary market, which is consistent with people relying more on superstition when facing more uncertainty.

## 7. Concluding Remarks

Applying empirical analysis to a very rich individual-level dataset of purchases of high-rise apartments in Singapore, we find evidence of both superstition and conspicuous spending, suggestive of heterogeneity among home buyers. The evidence of superstition is higher demand for lucky addresses from older buyers, those who suffered from more traffic accidents, and people buying new as contrasted with second-hand apartments, and lower demand from buyers with Western names and senior public-sector employees. The evidence of conspicuous spending is lower demand for lucky addresses in luxury districts and lower demand from buyers with Western names and senior public-sector employees. On balance, weighing the evidence for the two hypotheses, we consider that superstition explains a somewhat larger part of Chinese buyers' preference for lucky addresses.

However, it is important to appreciate that conspicuous spending is also essential to understanding the demand for housing. The theory of reference dependent preferences and loss aversion (Kahneman and Tversky 1979) predicts that a buyer's willingness to accept an unlucky address would be less than her willingness to pay for a lucky address. Yet, our estimates (Table 3, column (c)), show that the discounts for unlucky addresses are not larger than the premia for lucky addresses, which is not consistent with reference dependent preferences and loss aversion. Conspicuous spending helps to explain this pattern of relative prices.

Previous research emphasized the importance of understanding the role of conspicuous spending in the demand for housing (Charles et al. 2009; Heffetz 2011). Here, we find evidence of conspicuous spending on two particularly visible aspects of the home the address and location. While our study is specific to Singapore, the findings are suggestive of consumer behavior in other housing markets. It is quite intuitive that home buyers in other markets would pay more for homes in luxury neighborhoods to signal their wealth and status. For instance, Los Angeles realtors advertise homes near to Beverly Hills as "Beverly Hills adjacent". Conspicuous spending on home location is an interesting
direction for future research.
Our findings also tie in with previous research into reference dependence in residential housing markets (Genesove and Mayer 2001). To the extent that people pay less for unlucky addresses, they will demand lower prices when they subsequently sell their homes. Indeed, we find similar discounts for unlucky addresses in the secondary and primary markets. The interaction between reference dependence and conspicuous spending on home location is yet another interesting direction for future research.

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Table 1. Lucky addresses: Type of sale and buyer ethnicity

| VARIABLES | $(\mathrm{a})$ | $(\mathrm{b})$ | $(\mathrm{c})$ | $(\mathrm{d})$ |
| :--- | :---: | :---: | :---: | :---: |
|  |  |  |  |  |
|  | All buyers | New sales | Chinese | Non-Chinese | \(\left.\begin{array}{c}Subsales <br>

and resales: <br>
All buyers\end{array}\right\}\)

Notes: The samples comprise apartments sold in the years 2000-09. Each cell reports mean and standard deviation. The apartment is considered as bought by Chinese buyers as long as one of the buyers is a Chinese. By the Kolmogorov-Smirnov test, the distribution of addresses in the primary market (new sales) is significantly different from that in the secondary market (subsales and resales) at the $1 \%$ level, and the share of apartments with lucky or very lucky addresses is significantly higher in the primary market than the secondary market at the $1 \%$ level.

Table 2. Summary statistics: New sales

| VARIABLES | Unit | (a) <br> Mean | (b) <br> s.d. | (c) <br> Min. | (d) Max. |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Transactions |  | 54,098 |  |  |  |
| Apartment size | Square meters | 118.7 | 44.9 | 5.00 | 768.00 |
| Unit price | S\$ per square meter | 10,601 | 5,255 | 3,410 | 68,313 |
| Very unlucky |  | 0.013 | 0.114 | 0 | 1 |
| Unlucky |  | 0.186 | 0.389 | 0 | 1 |
| Neutral |  | 0.666 | 0.472 | 0 | 1 |
| Lucky |  | 0.129 | 0.335 | 0 | 1 |
| Very lucky |  | 0.006 | 0.077 | 0 | 1 |
| 99-year leasehold |  | 0.494 | 0.500 | 0 | 1 |
| 999-year leasehold |  | 0.044 | 0.205 | 0 | 1 |
| Luxury district (1, 9, 10, 11) |  | 0.224 | 0.417 | 0 | 1 |
| Condominium |  | 0.733 | 0.442 | 0 | 1 |
| Floor number |  | 9.588 | 8.129 | 1 | 68 |
| Ground floor |  | 0.056 | 0.230 | 0 | 1 |
| Top floor |  | 0.080 | 0.271 | 0 | 1 |
| At least one Singaporean buyer | indicator | 0.809 | 0.393 | 0 | 1 |
| At least one Chinese buyer | indicator | 0.910 | 0.287 | 0 | 1 |
| At least one Cantonese buyer | indicator | 0.032 | 0.177 | 0 | 1 |
| At least one Hokkien buyer | indicator | 0.060 | 0.237 | 0 | 1 |
| At least one Teochew buyer | indicator | 0.016 | 0.126 | 0 | 1 |
| At least one buyer with english name | indicator | 0.305 | 0.461 | 0 | 1 |
| Mean buyer age > $=46$ |  | 0.244 | 0.429 | 0 | 1 |
| Mean buyer age $<=34$ |  | 0.303 | 0.460 | 0 | 1 |
| At least one buyer is senior public sector employee | indicator | 0.066 | 0.248 | 0 | 1 |
| At least one buyer with univ. degree | indicator | 0.211 | 0.408 | 0 | 1 |
| At least one buyer with science degree | indicator | 0.113 | 0.317 | 0 | 1 |
| With mortgage |  | 0.707 | 0.455 | 0 | 1 |
| Mean number of traffic accidents as plaintiff |  | 0.038 | 0.187 | 0 | 9 |
| Mean number of traffic accidents as defendant |  | 0.063 | 0.264 | 0 | 10 |

Notes: The sample comprises new sales of apartments between the years 2000-09. Transaction prices are denominated in 2014 Singapore dollars per square meter (at the average exchange rate of one Singapore dollar to 61 U.S. cents, a price of $\mathbf{S} \$ 1000$ per square meter is equivalent to US $\$ 56.75$ per square foot). Very lucky = last digit of both floor and apartment numbers is " 8 "; lucky = last digit of either floor or apartment number is " 8 " and neither last digit is " 4 "; unlucky $=$ last digit of either floor or apartment number is " 4 " and neither last digit is " 8 "; very unlucky $=$ last digit of both floor and apartment numbers is " 4 "; neutral $=$ all other addresses. Data on buyer's age, working for public sector, university degree, and numbers of car accidents is only available for Singaporean buyers.

Table 3. Lucky addresses: Prices

| VARIABLES | (a) | (b) | (c) | (d) | (e) |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | New sales |  |  |  |  | Subsales |
|  | 2000-2009 |  |  |  | 1995-2012 | and resales |
| Size (log) | $\begin{aligned} & -0.001 \\ & (0.041) \end{aligned}$ | $\begin{gathered} -0.053 \\ (0.041) \end{gathered}$ | $\begin{gathered} -0.111 * * * \\ (0.018) \end{gathered}$ | $\begin{gathered} -0.111 * * * \\ (0.018) \end{gathered}$ | $\begin{gathered} -0.124^{* * *} \\ (0.015) \end{gathered}$ | $\begin{gathered} -0.130 * * * \\ (0.017) \end{gathered}$ |
| Chinese buyers | $\begin{gathered} -0.131^{* * *} \\ (0.013) \end{gathered}$ | $\begin{gathered} -0.036 * * * \\ (0.007) \end{gathered}$ | $\begin{gathered} -0.014 * * * \\ (0.003) \end{gathered}$ |  | $\begin{gathered} -0.013 * * * \\ (0.003) \end{gathered}$ | $\begin{gathered} -0.024^{* * *} \\ (0.003) \end{gathered}$ |
| Non-Chinese buyers | $\begin{gathered} 0.015 * * * \\ (0.003) \end{gathered}$ |  |  |  |  |  |
| Very unlucky | $\begin{gathered} -0.040^{* * *} \\ (0.010) \end{gathered}$ | $\begin{gathered} -0.014^{*} * \\ (0.005) \end{gathered}$ | $\begin{gathered} -0.015 * * * \\ (0.004) \end{gathered}$ | $\begin{gathered} -0.012 * * * \\ (0.004) \end{gathered}$ | $\begin{gathered} -0.014 * * * \\ (0.004) \end{gathered}$ | $\begin{gathered} -0.016 * * * \\ (0.005) \end{gathered}$ |
| Unlucky | $\begin{gathered} -0.019^{* * *} \\ (0.004) \end{gathered}$ | $\begin{gathered} -0.009 * * * \\ (0.003) \end{gathered}$ | $\begin{gathered} -0.007 * * * \\ (0.002) \end{gathered}$ | $\begin{gathered} -0.007 * * * \\ (0.002) \end{gathered}$ | $\begin{gathered} -0.006 * * * \\ (0.002) \end{gathered}$ | $\begin{gathered} -0.006 * * * \\ (0.002) \end{gathered}$ |
| Lucky | $\begin{gathered} 0.012 * * \\ (0.005) \end{gathered}$ | $\begin{gathered} 0.008 * * * \\ (0.003) \end{gathered}$ | $\begin{gathered} 0.005 * * \\ (0.002) \end{gathered}$ | $\begin{gathered} 0.006 * * \\ (0.002) \end{gathered}$ | $\begin{gathered} 0.007 * * * \\ (0.002) \end{gathered}$ | $\begin{aligned} & 0.005^{*} \\ & (0.002) \end{aligned}$ |
| Very lucky | $\begin{gathered} 0.040 * * * \\ (0.015) \end{gathered}$ | $\begin{gathered} 0.030 * * * \\ (0.010) \end{gathered}$ | $\begin{gathered} 0.022^{*} * * \\ (0.007) \end{gathered}$ | $\begin{gathered} 0.024 * * * \\ (0.008) \end{gathered}$ | $\begin{gathered} 0.020 * * * \\ (0.006) \end{gathered}$ | $\begin{gathered} 0.020^{* *} \\ (0.009) \end{gathered}$ |
| Non-Chinese | (0.013) |  |  |  |  |  |
| Non-Chinese | -0.001 |  |  |  |  |  |
| x unlucky | (0.005) |  |  |  |  |  |
| Non-Chinese |  | -0.005 |  |  |  |  |
| x lucky |  | (0.007) |  |  |  |  |
| Non-Chinese |  | -0.054 |  |  |  |  |
| x very lucky |  | (0.042) |  |  |  |  |
| Constant | 9.130*** | $9.461 * * *$ | $9.718^{* * *}$ | $9.704^{* * *}$ | 9.462*** | 9.799*** |
|  | (0.190) | (0.198) | (0.092) | (0.091) | (0.091) | (0.087) |
| Lease tenure | Yes | Yes | Yes | Yes | Yes | Yes |
| Condominium | Yes | Yes | Yes | Yes | Yes | Yes |
| Building age (quadratic) | No | No | No | No | No | Yes |
| Floor controls | Yes | Yes | Yes | Yes | Yes | Yes |
| Postal district f.e. | No | Yes | Yes | Yes | Yes | Yes |
| Year f.e. | Yes | Yes | Yes | Yes | Yes | Yes |
| Quarter f.e. | Yes | Yes | Yes | Yes | Yes | Yes |
| Observations | 54,098 | 54,098 | 54,098 | 54,098 | 69,075 | 50,347 |
| R-squared | 0.553 | 0.843 | 0.928 | 0.928 | 0.936 | 0.886 |

Notes: Dependent variable is the logarithm of price in Singapore dollars per square meter; Floor controls include a linear trend in floor number, and indicators for first and top floors; Estimates are robust to controlling for floor number with a quadratic trend; Robust standard errors clustered by postal district ( ${ }^{*} \mathrm{p}<0.1 ;{ }^{* *} \mathrm{p}<0.05$; ${ }^{* * *} \mathrm{p}<0.01$ ). Column (a): New sales to all buyers with controls; Column (b): New sales to all buyers with controls and 2-digit postal district fixed effects; Column (b): New sales to all buyers with controls and 4-digit postal district fixed effects; Column (d): New sales distinguishing Chinese and non-Chinese buyers; Column (e): New sales, including earlier transactions back to 1995 and forward to 2012; Column (f): Secondary market (subsales and resales) 1995-2012, including quadratic control for age of building.

Table 4. Lucky addresses: Buyer choice

| VARIABLES | (a) | (b) | (c) |  | (e) | (f) | (g) | (h) | (i) Subsales and resales |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | New sales |  |  |  |  |  |  |  |  |
|  | All buyers | All buyers | All buyers (preferred) | All buyers | Luxury areas | Excl mixed addresses | Buildings at least 8 floors | $\begin{aligned} & 1995- \\ & 2012 \end{aligned}$ |  |
| Chinese buyers | $\begin{gathered} \hline 0.183 * * * \\ (0.033) \end{gathered}$ | $\begin{gathered} \hline 0.220^{* * *} \\ (0.032) \end{gathered}$ | $\begin{gathered} \hline 0.223 * * * \\ (0.034) \end{gathered}$ | $\begin{gathered} \hline 0.169 * * * \\ (0.049) \end{gathered}$ | $\begin{gathered} \hline 0.274 * * * \\ (0.040) \end{gathered}$ | $\begin{gathered} \hline 0.223 * * * \\ (0.034) \end{gathered}$ | $\begin{gathered} \hline 0.266 * * * \\ (0.036) \end{gathered}$ | $\begin{gathered} \hline 0.218^{* * *} \\ (0.029) \end{gathered}$ | $\begin{gathered} \hline 0.057 * * \\ (0.025) \end{gathered}$ |
| Singaporean buyers |  |  |  | $\begin{gathered} -0.070 \\ (0.063) \end{gathered}$ |  |  |  |  |  |
| Singaporean x Chinese |  |  |  | $\begin{gathered} 0.097 \\ (0.066) \end{gathered}$ |  |  |  |  |  |
| Chinese x luxury |  |  |  |  | $\begin{gathered} -0.135^{*} \\ (0.074) \end{gathered}$ |  |  |  |  |
| Size (log) | $\begin{gathered} 0.407 * * * \\ (0.058) \end{gathered}$ | $\begin{gathered} 0.442 * * * \\ (0.061) \end{gathered}$ | $\begin{gathered} 0.558 * * * \\ (0.076) \end{gathered}$ | $\begin{gathered} 0.558 * * * \\ (0.076) \end{gathered}$ | $\begin{gathered} 0.557 * * * \\ (0.076) \end{gathered}$ | $\begin{gathered} 0.564 * * * \\ (0.077) \end{gathered}$ | $\begin{gathered} 0.623^{* * *} \\ (0.096) \end{gathered}$ | $\begin{gathered} 0.456 * * * \\ (0.064) \end{gathered}$ | $\begin{gathered} 0.263 * * * \\ (0.067) \end{gathered}$ |
| Lease tenure | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Condominium | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Floor control | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Postal district f.e. | No | 2-digit | 4-digit | 4-digit | 4-digit | 4-digit | 4-digit | 4-digit | 4-digit |
| Year f.e. | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Quarter f.e. | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Building age (quadratic) | No | No | No | No | No | No | No | No | Yes |
| Observations | 54,098 | 54,098 | 54,098 | 54,098 | 54,098 | 53,132 | 44,562 | 69,075 | 50,347 |

Notes: Ordered logit model of demand for very lucky, lucky, neutral, unlucky, and very unlucky addresses, estimated by maximum likelihood; Ordered logit cut-points not reported; Floor controls include a linear trend in floor number, and indicators for first and top floors; Robust standard errors clustered by postal district except for column (a) (* p<0.1; ** $p<0.05 ; * * * p<0.01$ ). Column (a): New sales with controls; Column (b): New sales with controls and 2-digit postal district fixed effects; Column (c): New sales with controls and 4-digit postal district fixed effects (preferred estimate); Column (d): New sales, distinguishing between Singapore and non-Singaporean Chinese buyers; Column (e): New sales, luxury areas (Districts 1, 9, 10, and 11); Column (f): New sales, excluding mixed addresses such as 08-04; Column (g): New sales, buildings with 8 or more floors; Column (h): New sales, including earlier transactions back to 1995 and forward to 2012; Column (i): Secondary market (subsales and resales) 1995-2012, including quadratic control for age of building.

Table 5. Singaporean Chinese buyers: Demographics and economics

| VARIABLES | (a) <br> Dialect groups | (b) <br> Western name | $\begin{gathered} \text { (c) } \\ \text { Age } \end{gathered}$ | (d) <br> University graduates | (e) <br> Science graduates | (f) <br> Science / other | (g) Publicsector | (h) <br> Mort- <br> gage | (j) <br> Combined | (j) <br> Combined |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Cantonese | $\begin{gathered} 0.022 \\ (0.056) \end{gathered}$ |  |  |  |  |  |  |  | $\begin{gathered} \hline 0.020 \\ (0.058) \end{gathered}$ | $\begin{gathered} \hline 0.020 \\ (0.058) \end{gathered}$ |
| Hokkien | $\begin{gathered} 0.059 \\ (0.042) \end{gathered}$ |  |  |  |  |  |  |  | $\begin{gathered} 0.051 \\ (0.042) \end{gathered}$ | $\begin{gathered} 0.051 \\ (0.042) \end{gathered}$ |
| Teochew | $\begin{gathered} 0.086 \\ (0.074) \end{gathered}$ |  |  |  |  |  |  |  | $\begin{gathered} 0.073 \\ (0.076) \end{gathered}$ | $\begin{gathered} 0.073 \\ (0.076) \end{gathered}$ |
| Western name |  | $\begin{gathered} -0.049 * * \\ (0.022) \end{gathered}$ |  |  |  |  |  |  | $\begin{aligned} & -0.042^{*} \\ & (0.023) \end{aligned}$ | $\begin{aligned} & -0.040^{*} \\ & (0.022) \end{aligned}$ |
| Mean age of buyers >=46 |  |  | $\begin{gathered} 0.124^{*} * * \\ (0.028) \end{gathered}$ |  |  |  |  |  | $\begin{gathered} 0.110^{* * *} \\ (0.029) \end{gathered}$ | $\begin{gathered} 0.113 * * * \\ (0.029) \end{gathered}$ |
| Mean age of buyers < = 34 |  |  | $\begin{gathered} -0.032 \\ (0.028) \end{gathered}$ |  |  |  |  |  | $\begin{gathered} -0.025 \\ (0.028) \end{gathered}$ | $\begin{gathered} -0.027 \\ (0.028) \end{gathered}$ |
| University degree |  |  |  | $\begin{gathered} -0.034 \\ (0.024) \end{gathered}$ |  | $\begin{gathered} 0.005 \\ (0.036) \end{gathered}$ |  |  |  | $\begin{gathered} 0.033 \\ (0.035) \end{gathered}$ |
| University science degree |  |  |  |  | $\begin{gathered} -0.069 * * \\ (0.033) \end{gathered}$ | $\begin{gathered} -0.073 \\ (0.047) \end{gathered}$ |  |  | $\begin{gathered} -0.040 \\ (0.035) \end{gathered}$ | $\begin{gathered} -0.067 \\ (0.047) \end{gathered}$ |
| Public-sector employee |  |  |  |  |  |  | $\begin{gathered} -0.139 * * * \\ (0.039) \end{gathered}$ |  | $\begin{gathered} -0.134 * * * \\ (0.041) \end{gathered}$ | $\begin{gathered} -0.138 * * * \\ (0.042) \end{gathered}$ |
| Mortgage |  |  |  |  |  |  |  | $\begin{aligned} & -0.047 \\ & (0.030) \end{aligned}$ | $\begin{aligned} & -0.018 \\ & (0.030) \end{aligned}$ | $\begin{gathered} -0.019 \\ (0.030) \end{gathered}$ |
| Apartment characteristics | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| 4-digit postal district f.e. | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Year f.e. | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Quarter f.e. | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Observations | 39,789 | 39,789 | 37,203 | 39,789 | 39,789 | 39,789 | 39,789 | 39,789 | 37,203 | 37,203 |

Notes: Singaporean Chinese buyers; Ordered logit model of demand for very lucky, lucky, neutral, unlucky, and very unlucky addresses, estimated by maximum likelihood; Apartment characteristics include size, floor, legal tenure, and condominium; Ordered logit cut-points not reported; Robust standard errors clustered by postal district (* p<0.1; ** p<0.05; *** p<0.01). Column (a): Including proportion of buyers by dialect groups; Column (b): Including proportion of buyers with Western given name; Column (c): Including buyers by mean age; Column (d): Including proportion of buyers who are university graduates; Column (e): Including proportion of buyers who are university graduates in science or technology disciplines; Column (f): Including proportions of buyers who are university graduates and buyers who are science or technology graduates; Column (g): Including proportion of buyers who are employed in the public sector; Column (h): Including indicator of purchases financed by mortgage loans; Column (i): Including all demographic and economic covariates except university graduate; Column (i): Including all demographic and economic covariates.

Table 6. Singaporean Chinese buyers: Misfortune and uncertainty

| VARIABLES | (a) <br> Traffic <br> accidents <br> plaintiff | (b) <br> Traffic <br> accidents <br> defendant | (c) <br> SARS | (d) <br> Combined |
| :--- | :---: | :---: | :---: | :---: |
| Mean number of traffic | $0.102^{* *}$ |  |  | $0.099^{* *}$ |
| accidents as plaintiff | $(0.042)$ |  |  | $(0.043)$ |
| Mean number of traffic |  | 0.023 |  | 0.012 |
| accidents as defendant |  | $(0.038)$ |  | $(0.038)$ |
| SARS |  |  | -0.002 | -0.005 |
|  |  |  | $(0.133)$ | $(0.132)$ |
| Apartment characteristics | Yes | Yes | Yes | Yes |
| 4-digit postal district f.e. | Yes | Yes | Yes | Yes |
| Year f.e. | Yes | Yes | Yes | Yes |
| Quarter f.e. | Yes | Yes | Yes | Yes |
| Observations | 41,568 | 41,568 | 41,568 | 41,568 |

Notes: Singaporean Chinese buyers; Ordered logit model of demand for very lucky, lucky, neutral, unlucky, and very unlucky addresses, estimated by maximum likelihood; Apartment characteristics include size, floor, legal tenure, and condominium; Ordered logit cut-points not reported; Robust standard errors clustered by 4 -digit postal district ( $* \mathrm{p}<0.1 ; * * \mathrm{p}<0.05$; *** $\mathrm{p}<0.01$ ). Column (a): Including mean number of traffic accidents buyer was involved as plaintiff; Column (b): Including mean number of traffic accidents buyer was involved as defendant; Column (c): SARS (March to June 2003);

Figure 1. Transactions by year


Note: Number of new sales, by year from 1995 to 2012. Solid red line depicts the number of new sales as recorded in the Urban Redevelopment Authority's REALIS database. Broken blue line depicts the number of new sales as recorded in our data-set of legal filings with the Registry of Land Titles.

Figure 2. Price and floor number


Note: The figure plots the estimated coefficients of the floor levels in a regression of the logarithm of unit price per square meter on indicators of floor number and last digit of unit number, apartment size, and building fixed effects.

Figure 3. Price and unit number


Notes: The Figure plots the estimated coefficients of the last digit of the unit number in a regression of the logarithm of price per square meter on indicators of the last digit of the unit number.

Table S1. Lucky addresses: Prices -- Additional estimates

| VARIABLES | (a) | (b) |
| :---: | :---: | :---: |
|  | 6-digit postal code fixed effects | Singaporean buyers |
| Size (log) | -0.150*** | -0.132*** |
|  | (0.011) | (0.020) |
| Chinese buyers | -0.008*** |  |
|  | (0.002) |  |
| Non-Chinese buyers |  | 0.007* |
|  |  | (0.004) |
| Very unlucky | $-0.014^{* * *}$ | $-0.011^{* *}$ |
|  | (0.003) | (0.005) |
| Unlucky | -0.004*** | $-0.008 * * *$ |
|  | (0.002) | (0.002) |
| Lucky | 0.006*** | 0.007*** |
|  | (0.002) | (0.002) |
| Very lucky | 0.019*** | $0.027 * * *$ |
|  | (0.007) | (0.007) |
| Non-Chinese x very unlucky |  | -0.031* |
|  |  | (0.018) |
| Non-Chinese x unlucky |  | -0.010 |
|  |  | (0.007) |
| Non-Chinese x lucky |  | -0.004 |
|  |  | (0.010) |
| Non-Chinese x very lucky |  | -0.079 |
|  |  | (0.086) |
| Constant | 9.741*** | 9.782*** |
|  | (0.056) | (0.098) |
| Lease tenure | Yes | Yes |
| Condominium | Yes | Yes |
| Floor controls | Yes | Yes |
| Postal district fixed effects | 6-digit | 4-digit |
| Year fixed effects | Yes | Yes |
| Quarter fixed effects | Yes | Yes |
| Observations | 54,098 | 43,784 |
| R-squared | 0.969 | 0.926 |

Notes: Dependent variable is the logarithm of price in Singapore dollars per square meter; Robust standard errors clustered by postal districts ( ${ }^{*} \mathrm{p}<0.1 ; * * \mathrm{p}<0.05$; *** $\mathrm{p}<0.01$ ). Floor controls include a linear trend in floor number, and indicators for first and top floors; Estimates are robust to controlling for floor number with a quadratic trend.

Table S2. Annualized rate of return

|  | (a) <br> All | (b) <br> Trim top <br> \& bottom $1 \%$ | (c) <br> Trim top <br> \& bottom $1 \%$ |
| :--- | :---: | :---: | :---: |
| VARIABLES | -0.002 | -0.004 | -0.002 |
| Very unlucky | $(0.008)$ | $(0.007)$ | $(0.005)$ |
|  | -0.002 | -0.000 | 0.000 |
| Unlucky | $(0.002)$ | $(0.002)$ | $(0.001)$ |
|  | 0.001 | -0.002 | $-0.003^{* *}$ |
| Lucky | $(0.003)$ | $(0.002)$ | $(0.001)$ |
|  | 0.000 | -0.007 | -0.004 |
| Very lucky | $(0.011)$ | $(0.005)$ | $(0.005)$ |
|  | $0.176^{* * *}$ | $0.181^{* * *}$ | $0.153^{* * *}$ |
| Constant | $(0.004)$ | $(0.004)$ | $(0.014)$ |
|  | No | No | Yes |
| Apartment characteristics | Yes | Yes | Yes |
| Year of new sale f.e. | Yes | Yes | Yes |
| Quarter of new sale f.e. | Yes | Yes | Yes |
| Year of resale f.e. | Yes | Yes | Yes |
| Quarter of resale f.e. | 5,284 | 5,126 | 5,126 |
| Observations | 0.465 | 0.476 | 0.702 |
| R-squared | .032 | .028 | .028 |
| Mean of dependent |  |  |  |

Notes: Sample comprises apartments with both new sale and resale transactions between 2000 and 2009;
Dependent variable is annualized rate of return; Apartment characteristics include size, floor, legal tenure, condominium, and 4-digit postal code; Robust standard errors clustered by 4 -digit postal district ( ${ }^{*} \mathrm{p}<0.1$; ** $\mathrm{p}<0.05$; *** $\mathrm{p}<0.01$ ).

Table S3: Uncertainty: Additional estimates

|  | (a) | (b) |
| :--- | :---: | :---: |
| VARIABLES | GDP | Consumption |
| Gross Domestic Product | $(0.071$ |  |
| forecast uncertainty |  | -0.044 |
| Private consumption |  | $(0.051)$ |
| forecast uncertainty | Yes | Yes |
| Apartment characteristics | Yes | Yes |
| 4-digit postal district fixed effects | Yes | Yes |
| Year f.e. | Yes | Yes |
| Quarter f.e. | 41,568 | 41,568 |
| Observations |  |  |

Notes: Singaporean Chinese buyers; Ordered logit model of demand for very lucky, lucky, neutral, unlucky, and very unlucky addresses, estimated by maximum likelihood; Apartment characteristics include size, floor, legal tenure, and condominium; Ordered logit cut-points not reported; Robust standard errors clustered by 4 -digit postal district ( $* \mathrm{p}<0.1$; ** $\mathrm{p}<0.05$; *** $\mathrm{p}<0.01$ ). Column (a): Gross Domestic Product forecast uncertainty is the standard deviation of analysts' predicted rate of GDP growth; Column (b): Private consumption forecast uncertainty is the standard deviation of analysts' predicted rate of growth of private consumption. Source of forecasts: Monetary Authority of Singapore.


[^0]:    * This version supersedes earlier versions entitled, "Can Superstitious Beliefs Affect Market Equilibrium? Personal Beliefs and Beliefs about Others" and "Superstition and Asset Markets: Evidence from Singapore Housing". The authors thank Gene Amromin, Zahi Ben-David, Souphala Chomsisengphet, David Hirshleifer, Teck-hua Ho, Amit Seru, Wenlan Qian, and participants at the Econometric Society Asian and Australasian Meetings, and seminars at the Nanyang Technological University, and National University of Singapore for comments, Ng Hwee Tou for supervising and Harta Wijaya for building the ethnic classification system, and William Lai and Christopher Ng Gon Chew for providing data, and the NUS Institute of Real Estate Studies for research funding.

[^1]:    ${ }^{1}$ Houses numbered " 4 " were older, while those numbered " 8 " were newer and had better features. The difference in age might be related to people buying houses numbered " 4 ", rebuilding and then, petitioning to renumber. In neighborhoods with more Chinese, house numbers ending in " 0 ", " 5 ", and " 8 " were more frequent (Fortin et al. 2014: Figure 3).
    ${ }^{2}$ Apparently, the Chinese vary in their definition of which numbers are lucky. Shum et al. (2014) defined both " 6 " and " 8 " as lucky, while Hirshleifer et al. (2012) considered " 6 ", " 8 ", and " 9 " to be lucky. The common thread in the previous research is that " 8 " is lucky and " 4 " is unlucky.

[^2]:    ${ }^{3}$ These and the following statistics of Singapore's population and housing market are drawn from the Yearbook of Statistics 2013 and Census of Population 2010, Department of Statistics, Singapore.
    ${ }^{4}$ A unique feature of Singapore's public housing is an explicit policy to ensure ethnic diversity (Wong 2013). The Housing and Development Board regulates the resale of apartments by building and area to prevent the percentage of any particular race from exceeding specified benchmarks.

[^3]:    ${ }^{5}$ By contrast, the authorities in other countries are more flexible. In the United States, many highrise hotels lack a 13th floor (USA Today, 2007). The Hong Kong authorities are particularly accommodating of real estate developers in floor numbering. The sales brochure for the Arch, a development in West Kowloon, boldly advertised that " $[t]$ here are no $14,24,34,40-49,53-54,58,60-$ $61,64 \& 74 \mathrm{~F}$ ". In the absence of units with unlucky numbers, any research would be limited to investigating the effects of lucky numbers.

[^4]:    ${ }^{6}$ Condominiums are high-rise buildings with exclusive access to amenities such as parking, sports facilities, and children's playgrounds. Apartment buildings are high-rise buildings without such exclusive amenities. Landed properties are single-family homes, and have unique street addresses and 6 -digit postal codes. Freehold is an unlimited tenure of ownership, while leasehold is a limited tenure of ownership, typically 99 years, after which ownership reverts to the government.

[^5]:    ${ }^{7}$ We limit analysis of the secondary market transactions (subsales and resales) to those for which we have data on the date of legal occupancy, from which we can calculate the age of the building. Age is surely an important factor in the price of apartments.

[^6]:    ${ }^{8}$ The insignificance of the interaction terms might seem to be inconsistent with Fortin et al. (2014), who find that, in the Greater Vancouver area, the discount on unlucky addresses and the premium for lucky addresses depend on the proportion of Chinese residents in the neighborhood. The difference is likely due to the Chinese being the majority throughout Singapore, and also being the marginal buyers and hence determining market prices. When numerological preference are prevalent, non-believers must take the future resale value of their residential investment into consideration even if they themselves have no numerological preference. Furthermore, the valuation of Chinese buyers is likely to appear as a salient opportunity cost to sellers.

[^7]:    ${ }^{9}$ We do not further refine the controls to the building level (6-digit postal district) as the estimation drops many observations due to perfect prediction, and the large number of dummy variables would very likely present the problem of incidental parameters in estimating the ordered logit model. The slight difference between the estimates controlling for 2 vis-à-vis 4 -digit postal districts (Table 4, columns (b) and (c)) suggest that the estimated coefficients are not sensitive to further refinement of the controls.

[^8]:    ${ }^{10}$ We classify given names that can be either Chinese or Western, such as "Lee", as Chinese.

