IZA DP No. 4242

Forecasting with Spatial Panel Data

Badi H. Baltagi Georges Bresson Alain Pirotte

June 2009

Forschungsinstitut zur Zukunft der Arbeit Institute for the Study of Labor

Forecasting with Spatial Panel Data

Badi H. Baltagi

Syracuse University and IZA

Georges Bresson

ERMES (CNRS), Université Panthéon-Assas Paris II

Alain Pirotte

ERMES (CNRS), Université Panthéon-Assas Paris II and INRETS-DEST

> Discussion Paper No. 4242 June 2009

> > IZA

P.O. Box 7240 53072 Bonn Germany

Phone: +49-228-3894-0 Fax: +49-228-3894-180 E-mail: iza@iza.org

Any opinions expressed here are those of the author(s) and not those of IZA. Research published in this series may include views on policy, but the institute itself takes no institutional policy positions.

The Institute for the Study of Labor (IZA) in Bonn is a local and virtual international research center and a place of communication between science, politics and business. IZA is an independent nonprofit organization supported by Deutsche Post Foundation. The center is associated with the University of Bonn and offers a stimulating research environment through its international network, workshops and conferences, data service, project support, research visits and doctoral program. IZA engages in (i) original and internationally competitive research in all fields of labor economics, (ii) development of policy concepts, and (iii) dissemination of research results and concepts to the interested public.

IZA Discussion Papers often represent preliminary work and are circulated to encourage discussion. Citation of such a paper should account for its provisional character. A revised version may be available directly from the author.

IZA Discussion Paper No. 4242 June 2009

ABSTRACT

Forecasting with Spatial Panel Data^{*}

This paper compares various forecasts using panel data with spatial error correlation. The true data generating process is assumed to be a simple error component regression model with spatial remainder disturbances of the autoregressive or moving average type. The best linear unbiased predictor is compared with other forecasts ignoring spatial correlation, or ignoring heterogeneity due to the individual effects, using Monte Carlo experiments. In addition, we check the performance of these forecasts under misspecification of the spatial error process, various spatial weight matrices, and heterogeneous rather than homogeneous panel data models.

JEL Classification: C33

Keywords: forecasting, BLUP, panel data, spatial dependence, heterogeneity

Corresponding author:

Badi H. Baltagi Department of Economics and Center for Policy Research 426 Eggers Hall Syracuse University Syracuse, NY 13244-1020 USA E-mail: bbaltagi@maxwell.syr.edu

^{*} This paper was presented at a conference in honor of Phoebus Dhrymes in Paphos, Cyprus, June 1-3, 2007. Also, at the 14th International Conference on Panel Data at the Wang Yanan Institute for Studies in Economics (WISE), Xiamen University, China, July 16-18, 2007, and the 63rd European Meeting of the Econometric Society (ESEM) held at the University of Bocconi in Milan, Italy, August 27-31, 2008.

1 Introduction

The literature on forecasting is rich with time series applications, but this is not the case for spatial panel data applications. Exceptions are Baltagi and Li (2004, 2006) with applications to forecasting sales of cigarettes and liquor per capita for U.S. states over time.¹ Best linear unbiased prediction (BLUP) in panel data using an error component model have been considered by Taub (1979), Baltagi and Li (1992), and Baillie and Baltagi (1999) to mention a few. Applications include Baltagi and Griffin (1997), Hsiao and Tahmiscioglu (1997), Schmalensee, Stoker and Judson (1998), Baltagi, Griffin and Xiong (2000), Hoogstrate, Palm and Pfann (2000), Baltagi, Bresson and Pirotte (2002, 2004), Frees and Miller (2004), Rapach and Wohar (2004), and Brucker and Siliverstovs (2006), see Baltagi (2008) for a recent survey. However, these panel forecasting applications do not deal with spatial dependence across the panel units. Spatial dependence models — popular in regional science and urban economics — deal with spatial interaction and spatial heterogeneity (see Anselin (1988) and Anselin and Bera (1998)). The structure of the dependence can be related to location and distance, both in a geographic space as well as a more general economic or social network space. Some commonly used spatial error processes include the spatial autoregressive (SAR) and the spatial moving average (SMA) error processes. Two different variants of these models for spatial panels are considered, one discussed in Anselin (1988) and another in Kapoor, Kelejian and Prucha (2007) and Fingleton (2007). The best linear unbiased predictors for the Anselin type model was derived by Baltagi and Li (2004). This paper derives the best linear unbiased predictors for the Kapoor, Kelejian and Prucha (2007) and Fingleton (2007) variants. More importantly, it compares the performance of sixteen various forecasts of the spatial panel data using Monte Carlo experiments. These include homogeneous as well as heterogeneous estimators of the spatial panel model and their corresponding forecasts. The true data generating process is assumed to be a simple error component regression model with spatial remainder disturbances of the autoregressive or moving average

¹In order to explain how spatial autocorrelation may arise in the demand for cigarettes, we note that cigarette prices vary among states primarily due to variation in state taxes on cigarettes. Border effect purchases not included in the cigarette demand equation can cause spatial autocorrelation among the disturbances. In forecasting sales of cigarettes, the spatial autocorrelation due to neighboring states and the individual heterogeneity across states is taken explicitly into account.

type. The best linear unbiased predictor is compared with other forecasts ignoring spatial correlation, or ignoring heterogeneity due to the individual effects. In addition, we check the performance of these forecasts under misspecification of the spatial error process, different spatial weight matrices, and various sample sizes. Section 2 introduces the error component model with spatially autocorrelated residuals of the SAR and SMA type. Section 3 describes the forecasts using the estimators considered in Section 2, while Section 4 gives the Monte Carlo design. Section 5 reports the results of the Monte Carlo simulations and Section 6 gives our summary and conclusion.

2 The Error Component Model with Spatially Autocorrelated Residuals

Consider a linear panel data regression model:

$$y_{it} = X_{it}\beta + \varepsilon_{it} , i = 1, ..., N; t = 1, ..., T$$
 (1)

where the disturbance term follows an error component model with spatially autocorrelated residuals. The disturbance vector for time t is given by:

$$\varepsilon_t = \mu + \phi_t \tag{2}$$

where $\varepsilon_t = (\varepsilon_{1t}, ..., \varepsilon_{Nt})'$, $\mu = (\mu_1, ..., \mu_N)'$ denotes the vector of specific effects assumed to be *iid* $(0, \sigma_{\mu}^2)$ and $\phi_t = (\phi_{1t}, ..., \phi_{Nt})'$ are the remainder disturbances which are independent of μ . We let the ϕ_t 's follow a spatial autoregressive (SAR) or a spatial moving average (SMA) error model. The SAR process is known to transmit the shocks globally while the SMA process transmits these shocks locally, see Anselin, Le Gallo and Jayet (2008).

The SAR specification for the $(N \times 1)$ error vector ϕ_t at time t can be expressed as:

$$\phi_t = \rho W_N \phi_t + v_t = (I_N - \rho W_N)^{-1} v_t = B_N^{-1} v_t$$
(3)

where W_N is an $(N \times N)$ known spatial weights matrix², ρ is the spatial autoregressive parameter and v_t is an $(N \times 1)$ error vector assumed to be dis-

²In the simplest case, the weights matrix is binary, with $w_{ij} = 1$ when *i* and *j* are neighbors and $w_{ij} = 0$ when they are not. By convention, diagonal elements are null: $w_{ii} = 0$ and the weights are almost always standardized such that the elements of each row sum to 1.

tributed independently across cross-sectional dimension with constant variance $\sigma_v^2 I_N$. $B_N = (I_N - \rho W_N)$ and is assumed to be non-singular. The error covariance matrix for the cross-section at time t becomes:

$$\Omega_t = E\left[\varepsilon_t \varepsilon_t'\right] = \sigma_\mu^2 I_N + \sigma_v^2 \left(B_N' B_N\right)^{-1} \tag{4}$$

For the full $(NT \times 1)$ vector of disturbances:

$$\varepsilon = (\iota_T \otimes I_N) \, \mu + \left(I_T \otimes B_N^{-1} \right) v \tag{5}$$

the corresponding $(NT \times NT)$ covariance matrix is given by:

$$\Omega = \sigma_{\mu}^{2} \left(J_{T} \otimes I_{N} \right) + \sigma_{v}^{2} \left[I_{T} \otimes \left(B_{N}^{\prime} B_{N} \right)^{-1} \right]$$
(6)

where ι_T is a $(T \times 1)$ vector of ones and $J_T = \iota_T \iota'_T$ is a $(T \times T)$ matrix of ones.

The spatial moving average (SMA) specification for the $(N \times 1)$ error vector ϕ_t at time t can be expressed as:

$$\phi_t = \lambda W_N v_t + v_t = (I_N + \lambda W_N) v_t = D_N v_t \tag{7}$$

where $D_N = (I_N + \lambda W_N)$. The error covariance matrix for the cross-section at time t becomes:

$$\Omega_t = E\left[\varepsilon_t \varepsilon_t'\right] = \sigma_\mu^2 I_N + \sigma_v^2 \left(D_N D_N'\right) \tag{8}$$

For the full $(NT \times 1)$ vector of disturbances:

$$\varepsilon = (\iota_T \otimes I_N) \,\mu + (I_T \otimes D_N) \,v \tag{9}$$

the corresponding $(NT \times NT)$ covariance matrix is given by:

$$\Omega = \sigma_{\mu}^{2} \left(J_{T} \otimes I_{N} \right) + \sigma_{v}^{2} \left[I_{T} \otimes \left(D_{N} D_{N}^{\prime} \right) \right]$$
(10)

MLE under normality of the disturbances using these error component models with spatial autocorrelation have been derived by Anselin (1988). The log-likelihood is given by:

$$L \propto -\frac{NT}{2} \ln \left(2\pi\sigma_v^2\right) - \frac{1}{2} \ln |\Sigma| - \frac{1}{2\sigma_v^2} \varepsilon' \Sigma^{-1} \varepsilon$$
(11)

where

$$\varepsilon = y - X\beta , \ \Omega = \sigma_v^2 \Sigma$$

$$\Sigma = \begin{cases} (J_T \otimes \theta I_N) + [I_T \otimes (B'_N B_N)^{-1}] & \text{for SAR} \\ (J_T \otimes \theta I_N) + [I_T \otimes (D_N D'_N)] & \text{for SMA} \end{cases}$$
(12)
with $\theta = \sigma_\mu^2 / \sigma_v^2$.

Regression models containing spatially correlated disturbance terms based on the SAR or SMA models are typically estimated using MLE, where the likelihood function corresponds to the normal distribution. However, this can be computationally demanding for large N. Kelejian and Prucha (1999) suggested a generalized moments (GM) estimation method for the SAR model in a cross-section setting, and Fingleton (2007) extended this generalized moments estimator to the SMA model. Kapoor, Kelejian and Prucha (2007) generalized this GM procedure from cross-section to panel data and derived its large sample properties when T is fixed and $N \to \infty$. However, their SAR random effects model (SAR-RE) differs from that described in (2) which we will call (RE-SAR). In fact, in their specification, the disturbance term ε_t itself follows a SAR process and the remainder term follows an error component structure. This allows the individual effects, i.e., the μ 's themselves to be spatially correlated but with the same ρ . In particular, the disturbance vector for time t is given by:

$$\varepsilon_t = \rho W_N \varepsilon_t + u_t \tag{13}$$

where u_t follows an error component structure :

$$u_t = \mu + v_t \tag{14}$$

The SAR-RE specification for the $(N \times 1)$ error vector ε_t at time t can be expressed as:

$$\varepsilon_t = (I_N - \rho W_N)^{-1} u_t = B_N^{-1} u_t \tag{15}$$

where $B_N = (I_N - \rho W_N)$. For the full $(NT \times 1)$ vector of disturbances:

$$\varepsilon = \left(\iota_T \otimes B_N^{-1}\right)\mu + \left(I_T \otimes B_N^{-1}\right)v \tag{16}$$

and the corresponding $(NT \times NT)$ covariance matrix is given by:

$$\Omega = \sigma_{\mu}^{2} \left(J_{T} \otimes \left(B_{N}^{\prime} B_{N} \right)^{-1} \right) + \sigma_{v}^{2} \left[I_{T} \otimes \left(B_{N}^{\prime} B_{N} \right)^{-1} \right]$$
(17)

Kapoor, et al. (2007) proposed three generalized moments (GM) estimators of ρ , σ_v^2 and $\sigma_1^2 \left(=\sigma_v^2 + T\sigma_\mu^2\right)$ based on the following six moment conditions:

$$E\begin{bmatrix}\frac{\frac{1}{N(T-1)}u'_{N}Q_{0,N}u_{N}}{\frac{1}{N(T-1)}\overline{u}'_{N}Q_{0,N}\overline{u}_{N}}\\\frac{\frac{1}{N(T-1)}\overline{u}'_{N}Q_{0,N}u_{N}}{\frac{1}{N}u'_{N}Q_{1,N}u_{N}}\\\frac{\frac{1}{N}\overline{u}'_{N}Q_{1,N}\overline{u}_{N}}{\frac{1}{N}\overline{u}'_{N}Q_{1,N}u_{N}}\end{bmatrix} = \begin{bmatrix}\sigma_{v}^{2}\\\sigma_{v}^{2}\frac{1}{N}tr\left(W'_{N}W_{N}\right)\\0\\\sigma_{1}^{2}\\\sigma_{1}^{2}\frac{1}{N}tr\left(W'_{N}W_{N}\right)\\0\end{bmatrix}$$
(18)

where

$$u_N = \varepsilon_N - \rho \overline{\varepsilon}_N \tag{19}$$

$$\overline{u}_N = \overline{\varepsilon}_N - \rho \overline{\overline{\varepsilon}}_N \tag{20}$$

$$\overline{\varepsilon}_N = (I_T \otimes W_N) \varepsilon_N \tag{21}$$

$$\overline{\overline{\varepsilon}}_N = (I_T \otimes W_N) \overline{\varepsilon}_N \tag{22}$$

$$Q_{0,N} = \left(I_T - \frac{J_T}{T}\right) \otimes I_N \tag{23}$$

$$Q_{1,N} = \frac{J_T}{T} \otimes I_N \tag{24}$$

Under the random effects specification considered, the OLS estimator of β is consistent. Using $\hat{\beta}_{OLS}$ one gets a consistent estimator of the disturbances $\hat{\varepsilon} = y - X \hat{\beta}_{OLS}$. The GM estimators of σ_1^2 , σ_{ν}^2 and ρ are the solution of the sample counterpart of the six equations given above. Kapoor, et al. (2007) suggest three GM estimators. The first involves only the first three moments which do not involve σ_1^2 and yield estimates of ρ and σ_{ν}^2 . The fourth moment condition is then used to solve for σ_1^2 given estimates of ρ and σ_{ν}^2 . The second GM estimator is based upon weighing the moment equations by the inverse of a properly normalized variance-covariance matrix of the sample moments evaluated at the true parameter values. A simple version of this weighting matrix is derived under normality of the disturbances. The third GM estimator is motivated by computational considerations and replaces a component of the weighting matrix for the second GM estimator by an identity matrix. Kapoor, et al. (2007) perform Monte Carlo experiments comparing MLE and these three GM estimation methods. They find that on average, the RMSE of MLE and their weighted GM estimators are quite

similar. The feasible GLS estimator of β is then obtained by replacing ρ , σ_v^2 and σ_1^2 by their GM estimators.³

Recently, Fingleton (2007) extended this GM estimator for the SMA panel data model with random effects. We call this SMA-RE to distinguish it from the RE-SMA procedure described in Anselin, et al. (2008). In fact, for the Fingleton (2007) SMA-RE, the disturbance term ε_t in (2) follows a SMA process and the remainder term follows an error component structure. Unlike the Anselin, et al. (2008) RE-SMA, the individual effects, i.e., the μ 's themselves are allowed to be spatially correlated but with the same λ . In particular, the disturbance vector for time t is given by:

$$\varepsilon_t = (I_N + \lambda W_N) u_t = D_N u_t \tag{25}$$

where $D_N = (I_N + \lambda W_N)$, and u_t follows an error component structure (14). So, the full SMA-RE $(NT \times 1)$ vector of disturbances is given by:

$$\varepsilon = (\iota_T \otimes D_N) \, \mu + (I_T \otimes D_N) \, v \tag{26}$$

and the corresponding $(NT \times NT)$ covariance matrix is given by:

$$\Omega = \sigma_{\mu}^2 \left(J_T \otimes (D_N D'_N) \right) + \sigma_v^2 \left[I_T \otimes (D_N D'_N) \right]$$
(27)

The moment conditions for SMA-RE are similar to those derived by Kapoor, et al. (2007), see Fingleton (2007).

3 Prediction

Goldberger (1962) has shown that, for a given Ω , the best linear unbiased predictor (BLUP) for the *i*th individual at a future period $T + \tau$ is given by:

$$\widehat{y}_{i,T+\tau} = X_{i,T+\tau}\widehat{\beta}_{GLS} + \omega' \Omega^{-1}\widehat{\varepsilon}_{GLS}$$
(28)

where $\omega = E\left[\varepsilon_{i,T+\tau}\varepsilon\right]$ is the covariance between the future disturbance $\varepsilon_{i,T+\tau}$ and the sample disturbances ε . $\hat{\beta}_{GLS}$ is the GLS estimator of β from equation (1) based on Ω and $\hat{\varepsilon}_{GLS}$ denotes the corresponding GLS residual vector.

³Later, in our Monte Carlo experiments, we computed the predictors for all three GM estimators suggested by Kapoor, et al. (2007). However, the differences in root mean squared error performance were minor. To save space, we only report the second GM estimator, called weighted GM estimator by Kapoor, et al. (2007).

For the error component without spatial autocorrelation ($\lambda = 0$), this BLUP reduces to:

$$\widehat{y}_{i,T+\tau} = X_{i,T+\tau} \widehat{\beta}_{GLS} + \frac{\sigma_{\mu}^2}{\sigma_1^2} \left(\iota_T' \otimes l_i' \right) \widehat{\varepsilon}_{GLS}$$
⁽²⁹⁾

where $\sigma_1^2 = T\sigma_{\mu}^2 + \sigma_v^2$ and l_i is the *i*th column of I_N . This predictor was considered by Wansbeek and Kapteyn (1978), Lee and Griffiths (1979) and Taub (1979). The typical element of the last term of equation (29) is $(T\sigma_{\mu}^2/\sigma_1^2) \overline{\varepsilon}_{i,GLS}$ where $\overline{\varepsilon}_{i,GLS} = \sum_{t=1}^T \widehat{\varepsilon}_{ti,GLS}/T$. Therefore, the BLUP of $y_{i,T+\tau}$ for the RE model modifies the usual GLS forecasts by adding a fraction of the mean of the GLS residuals corresponding to the *i*th individual. In order to make this forecast operational, $\widehat{\beta}_{GLS}$ is replaced by its feasible GLS estimate and the variance components are replaced by their feasible estimates.

Baltagi and Li (2004, 2006) derived the BLUP correction term when both error components and spatial autocorrelation are present and ϕ_t follows a SAR process. So, the predictors for the SAR and the SMA are given by:

$$\widehat{y}_{i,T+\tau} = \begin{cases}
X_{i,T+\tau} \widehat{\beta}_{MLE} + \theta \left(\iota'_T \otimes l'_i C_1^{-1} \right) \widehat{\varepsilon}_{MLE} \\
= X_{i,T+\tau} \widehat{\beta}_{MLE} + T \theta \sum_{j=1}^N c_{1,j} \overline{\varepsilon}_{j,MLE} & \text{for SAR} \\
X_{i,T+\tau} \widehat{\beta}_{MLE} + \theta \left(\iota'_T \otimes l'_i C_2^{-1} \right) \widehat{\varepsilon}_{MLE} \\
= X_{i,T+\tau} \widehat{\beta}_{MLE} + T \theta \sum_{j=1}^N c_{2,j} \overline{\varepsilon}_{j,MLE} & \text{for SMA}
\end{cases}$$
(30)

where c_{1j} (resp. $c_{2,j}$) is the *j*th element of the *i*th row of C_1^{-1} (resp. C_2^{-1}) with $C_1 = [T\theta I_N + (B'_N B_N)^{-1}]$ (resp. $C_2 = [T\theta I_N + (D_N D'_N)]$) and $\overline{\varepsilon}_{j,,MLE} = \sum_{t=1}^T \widehat{\varepsilon}_{tj,MLE}/T$. In other words, the BLUP of $y_{i,T+\tau}$ adds to $X_{i,T+\tau}\widehat{\beta}_{MLE}$ a weighted average of the MLE residuals for the *N* individuals averaged over time. The weights depend upon the spatial matrix W_N and the spatial autoregressive (or moving average) coefficients ρ and λ . To make these predictors operational, we replace θ, ρ and λ by their estimates from the RE-spatial MLE with SAR or SMA. When there are no random individual effects, so that $\sigma_{\mu}^2 = 0$, then $\theta = 0$ and the BLUP prediction terms drop out completely from equation (30). In these cases, Ω in equation (12) reduces to $\sigma_v^2 [I_T \otimes (B'_N B_N)^{-1}]$ for SAR and $\sigma_v^2 [I_T \otimes (D_N D'_N)]$ for SMA, and the corresponding MLE for these models yield the pooled spatial MLE with SAR or SMA remainder disturbances.

For the Kapoor, et al. (2007) model, the BLUP of $y_{i,T+\tau}$ for the SAR-RE also modifies the usual GLS forecasts by adding a fraction of the mean of the GLS residuals corresponding to the *i*th individual. More specifically, the predictor is given by:

$$\widehat{y}_{i,T+\tau} = X_{i,T+\tau} \widehat{\beta}_{GLS} + \left(\frac{\sigma_{\mu}^2}{\sigma_1^2}\right) b_i \left(\iota_T' \otimes B_N\right) \widehat{\varepsilon}_{GLS}$$
(31)

where b_i is the *i*th row of the matrix B_N^{-1} . This is derived in the Appendix of this paper which also shows the resulting predictor has the same form as that of the RE model (29). This proof applies to both the Kapoor, et al. (2007) SAR-RE specification and the Fingleton (2007) SMA-RE specification. Therefore, the BLUP of $y_{i,T+\tau}$ for the SAR-RE and the SMA-RE, like the usual RE model with no spatial effects, modifies the usual GLS forecasts by adding a fraction of the mean of the GLS residuals corresponding to the *i*th individual. While the predictor formula is the same, the MLEs for these specifications yield different estimates which in turn yield different residuals and hence different forecasts.

4 Monte Carlo Design

In this section, we consider the small sample performance of several predictors for an error component model with spatially autocorrelated residuals. The data generating process (DGP) consider two specifications on the remainder errors, namely SAR and SMA:

$$y_{it} = \beta_0 + \beta_1 x_{it} + \varepsilon_{it} , \ \varepsilon_{it} = \mu_i + \phi_{it}, \ i = 1, ..., N; \ t = 1, ..., T$$
(32)

where⁴

$$x_{it} = \delta_i + \xi_{it}$$

with

$$\begin{array}{ll} \mu_i & \sim & iid.N\left(0,\sigma_{\mu}^2\right), \, \delta_i \sim iid.U\left(-7.5,7.5\right) \\ \xi_{it} & \sim & iid.U\left(-5,5\right), \, \beta_0 = 5, \, \beta_1 = 0.5 \end{array}$$

⁴In the spirit of Nerlove (1971), we have tried another DGP for x_{it} . We obtain the same ranking as those which appear in the reported tables. The only difference is that the gap between the average heterogeneous estimators and the homogeneous estimators widens with a Nerlove (1971) type design. In other words, the forecast performance of the heterogeneous estimators becomes worse.

$$\phi_t = \begin{cases} \rho W_N \phi_t + v_t & \text{for SAR} \\ \lambda W_N v_t + v_t & \text{for SMA} \end{cases} \text{ with } \rho, \lambda = \begin{cases} 0.8 \\ 0.4 \end{cases}$$
(33)

and

$$v_{it} \sim iid.N\left(0,\sigma_v^2\right) \tag{34}$$

We consider the simple regressions (32) and (33) with N = (50, 100), T = (10, 20) and two cases for the residuals variances:

$$\begin{cases} \sigma_{\mu}^{2} = 4, & \sigma_{v}^{2} = 16\\ \sigma_{\mu}^{2} = 16, & \sigma_{v}^{2} = 4 \end{cases}$$
(35)

Following Kelejian and Prucha (1999), we use two weight matrices which essentially differ in their degree of sparseness. The weight matrices are labelled as "j ahead and j behind" with the non-zero elements being 1/2j, j = 1 and 5. Even with this modest design we have 64 experiments.

For each experiment, we obtain the following 16 estimators:

- 1. Pooled OLS which ignores the individual heterogeneity and the spatial autocorrelation.
- 2. The average heterogeneous OLS which estimates the cross-sectional equation using OLS for each time period and averages these heterogeneous estimates to obtain a pooled estimator, see Pesaran and Smith (1995).
- 3. The fixed-effects (FE) estimator which accounts for fixed individual effects but does not take into account the spatial autocorrelation.
- 4. The random effects (RE) estimator which assumes that the μ_i 's are $iid(0, \sigma_{\mu}^2)$, and independent of the remainder disturbances ϕ_{it} 's. This estimator accounts for random individual effects but does not take into account the spatial autocorrelation.
- 5. The RE-spatial MLE assuming a SAR specification (RE-SAR) on the remainder disturbances. In this case, the μ_i 's are $iid(0, \sigma_{\mu}^2)$ and are independent of the ϕ_{it} 's which follow a SAR process, see Anselin, et al. (2008).

- 6. The RE-spatial MLE assuming a SMA specification (RE-SMA) on the remainder disturbances. In this case, the μ_i 's are $iid(0, \sigma_{\mu}^2)$ and are independent of the ϕ_{it} 's which follow a SMA process, see Anselin, et al. (2008).
- 7. The pooled spatial MLE assuming a SAR specification (Pooled SAR) on the remainder disturbances. This estimator ignores the individual heterogeneity but takes into account the spatial autocorrelation of the SAR type.
- 8. The pooled spatial MLE assuming a SMA specification (Pooled SMA) on the remainder disturbances. This estimator ignores the individual heterogeneity but takes into account the spatial autocorrelation of the SMA type.
- 9. The average heterogeneous spatial MLE assuming a SAR specification on the remainder disturbances. This estimates cross-sectional MLE with SAR disturbances for each time period and averages the estimates over time.
- 10. The average heterogeneous spatial GM estimator assuming a SAR specification on the remainder disturbances proposed by Kelejian and Prucha (1999). This estimates cross-sectional GM estimator with SAR disturbances for each time period and averages the estimates over time.
- 11. The average heterogeneous spatial MLE assuming a SMA specification on the remainder disturbances. This estimates cross-sectional MLE with SMA disturbances for each time period and averages the estimates over time.
- 12. The average heterogeneous spatial GM estimator assuming a SMA specification on the remainder disturbances proposed by Fingleton (2007). This estimates cross-sectional GM estimator with SMA disturbances for each time period and averages the estimates over time.
- 13. The FE-spatial MLE assuming a SAR specification (FE-SAR) on the remainder disturbances.
- 14. The FE-spatial MLE assuming a SMA specification (FE-SMA) on the remainder disturbances.

- 15. The (SAR-RE) model following Kapoor, et al. (2007). This utilizes a panel data GM estimator where the disturbance term itself follows a SAR process and the remainder term follows an error component structure.
- 16. The (SMA-RE) model following Fingleton (2007). This utilizes a panel data GM estimator where the disturbance term itself follows a SMA process and the remainder term follows an error component structure.

Next, we compute the following predictors for the *i*th individual at a future period $T + \tau$ for $\tau = 1, 2, ..., 5$:

$\widehat{y}_{i,T+\tau} = X_{i,T+\tau} \widehat{\beta}_{OLS}$
$\widehat{y}_{i,T+\tau} = X_{i,T+\tau} \widehat{\beta}_{av.OLS}$
$\begin{cases} \widehat{y}_{i,T+\tau} = X_{i,T+\tau} \widehat{\beta}_{FE} + \widehat{\mu}_i \\ \text{with } \widehat{\mu}_i = \overline{y}_i - \overline{X}_i \widehat{\beta}_{FE}, \ \overline{y}_i = \sum_{t=1}^T y_{it}/T \end{cases}$
$\widehat{y}_{i,T+\tau} = X_{i,T+\tau} \widehat{\beta}_{RE} + \frac{\sigma_{\mu}^2}{\sigma_1^2} \left(\iota_T' \otimes l_i' \right) \widehat{\varepsilon}_{RE}$
$\begin{cases} \widehat{y}_{i,T+\tau} = X_{i,T+\tau} \widehat{\beta}_{MLE,RE-SAR} + \theta \left(\iota'_T \otimes l'_i C_1^{-1} \right) \widehat{\varepsilon}_{MLE,RE-SAR} \\ \text{with } C_1 = \left[T \theta I_N + \left(B'_N B_N \right)^{-1} \right] \text{ and } \theta = \sigma_{\mu}^2 / \sigma_v^2 \end{cases}$
$\begin{cases} \widehat{y}_{i,T+\tau} = X_{i,T+\tau} \widehat{\beta}_{MLE,RE-SMA} + \theta \left(\iota'_T \otimes l'_i C_2^{-1} \right) \widehat{\varepsilon}_{MLE,RE-SMA} \\ \text{with } C_2 = \left[T \theta I_N + (D_N D'_N) \right] \text{ and } \theta = \sigma_{\mu}^2 / \sigma_v^2 \end{cases}$
$\widehat{y}_{i,T+ au} = X_{i,T+ au} \widehat{eta}_{MLE,SAR}$
$\widehat{y}_{i,T+ au} = X_{i,T+ au} \widehat{eta}_{MLE,SMA}$
$\widehat{y}_{i,T+\tau} = \begin{cases} X_{i,T+\tau} \widehat{\beta}_{av.MLE,SAR} \\ X_{i,T+\tau} \widehat{\beta}_{av.GM,SAR} \end{cases}$
$\widehat{y}_{i,T+\tau} = \begin{cases} X_{i,T+\tau} \widehat{\beta}_{av.MLE,SMA} \\ X_{i,T+\tau} \widehat{\beta}_{av.GM,SMA} \end{cases}$
$\begin{cases} \widehat{y}_{i,T+\tau} = X_{i,T+\tau} \widehat{\beta}_{MLE,FE-SAR} + \widehat{\mu}_i \\ \text{with } \widehat{\mu}_i = \overline{y}_i - \overline{X}_i \widehat{\beta}_{MLE,FE-SAR} , \ \overline{y}_i = \sum_{t=1}^T y_{it}/T \end{cases}$
$\begin{cases} \widehat{y}_{i,T+\tau} = X_{i,T+\tau} \widehat{\beta}_{MLE,FE-SMA} + \widehat{\mu}_i \\ \text{with } \widehat{\mu}_i = \overline{y}_i - \overline{X}_i \widehat{\beta}_{MLE,FE-SMA}, \ \overline{y}_i = \sum_{t=1}^T y_{it}/T \end{cases}$
$\widehat{y}_{i,T+\tau} = X_{i,T+\tau} \overline{\widehat{\beta}_{MLE,SAR-RE} + \left(\frac{\sigma_{\mu}^2}{\sigma_1^2}\right) (\iota_T' \otimes l_i') \widehat{\varepsilon}_{MLE,SAR-RE}}$
$\widehat{y}_{i,T+\tau} = X_{i,T+\tau} \widehat{\beta}_{MLE,SMA-RE} + \left(\frac{\sigma_{\mu}^2}{\sigma_1^2}\right) \left(l_T' \otimes l_i'\right) \widehat{\varepsilon}_{MLE,SMA-RE}$

For all experiments, 1000 replications are performed and the RMSE for one step to five step ahead forecasts are reported.

5 Monte Carlo Results

5.1 The Spatial Dependence Specification Effect

Table 1 gives the RMSE for the one year, two year,..., and five year ahead forecasts along with the average RMSE for all 5 years. These are out of sample forecasts when the true DGP is a RE panel model with SAR remainder disturbances. The sample size is N = 50 and T = 10, the weight matrix is W(1,1), i.e., one neighbor behind and one neighbor ahead. In general, for $\rho = 0.4, 0.8$ and $\sigma_{\mu}^2 = 4, 16$, the lowest RMSE is that of RE-SAR. This is followed closely by SAR-RE and SMA-RE. It confirms the findings of Kapoor, et al. (2007) that, on average, RMSE of MLE and their GM estimators are quite similar. It also seems like misspecifying the SAR by an SMA in an error component model does not affect the forecast performance as long as it is taken into account. As the spatial autoregressive parameter ρ doubles from 0.4 to 0.8, the RMSE also doubles. The RMSE improves as σ_{μ}^2 gets large, i.e., 16 rather than 4, for estimators that take heterogeneity into account. Pooled OLS, average heterogeneous OLS, pooled SAR, pooled SMA, average heterogeneous SAR (MLE and GM) and average heterogeneous SMA (MLE and GM) perform worse in terms of RMSE than spatial/panel homogeneous estimators. This forecast comparison is robust whether we are predicting one period, two periods or 5 periods ahead and is also reflected in the average over the five years. The gain in forecast performance is substantial once we account for RE or FE and is only slightly improved by additionally accounting for spatial autocorrelation, i.e., FE-SAR or RE-SAR, FE-SMA, or RE-SMA.

Table 2 gives the RMSE results when the true DGP is a RE panel model with SMA remainder disturbances. The sample size is still N = 50, T = 10, and the weight matrix is W(1,1). In general, for $\rho = 0.4, 0.8$ and $\sigma_{\mu}^2 = 4, 16$, the lowest RMSE is that of RE-SMA. This is followed closely by RE-SAR.

⁵See Baillie and Baltagi (1998).

Misspecifying the SMA by an SAR in an error component model does not seem to affect the forecast performance as long as it is taken into account. However, the magnitudes of the RMSE in Table 2 (where the true DGP is a RE-SMA process) are much lower than those in Table 1 (where the true DGP is a RE-SAR process). Once again, the forecast RMSE of based on MLE and their GM counterparts are quite similar, compare SAR-RE and SMA-RE with RE-SAR and RE-SMA. The RMSE improves as σ_{μ}^2 gets large, i.e., 16 rather than 4, for estimators that take heterogeneity into account. As the spatial autoregressive parameter λ increases from 0.4 to 0.8, the RMSE also increases but not as much as it did for the SAR process in Table 1. Pooled OLS, average heterogeneous OLS, pooled SAR, pooled SMA, average heterogeneous SAR (MLE and GM) and average heterogeneous SMA (MLE and GM) perform worse in terms of RMSE than spatial/panel homogeneous estimators. This forecast performance is robust whether we are predicting one period, two periods or 5 periods ahead and is also reflected in the average over the five years. Once again, the gain in forecast performance is substantial once we account for RE or FE and is only slightly improved by additionally accounting for spatial autocorrelation, i.e., FE-SMA, or RE-SMA, FE-SAR or RE-SAR.

5.2 Sensitivity Analysis

5.2.1 The Spatial Weight Matrix effect

Tables 3 and 4 report the RMSE results as Tables 1 and 2 except that the weight matrix is changed from a W(1,1) to W(5,5), i.e., five neighbors behind and five neighbors ahead. Except for the magnitudes of the RMSE, the same rankings in terms of RMSE performance are exhibited as before.

Tables 5 and 6 report the RMSE results as Tables 1 and 2 except that T is now doubled from 10 to 20 holding N fixed at 50. Except for the magnitudes of the RMSE, the same rankings in terms of RMSE performance are exhibited as before.

Table 7 reports the RMSE results when $\rho = \lambda = 0.8$, the weight matrix is W(1, 1), and N is doubled from 50 to 100 holding T fixed at 10. While Table 8 reports the RMSE results as Table 7 except that the weight matrix is W(5, 5). Except for the magnitudes of the RMSE, the same rankings in terms of RMSE performance are exhibited as before.⁶

⁶Other Tables for W(5,5) and (N,T) = (100,20) show the same rankings in terms of

5.2.2 Sensitivity to Irregular Lattice Structures

The spatial weights matrices considered in the paper are regular lattice structures. Using real irregular lattices structures, as in Anselin and Moreno (2003) and in Kelejian and Prucha (1999), does not change the conclusions of the Monte Carlo study. We used real-world matrices by taking spatial groupings of French administrative communes for dimension $N = 50.^7$ Those spatial matrices have been used by Baltagi, Bresson and Pirotte (2007). Spatial weight matrices may represent high-order contiguity relationships. We use a k-order contiguity matrix containing N-1 potential neighborhoods in French municipalities. We have patterns of 0 and 1 values in an (N-1)by (N-1) grid for the k-nearest neighborhoods and we use the 1-nearest neighborhood (k = 1) and the 5-nearest neighborhoods $(k = 1)^8$. Results of Tables 9 to 12 are very similar to those of Tables 1 to 4. Using irregular lattice structures do not change the main conclusions in terms of the RMSE forecast performance of the various estimators considered. These are similar to the rankings obtained when regular lattice structures are used, only the magnitudes of the RMSE differ.

5.2.3 Robustness to Non-Normality

So far, we have been assuming that the error components have been generated by the normal distribution. In this section, we check the sensitivity of our results to non-normal disturbances. In particular, we generate the μ_i 's from a χ^2 distribution and we let the remainder disturbances follow the normal distribution. Tables 13 and 14 give similar results as those of Tables 1 and 2 (when the individual effects follow a normal distribution). So, the results seem to be robust to non-normality of the disturbances of the χ^2 type.

RMSE forecast performance and are not shown here to save space. These are available upon request from the authors.

⁷Other Tables for N = 100 are available upon request from the authors.

⁸Note that a non-zero entry in row *i*, column *j* denotes that neighborhoods *i* and *j* have borders that touch and are therefore considered "neighbors". For N = 50 and for k = 5, and for the 2401 possible elements in the 49 by 49 matrix, there are only 250 non-zero elements. So, the sparseness value is 10% (= 250/2500). These non-zero entries reflect the contiguity relations between the 5-nearest neighborhoods.

6 Summary and Conclusion

Our Monte Carlo study finds that when the true DGP is RE with a SAR or SMA remainder disturbances, estimators that ignore heterogeneity/spatial correlation perform badly in RMSE forecasts. For our experiments, accounting for heterogeneity improves the forecast performance by a big margin and accounting for spatial correlation improves the forecast but by a smaller margin. Ignoring both leads to the worst forecasting performance. Heterogeneous estimators based on averaging perform worse than homogeneous estimators in forecasting performance. This performance improves with a larger sample size and seems robust to the type of spatial error structure imposed on the remainder disturbances. These Monte Carlo experiments confirm earlier empirical studies that report similar findings.

7 Appendix

This appendix first derives the BLUP for the KKP model which we are calling the (SAR-RE) model described in (13) and (14). The variance-covariance matrix Ω is given in (17). The inverse of Ω is given by:

$$\Omega^{-1} = \frac{1}{\sigma_v^2} \left[\left(I_T - \frac{T \sigma_\mu^2}{\sigma_1^2} \overline{J}_T \right) \otimes \left(B'_N B_N \right) \right]$$

where $\overline{J}_T = J_T/T$ and $\sigma_1^2 = T\sigma_\mu^2 + \sigma_v^2$ and $B_N = (I_N - \rho W_N)$. From (13) and (14), we have :

$$\varepsilon_{T+\tau} = B_N^{-1} u_{T+\tau} = B_N^{-1} (\mu + v_{T+\tau})$$

so that,

$$E\left[\varepsilon_{T+\tau}\varepsilon'\right] = E\left[B_N^{-1}\left(\mu + v_{T+\tau}\right)\left(\left(\iota_T \otimes B_N^{-1}\right)\mu + \left(I_T \otimes B_N^{-1}\right)v\right)'\right]$$
$$= \sigma_{\mu}^2 B_N^{-1}\left(\iota'_T \otimes B_N^{-1'}\right)$$
$$\omega' = E\left[\varepsilon_{i,T+\tau}\varepsilon'\right] = \sigma_{\mu}^2 b_i\left(\iota'_T \otimes B_N^{-1'}\right)$$

where b_i is the *i*th row of the matrix B_N^{-1} . In this case,

$$\omega'\Omega^{-1} = \frac{\sigma_{\mu}^2}{\sigma_v^2} b_i \left(\iota'_T \otimes B_N^{-1'}\right) \left[\left(I_T - \frac{T\sigma_{\mu}^2}{\sigma_1^2} \overline{J}_T \right) \otimes \left(B'_N B_N \right) \right]$$

$$= \frac{\sigma_{\mu}^{2}}{\sigma_{v}^{2}} b_{i} \left[\left(\iota_{T}^{'} \otimes B_{N} \right) - \frac{T \sigma_{\mu}^{2}}{\sigma_{1}^{2}} \left(\iota_{T}^{'} \otimes B_{N} \right) \right]$$
$$= \frac{\sigma_{\mu}^{2}}{\sigma_{1}^{2}} b_{i} \left(\iota_{T}^{'} \otimes B_{N} \right)$$

But $b_i (\iota'_T \otimes B_N) = (1 \otimes b_i) (\iota'_T \otimes B_N) = (\iota'_T \otimes l'_i)$, where l'_i is the *i*th row of I_N . This holds because $B_N^{-1}B_N = I_N$ and therefore $b_iB_N = l'_i$. This means that the predictor of the KKP model from (28) is given by:

$$\widehat{y}_{i,T+\tau} = X_{i,T+\tau} \widehat{\beta}_{GLS} + \frac{\sigma_{\mu}^2}{\sigma_1^2} \left(\iota_T' \otimes l_i' \right) \widehat{\varepsilon}_{GLS}$$
(36)

which is the same as that of the RE model with no spatial correlation. While the predictor formula is the same, the MLEs for these specifications yield different estimates which in turn yield different residuals and hence different forecasts.

The proof is the similar for the Fingleton (2007) specification which we are calling the (SMA-RE) model described in (25) and (14). The variance-covariance matrix Ω is given in (27). The inverse of Ω is given by:

$$\Omega^{-1} = \frac{1}{\sigma_v^2} \left[\left(I_T - \frac{T \sigma_\mu^2}{\sigma_1^2} \overline{J}_T \right) \otimes \left(D_N D_N' \right)^{-1} \right]$$

where $D_N = (I_N + \lambda W_N)$. From (25) and (14), we have :

$$\varepsilon_{T+\tau} = D_N u_{T+\tau} = D_N \left(\mu + v_{T+\tau} \right)$$

so that,

$$E \left[\varepsilon_{T+\tau} \varepsilon' \right] = E \left[D_N \left(\mu + v_{T+\tau} \right) \left(\left(\iota_T \otimes D_N \right) \mu + \left(I_T \otimes D_N \right) v \right)' \right] \\ = \sigma_{\mu}^2 D_N \left(\iota'_T \otimes D'_N \right) \\ \omega' = E \left[\varepsilon_{i,T+\tau} \varepsilon' \right] = \sigma_{\mu}^2 d_i \left(\iota'_T \otimes D'_N \right)$$

where d_i is the *i*th row of the matrix D_N . In this case,

$$\begin{split} \omega' \Omega^{-1} &= \frac{\sigma_{\mu}^2}{\sigma_v^2} d_i \left(\iota'_T \otimes D'_N \right) \left[\left(I_T - \frac{T \sigma_{\mu}^2}{\sigma_1^2} \overline{J}_T \right) \otimes (D_N D'_N)^{-1} \right] \\ &= \frac{\sigma_{\mu}^2}{\sigma_v^2} d_i \left[\left(\iota'_T \otimes D_N^{-1} \right) - \frac{T \sigma_{\mu}^2}{\sigma_1^2} \left(\iota'_T \otimes D_N^{-1} \right) \right] \\ &= \frac{\sigma_{\mu}^2}{\sigma_1^2} d_i \left(\iota'_T \otimes D_N^{-1} \right) \end{split}$$

But $d_i (\iota'_T \otimes D_N^{-1}) = (1 \otimes d_i) (\iota'_T \otimes D_N^{-1}) = (\iota'_T \otimes l'_i)$, where l'_i is the *i*th row of I_N . This holds because $D_N D_N^{-1} = I_N$ and therefore $d_i D_N^{-1} = l'_i$. This means that the predictor of the Fingleton (2007) model is again the same as that of the RE model with no spatial correlation. While the predictor formula is the same, the MLEs for these specifications yield different estimates which in turn yield different residuals and hence different forecasts.

References

- Anselin, L., 1988, Spatial Econometrics: Methods and Models, Kluwer Academic Publishers, Dordrecht.
- Anselin, L. and A.K. Bera, 1998, Spatial dependence in linear regression models with an introduction to spatial econometrics. In A. Ullah and D.E.A. Giles, eds., Handbook of Applied Economic Statistics, Marcel Dekker, New York.
- Anselin, L. and R. Moreno, 2003, Properties of tests for spatial error components, Regional Science and Urban Economics 33, 595-618.
- Anselin, L., J. Le Gallo and H. Jayet, 2008, Spatial panel econometrics. Ch. 19 in L. Mátyás and P. Sevestre, eds., The Econometrics of Panel Data: Fundamentals and Recent Developments in Theory and Practice, Springer-Verlag, Berlin, 625-660.
- Baillie, R.T. and B.H. Baltagi, 1999, Prediction from the regression model with one-way error components, Chapter 10 in C. Hsiao, K. Lahiri, L.F. Lee and H. Pesaran, eds., Analysis of Panels and Limited Dependent Variable Models, Cambridge University Press, Cambridge, 255–267.
- Baltagi, B.H., 2008, Forecasting with panel data, Journal of Forecasting 27, 153-173...
- Baltagi, B.H. and J.M. Griffin, 1997, Pooled estimators vs. their heterogeneous counterparts in the context of dynamic demand for gasoline, Journal of Econometrics 77, 303–327.
- Baltagi, B.H. and D. Li, 2004, Prediction in the panel data model with spatial correlation, Chapter 13 in L. Anselin, R.J.G.M. Florax and S.J. Rey, eds., Advances in Spatial Econometrics: Methodology, Tools and Applications, Springer, Berlin, 283–295.
- Baltagi, B.H. and D. Li, 2006, Prediction in the panel data model with spatial correlation: The case of liquor, Spatial Economic Analysis 1, 175-185.
- Baltagi, B.H. and Q. Li, 1992, Prediction in the one-way error component model with serial correlation, Journal of Forecasting 11, 561–567.
- Baltagi, B.H., G. Bresson and A. Pirotte, 2002, Comparison of forecast performance for homogeneous, heterogeneous and shrinkage estimators: Some empirical evidence from US electricity and natural-gas consumption, Economics Letters 76, 375-382.

- Baltagi, B.H., G. Bresson and A. Pirotte, 2004, Tobin q: forecast performance for hierarchical Bayes, shrinkage, heterogeneous and homogeneous panel data estimators, Empirical Economics 29, 107-113.
- Baltagi, B.H., G. Bresson and A. Pirotte, 2007, Panel unit root tests and spatial dependence, Journal of Applied Econometrics 22, 339-360.
- Baltagi, B.H., J.M. Griffin and W. Xiong, 2000, To pool or not to pool: Homogeneous versus heterogeneous estimators applied to cigarette demand, Review of Economics and Statistics 82, 117–126.
- Brucker, H. and B. Siliverstovs, 2006, On the estimation and forecasting of international migration: how relevant is heterogeneity across countries, Empirical Economics 31, 735-754.
- Fingleton, B., 2007a, A generalized method of moments estimator for a spatial model with endogenous spatial lag and spatial moving average errors, paper presented at the 13th international conference on panel data, University of Cambridge, forthcoming Spatial Economic Analysis.
- Fingleton, B., 2007b, A generalized method of moments estimator for a spatial model with moving average errors with application to real estate prices, forthcoming in Empirical Economics.
- Frees, E.W. and T.W. Miller, 2004, Sales forecasting using longitudinal data models. International Journal of Forecasting 20, 99–114.
- Goldberger, A.S., 1962, Best linear unbiased prediction in the generalized linear regression model, Journal of the American Statistical Association 57, 369–375.
- Kapoor, M., H.H. Kelejian and I.R. Prucha, 2007, Panel data models with spatially correlated error components, Journal of Econometrics 140, 97-130.
- Kelejian, H.H. and I.R. Prucha, 1999, A generalized moments estimator for the autoregressive parameter in a spatial model, International Economic Review 40, 509-533.
- Lee, L.F. and W.E. Griffiths, 1979, The prior likelihood and best linear unbiased prediction in stochastic coefficient linear models, working paper, Department of Economics, University of Minnesota.

- Hoogstrate, A.J., F.C. Palm and G.A. Pfann, 2000, Pooling in dynamic panel-data models: An application to forecasting GDP growth rates, Journal of Business and Economic Statistics 18, 274-283.
- Hsiao, C. and A.K. Tahmiscioglu, 1997, A panel analysis of liquidity constraints and firm investment, Journal of the American Statistical Association 92, 455–465.
- Nerlove, M., 1971, Futher evidence on the estimation of dynamic economic relations from a time-series of cross-sections, Econometrica 39, 359-382.
- Pesaran, M.H. and R. Smith, 1995, Estimating long-run relationships from dynamic heterogenous panels, Journal of Econometrics 68, 79–113.
- Rapach, D.E. and M.E. Wohar, 2004, Testing the monetary model of exchange rate determination: a closer look at panels, Journal of International Money and Finance 23, 867–895.
- Schmalensee, R., T.M. Stoker and R.A. Judson, 1998, World carbon dioxide emissions: 1950-2050, Review of Economics and Statistics 80, 15–27.
- Spanos, A., 2002, The ET interview: Professor Phoebus J. Dhrymes, Econometric Theory 18, 1221-1272.
- Taub, A.J., 1979, Prediction in the context of the variance-components model, Journal of Econometrics 10, 103–108.
- Theil, H., 1961, Economic Forecasts and Policy, North-Holland, Amsterdam.
- Wansbeek, T.J. and A. Kapteyn, 1978, The separation of individual variation and systematic change in the analysis of panel data, Annales de l'INSEE 30-31, 659-680.

										Esti	mators							
			Pooled	Av. hetero.			Pooled SAR	Av. hete	ero. SAR	FE-SAR	RE-SAR	Pooled SMA	Av. hete	ro. SMA	FE-SMA	RE-SMA	SAR-RE	SMA-RE
	ρ	σ_{μ}^{2}	OLS	OLS	FE	RE	MLE	MLE	GM	MLE	MLE	MLE	MLE	GM	MLE	MLE	GM	GM
	0.4	4	3.9781	3.9782	3.8102	3.7645	3.9781	3.9781	3.9606	3.8093	3.7558	3.9782	3.9782	3.9464	3.8093	3.7558	3.7610	3.7765
1st vear		16	3.6289	3.6290	1.9019	1.8989	3.6300	3.6299	3.6522	1.9007	1.8971	3.6301	3.6300	3.6569	1.9007	1.8973	1.8978	1.9134
Tot your	0.8	4	7.0556	7.0552	7.1957	7.0218	7.0529	7.0389	7.0558	7.1917	6.9564	7.0541	7.0403	7.0796	7.1917	6.9668	7.0187	7.0382
		16	4.6529	4.6533	3.5908	3.5764	4.6584	4.6569	4.6697	3.5863	3.5518	4.6589	4.6576	4.6644	3.5867	3.5603	3.6047	3.5908
	0.4	4	4.4164	4.4165	4.2360	4.1840	4.4162	4.4162	4.3423	4.2354	4.1763	4.4164	4.4164	4.3721	4.2353	4.1755	4.1808	4.1739
2nd year		16	3.8731	3.8733	2.1207	2.1175	3.8742	3.8743	3.8849	2.1194	2.1155	3.8742	3.8744	3.8918	2.1195	2.1156	2.1164	2.1216
2nd year	0.8	4	7.8106	7.8106	7.9469	7.7633	7.8066	7.7911	7.8100	7.9408	7.6956	7.8073	7.7920	7.8306	7.9407	7.7034	7.7832	7.8190
		16	5.1174	5.1177	4.0090	3.9942	5.1221	5.1209	5.1206	4.0039	3.9661	5.1225	5.1213	5.1084	4.0042	3.9754	3.9923	3.9833
	0.4	4	4.5807	4.5808	4.3992	4.3445	4.5805	4.5805	4.5627	4.3986	4.3364	4.5806	4.5807	4.5560	4.3985	4.3357	4.3414	4.3475
3rd year		16	3.9582	3.9585	2.2004	2.1972	3.9591	3.9592	3.9660	2.1992	2.1954	3.9591	3.9594	3.9682	2.1993	2.1956	2.1963	2.2023
Siù yeai	0.8	4	8.1467	8.1467	8.2921	8.1023	8.1424	8.1273	8.1458	8.2853	8.0289	8.1430	8.1279	8.1618	8.2854	8.0382	8.1016	8.1402
		16	5.2892	5.2894	4.1685	4.1529	5.2936	5.2923	5.2763	4.1636	4.1234	5.2940	5.2928	5.2674	4.1640	4.1337	4.1387	4.1450
	0.4	4	4.6719	4.6720	4.4891	4.4335	4.6718	4.6717	4.6676	4.4882	4.4250	4.6719	4.6720	4.6583	4.4881	4.4245	4.4301	4.4332
Ath year		16	4.0024	4.0026	2.2471	2.2440	4.0031	4.0033	4.0117	2.2460	2.2423	4.0032	4.0034	4.0125	2.2461	2.2424	2.2432	2.2451
401 year	0.8	4	8.3035	8.3035	8.4435	8.2560	8.2997	8.2836	8.3011	8.4367	8.1826	8.3005	8.2843	8.3225	8.4370	8.1922	8.3142	8.3214
		16	5.3799	5.3802	4.2531	4.2377	5.3838	5.3826	5.3662	4.2481	4.2085	5.3841	5.3829	5.3626	4.2485	4.2183	4.2296	4.2143
	0.4	4	4.7238	4.7239	4.5443	4.4870	4.7238	4.7238	4.7274	4.5432	4.4778	4.7239	4.7240	4.7199	4.5432	4.4775	4.4836	4.4906
5th year		16	4.0283	4.0285	2.2727	2.2698	4.0288	4.0290	4.0362	2.2716	2.2681	4.0289	4.0291	4.0374	2.2716	2.2682	2.2689	2.2718
Juryea	0.8	4	8.4195	8.4197	8.5680	8.3756	8.4158	8.3995	8.4173	8.5606	8.2997	8.4164	8.4001	8.4331	8.5608	8.3100	8.4299	8.4256
		16	5.4280	5.4282	4.2962	4.2812	5.4313	5.4301	5.4156	4.2911	4.2526	5.4317	5.4305	5.4125	4.2914	4.2618	4.2808	4.2710
	0.4	4	4.4742	4.4743	4.2957	4.2427	4.4741	4.4740	4.4601	4.2949	4.2343	4.4742	4.4743	4.4505	4.2949	4.2338	4.2394	4.2444
Average		16	3.8982	3.8984	2.1486	2.1455	3.8990	3.8992	3.9102	2.1474	2.1437	3.8991	3.8993	3.9133	2.1474	2.1438	2.1445	2.1509
Average	0.8	4	7.9472	7.9471	8.0892	7.9038	7.9435	7.9281	7.9460	8.0830	7.8326	7.9443	7.9289	7.9655	8.0831	7.8421	7.9295	7.9489
		16	5.1735	5.1738	4.0635	4.0485	5.1778	5.1766	5.1697	4.0586	4.0205	5.1782	5.1770	5.1631	4.0589	4.0299	4.0492	4.0409

Table 1 - Forecasts RMSE - (N,T)=(50,10), SAR data generating process for φ, W(1,1), 1000 replications

										Estir	mators							
			Pooled	Av. hetero.			Pooled SAR	Av. hete	ero. SAR	FE-SAR	RE-SAR	Pooled SMA	Av. hete	ro. SMA	FE-SMA	RE-SMA	SAR-RE	SMA-RE
	λ	$\sigma_{_{\!$	OLS	OLS	FE	RE	MLE	MLE	GM	MLE	MLE	MLE	MLE	GM	MLE	MLE	GM	GM
	0.4	4	3.6702	3.6703	3.4717	3.4261	3.6704	3.6701	3.6706	3.4707	3.4193	3.6705	3.6702	3.6669	3.4707	3.4187	3.4330	3.4375
1st vear		16	3.5582	3.5582	1.7481	1.7455	3.5583	3.5584	3.5606	1.7478	1.7450	3.5584	3.5585	3.5569	1.7479	1.7449	1.7444	1.7323
ist year	0.8	4	3.9870	3.9873	3.8507	3.7906	3.9857	3.9856	3.9878	3.8481	3.7653	3.9858	3.9859	3.9770	3.8480	3.7635	3.7915	3.8213
		16	3.6364	3.6364	1.9117	1.9095	3.6381	3.6377	3.6235	1.9098	1.9068	3.6380	3.6376	3.6316	1.9097	1.9062	1.9270	1.9078
	0.4	4	4.0793	4.0793	3.8608	3.8133	4.0794	4.0792	4.0796	3.8600	3.8060	4.0795	4.0793	4.0816	3.8600	3.8056	3.8269	3.8097
2nd year		16	3.7747	3.7748	1.9380	1.9354	3.7751	3.7754	3.7756	1.9375	1.9346	3.7752	3.7755	3.7759	1.9376	1.9345	1.9282	1.9255
zna year	0.8	4	4.4390	4.4391	4.2819	4.2224	4.4388	4.4386	4.4209	4.2783	4.1971	4.4396	4.4396	4.4105	4.2777	4.1952	4.2168	4.2313
		16	3.8696	3.8697	2.1223	2.1191	3.8716	3.8714	3.8791	2.1201	2.1157	3.8718	3.8717	3.8821	2.1199	2.1149	2.1374	2.1270
	0.4	4	4.2357	4.2358	4.0121	3.9644	4.2358	4.2358	4.2367	4.0111	3.9563	4.2360	4.2359	4.2393	4.0111	3.9560	3.9785	3.9661
3rd voar		16	3.8526	3.8527	2.0109	2.0084	3.8531	3.8534	3.8521	2.0104	2.0074	3.8531	3.8535	3.8537	2.0104	2.0074	2.0076	2.0047
Siù yeai	0.8	4	4.6176	4.6177	4.4527	4.3926	4.6176	4.6175	4.5966	4.4490	4.3684	4.6184	4.6184	4.5887	4.4483	4.3652	4.3835	4.3981
		16	3.9613	3.9614	2.2116	2.2084	3.9636	3.9634	3.9673	2.2095	2.2054	3.9638	3.9636	3.9692	2.2094	2.2045	2.2194	2.2168
	0.4	4	4.3113	4.3114	4.0834	4.0354	4.3110	4.3111	4.3131	4.0823	4.0263	4.3112	4.3112	4.3161	4.0823	4.0263	4.0608	4.0475
Ath year		16	3.8901	3.8902	2.0500	2.0474	3.8905	3.8908	3.8901	2.0494	2.0464	3.8906	3.8909	3.8917	2.0494	2.0463	2.0465	2.0482
4iii yeai	0.8	4	4.7133	4.7133	4.5470	4.4856	4.7135	4.7134	4.6870	4.5433	4.4613	4.7146	4.7144	4.6837	4.5428	4.4581	4.4617	4.4901
		16	4.0100	4.0101	2.2659	2.2627	4.0122	4.0121	4.0134	2.2642	2.2598	4.0123	4.0122	4.0152	2.2639	2.2590	2.2619	2.2631
	0.4	4	4.3637	4.3638	4.1367	4.0876	4.3635	4.3635	4.3653	4.1357	4.0786	4.3637	4.3637	4.3695	4.1357	4.0786	4.1094	4.0975
5th year		16	3.9122	3.9124	2.0748	2.0722	3.9126	3.9129	3.9147	2.0743	2.0712	3.9127	3.9131	3.9147	2.0742	2.0711	2.0729	2.0752
Juryear	0.8	4	4.7697	4.7698	4.6004	4.5394	4.7702	4.7700	4.7457	4.5967	4.5160	4.7713	4.7712	4.7428	4.5963	4.5125	4.5223	4.5371
		16	4.0405	4.0405	2.2956	2.2926	4.0426	4.0425	4.0396	2.2937	2.2898	4.0428	4.0427	4.0411	2.2934	2.2889	2.2890	2.2901
	0.4	4	4.1321	4.1321	3.9129	3.8653	4.1320	4.1319	4.1331	3.9120	3.8573	4.1322	4.1320	4.1347	3.9120	3.8570	3.8817	3.8717
Average		16	3.7976	3.7977	1.9644	1.9618	3.7979	3.7982	3.7986	1.9639	1.9609	3.7980	3.7983	3.7986	1.9639	1.9608	1.9599	1.9572
Average	0.8	4	4.5053	4.5054	4.3466	4.2861	4.5051	4.5050	4.4876	4.3431	4.2616	4.5059	4.5059	4.4805	4.3426	4.2589	4.2752	4.2956
		16	3.9036	3.9036	2.1614	2.1584	3.9056	3.9054	3.9046	2.1594	2.1555	3.9058	3.9055	3.9079	2.1592	2.1547	2.1669	2.1610

Table 2 - Forecasts RMSE - (N,T)=(50,10), SMA data generating process for φ, W(1,1), 1000 replications

										Estim	ators							
			Pooled	Av. hetero.		DE	Pooled SAR	Av. hete	ro. SAR	FE-SAR	RE-SAR	Pooled SMA	Av. hete	ero. SMA	FE-SMA	RE-SMA	SAR-RE	SMA-RE
	ρ	σ_{μ}^{2}	OLS	OLS	FE	ĸe	MLE	MLE	GM	MLE	MLE	MLE	MLE	GM	MLE	MLE	GM	GM
	0.4	4	3.6604	3.6604	3.4537	3.4118	3.6600	3.6601	3.6448	3.4535	3.4095	3.6602	3.6603	3.6338	3.4536	3.4095	3.4415	3.4417
1st vear		16	3.5736	3.5735	1.7414	1.7405	3.5737	3.5739	3.5697	1.7410	1.7402	3.5736	3.5739	3.5671	1.7411	1.7403	1.7351	1.7342
i ot your	0.8	4	4.9150	4.9149	4.8055	4.7733	4.9133	4.8926	4.8984	4.8040	4.7355	4.9135	4.8926	4.9076	4.8040	4.7365	4.7109	4.7822
		16	3.9279	3.9279	2.4155	2.4138	3.9283	3.9272	3.8975	2.4131	2.4063	3.9284	3.9270	3.9042	2.4133	2.4080	2.3947	2.3975
	0.4	4	4.0515	4.0516	3.8391	3.7896	4.0515	4.0515	4.0529	3.8386	3.7868	4.0515	4.0516	4.0498	3.8386	3.7868	3.8221	3.8177
2nd vear		16	3.7783	3.7784	1.9310	1.9294	3.7785	3.7786	3.7817	1.9307	1.9290	3.7786	3.7790	3.7805	1.9307	1.9291	1.9233	1.9236
Lina your	0.8	4	5.4516	5.4517	5.3368	5.2966	5.4509	5.4281	5.4571	5.3352	5.2602	5.4510	5.4283	5.4597	5.3351	5.2595	5.2384	5.3148
		16	4.2189	4.2188	2.6745	2.6715	4.2195	4.2180	4.2021	2.6725	2.6625	4.2198	4.2181	4.1977	2.6725	2.6652	2.6551	2.6764
	0.4	4	4.2132	4.2133	3.9946	3.9444	4.2133	4.2134	4.2197	3.9941	3.9419	4.2133	4.2135	4.2203	3.9942	3.9419	3.9698	3.9781
3rd vear		16	3.8499	3.8500	2.0047	2.0029	3.8500	3.8503	3.8582	2.0045	2.0025	3.8501	3.8506	3.8556	2.0045	2.0025	2.0018	1.9992
ora your	0.8	4	5.6484	5.6484	5.5355	5.4903	5.6473	5.6280	5.6924	5.5331	5.4516	5.6475	5.6282	5.6855	5.5331	5.4521	5.4781	5.5263
		16	4.3224	4.3224	2.7734	2.7701	4.3232	4.3219	4.3141	2.7716	2.7613	4.3235	4.3221	4.3130	2.7717	2.7640	2.7746	2.7792
	0.4	4	4.3083	4.3083	4.0871	4.0372	4.3086	4.3085	4.3133	4.0867	4.0341	4.3085	4.3087	4.3118	4.0867	4.0342	4.0471	4.0522
4th vear		16	3.8902	3.8902	2.0461	2.0442	3.8903	3.8904	3.8991	2.0458	2.0437	3.8904	3.8908	3.8944	2.0458	2.0437	2.0420	2.0440
nin your	0.8	4	5.7632	5.7632	5.6516	5.6042	5.7617	5.7412	5.7872	5.6492	5.5624	5.7619	5.7413	5.7905	5.6492	5.5644	5.5835	5.6212
		16	4.3837	4.3836	2.8346	2.8315	4.3844	4.3831	4.3727	2.8325	2.8224	4.3847	4.3833	4.3731	2.8326	2.8250	2.8343	2.8335
	0.4	4	4.3621	4.3621	4.1403	4.0901	4.3623	4.3622	4.3606	4.1399	4.0869	4.3622	4.3624	4.3587	4.1399	4.0870	4.0913	4.1018
5th vear		16	3.9133	3.9134	2.0714	2.0695	3.9135	3.9137	3.9247	2.0712	2.0691	3.9136	3.9141	3.9197	2.0712	2.0691	2.0665	2.0674
our you	0.8	4	5.8382	5.8382	5.7313	5.6808	5.8371	5.8162	5.8620	5.7293	5.6378	5.8372	5.8164	5.8640	5.7292	5.6407	5.6510	5.7074
		16	4.4187	4.4186	2.8668	2.8640	4.4195	4.4182	4.4019	2.8648	2.8551	4.4198	4.4184	4.3993	2.8649	2.8576	2.8702	2.8755
	0.4	4	4.1191	4.1191	3.9030	3.8546	4.1191	4.1191	4.1182	3.9025	3.8518	4.1191	4.1193	4.1149	3.9026	3.8519	3.8744	3.8783
Average		16	3.8010	3.8011	1.9589	1.9573	3.8012	3.8014	3.8067	1.9586	1.9569	3.8012	3.8017	3.8035	1.9587	1.9569	1.9537	1.9537
, we age	0.8	4	5.5233	5.5233	5.4121	5.3690	5.5221	5.5012	5.5394	5.4102	5.3295	5.5222	5.5014	5.5414	5.4101	5.3306	5.3324	5.3904
		16	4.2543	4.2542	2.7129	2.7102	4.2550	4.2537	4.2376	2.7109	2.7015	4.2552	4.2538	4.2375	2.7110	2.7040	2.7058	2.7124

Table 3 - Forecasts RMSE - (N,T)=(50,10), SAR data generating process for φ, W(5,5), 1000 replications

										Estima	ators							
			Pooled	Av. hetero.		DE	Pooled SAR	Av. hete	ero. SAR	FE-SAR	RE-SAR	Pooled SMA	Av. hete	ero. SMA	FE-SMA	RE-SMA	SAR-RE	SMA-RE
	λ	$\sigma_{_{\mu}}^{_{2}}$	OLS	OLS	FE	KE	MLE	MLE	GM	MLE	MLE	MLE	MLE	GM	MLE	MLE	GM	GM
	0.4	4	3.5909	3.5909	3.3764	3.3363	3.5911	3.5912	3.5823	3.3759	3.3356	3.5911	3.5911	3.5859	3.3759	3.3354	3.3417	3.3338
1st vear		16	3.5114	3.5116	1.6868	1.6846	3.5115	3.5116	3.5249	1.6867	1.6843	3.5115	3.5121	3.5175	1.6867	1.6843	1.6906	1.6851
ist year	0.8	4	3.6763	3.6763	3.4731	3.4312	3.6759	3.6761	3.6593	3.4724	3.4280	3.6760	3.6763	3.6546	3.4722	3.4275	3.4035	3.3943
		16	3.5299	3.5300	1.7293	1.7273	3.5300	3.5304	3.5553	1.7291	1.7269	3.5301	3.5306	3.5662	1.7290	1.7267	1.7231	1.7248
	0.4	4	3.9719	3.9720	3.7426	3.6993	3.9721	3.9724	3.9858	3.7424	3.6988	3.9721	3.9726	3.9829	3.7424	3.6985	3.7173	3.7026
2nd year		16	3.7285	3.7287	1.8800	1.8778	3.7289	3.7291	3.7246	1.8798	1.8774	3.7289	3.7295	3.7237	1.8798	1.8774	1.8757	1.8718
Zhu year	0.8	4	4.0662	4.0662	3.8506	3.8034	4.0658	4.0659	4.0556	3.8497	3.7987	4.0659	4.0660	4.0490	3.8496	3.7984	3.7954	3.7802
		16	3.7328	3.7328	1.9132	1.9106	3.7330	3.7333	3.7692	1.9130	1.9100	3.7331	3.7335	3.7823	1.9129	1.9101	1.9116	1.9099
	0.4	4	4.1258	4.1258	3.8950	3.8495	4.1259	4.1262	4.1356	3.8946	3.8482	4.1259	4.1264	4.1315	3.8946	3.8480	3.8544	3.8491
Ordvoor		16	3.8057	3.8059	1.9534	1.9513	3.8059	3.8061	3.8035	1.9533	1.9510	3.8060	3.8065	3.8018	1.9533	1.9510	1.9528	1.9464
Sid year	0.8	4	4.2227	4.2227	3.9958	3.9493	4.2222	4.2224	4.2096	3.9949	3.9452	4.2223	4.2226	4.2038	3.9947	3.9447	3.9422	3.9333
		16	3.8127	3.8128	1.9925	1.9899	3.8129	3.8131	3.8410	1.9923	1.9896	3.8130	3.8134	3.8527	1.9923	1.9896	1.9869	1.9913
	0.4	4	4.2050	4.2051	3.9729	3.9270	4.2050	4.2053	4.2140	3.9726	3.9258	4.2050	4.2055	4.2135	3.9726	3.9255	3.9288	3.9275
1th year		16	3.8465	3.8466	1.9921	1.9902	3.8467	3.8469	3.8420	1.9919	1.9898	3.8467	3.8473	3.8395	1.9919	1.9899	1.9900	1.9894
4th year	0.8	4	4.3004	4.3004	4.0741	4.0261	4.2999	4.3001	4.2914	4.0734	4.0214	4.3000	4.3002	4.2846	4.0732	4.0209	4.0191	4.0143
		16	3.8530	3.8531	2.0306	2.0282	3.8531	3.8533	3.8774	2.0304	2.0280	3.8532	3.8536	3.8810	2.0303	2.0279	2.0313	2.0293
	0.4	4	4.2560	4.2561	4.0203	3.9746	4.2560	4.2562	4.2663	4.0200	3.9735	4.2560	4.2564	4.2663	4.0200	3.9733	3.9762	3.9787
Ethylogr		16	3.8694	3.8696	2.0158	2.0139	3.8697	3.8700	3.8634	2.0157	2.0136	3.8697	3.8704	3.8610	2.0157	2.0136	2.0141	2.0161
our year	0.8	4	4.3474	4.3474	4.1229	4.0736	4.3469	4.3472	4.3447	4.1222	4.0683	4.3470	4.3473	4.3401	4.1221	4.0678	4.0724	4.0629
		16	3.8766	3.8767	2.0576	2.0551	3.8767	3.8769	3.8999	2.0573	2.0548	3.8768	3.8771	3.9069	2.0572	2.0547	2.0545	2.0553
	0.4	4	4.0299	4.0300	3.8014	3.7573	4.0300	4.0303	4.0368	3.8011	3.7564	4.0300	4.0304	4.0360	3.8011	3.7562	3.7637	3.7583
Average		16	3.7523	3.7525	1.9056	1.9035	3.7525	3.7527	3.7517	1.9055	1.9032	3.7525	3.7532	3.7487	1.9055	1.9032	1.9046	1.9018
Average	0.8	4	4.1226	4.1226	3.9033	3.8567	4.1222	4.1223	4.1121	3.9025	3.8523	4.1222	4.1225	4.1064	3.9024	3.8519	3.8465	3.8370
		16	3.7610	3.7611	1.9446	1.9422	3.7611	3.7614	3.7885	1.9444	1.9419	3.7612	3.7616	3.7989	1.9444	1.9418	1.9415	1.9421

Table 4 - Forecasts RMSE - (N,T)=(50,10), SMA data generating process for φ, W(5,5), 1000 replications

										Estima	ators							
			Pooled	Av. hetero.			Pooled SAR	Av. hete	ro. SAR	FE-SAR	RE-SAR	Pooled SMA	Av. hete	ero. SMA	FE-SMA	RE-SMA	SAR-RE	SMA-RE
	ρ	$\sigma_{_{\!$	OLS	OLS	FE	RE	MLE	MLE	GM	MLE	MLE	MLE	MLE	GM	MLE	MLE	GM	GM
	0.4	4	3.9511	3.9513	3.7364	3.7155	3.9516	3.9516	3.9631	3.7358	3.7105	3.9516	3.9517	3.9672	3.7358	3.7096	3.7289	3.7116
1st voar		16	3.6164	3.6164	1.8633	1.8625	3.6180	3.6180	3.6429	1.8629	1.8618	3.6182	3.6183	3.6441	1.8629	1.8618	1.8654	1.8654
ist year	0.8	4	7.0636	7.0637	7.0399	6.9797	7.0634	7.0543	7.1286	7.0359	6.9436	7.0642	7.0551	7.1390	7.0356	6.9442	6.9870	6.9751
		16	4.6604	4.6608	3.5080	3.5046	4.6610	4.6607	4.7277	3.5066	3.4980	4.6604	4.6601	4.7194	3.5067	3.4997	3.4914	3.4981
	0.4	4	4.4075	4.4079	4.1572	4.1384	4.4081	4.4083	4.3916	4.1563	4.1335	4.4081	4.4084	4.3948	4.1563	4.1329	4.1372	4.1361
2nd year		16	3.8526	3.8526	2.0675	2.0667	3.8538	3.8539	3.8684	2.0672	2.0662	3.8539	3.8541	3.8755	2.0673	2.0663	2.0692	2.0659
znu year	0.8	4	7.8524	7.8523	7.8121	7.7512	7.8511	7.8442	7.8893	7.8082	7.7158	7.8530	7.8459	7.9029	7.8083	7.7174	7.7551	7.7475
		16	5.1228	5.1230	3.9200	3.9160	5.1233	5.1228	5.1328	3.9180	3.9080	5.1233	5.1228	5.1272	3.9180	3.9100	3.8881	3.9055
	0.4	4	4.5841	4.5843	4.3239	4.3050	4.5846	4.5847	4.5668	4.3234	4.2998	4.5847	4.5849	4.5692	4.3234	4.2997	4.3076	4.3034
3rd year		16	3.9438	3.9438	2.1507	2.1500	3.9446	3.9447	3.9571	2.1504	2.1495	3.9447	3.9448	3.9611	2.1504	2.1495	2.1549	2.1467
Siù year	0.8	4	8.1797	8.1796	8.1425	8.0788	8.1789	8.1712	8.1870	8.1380	8.0444	8.1804	8.1727	8.1969	8.1380	8.0442	8.1139	8.0629
		16	5.2836	5.2838	4.0774	4.0722	5.2842	5.2837	5.3024	4.0753	4.0618	5.2844	5.2839	5.2963	4.0754	4.0656	4.0413	4.0582
	0.4	4	4.6767	4.6769	4.4123	4.3931	4.6772	4.6773	4.6529	4.4118	4.3881	4.6773	4.6775	4.6563	4.4118	4.3879	4.3893	4.3832
4th year		16	3.9904	3.9904	2.1964	2.1957	3.9916	3.9916	4.0038	2.1961	2.1953	3.9917	3.9918	4.0078	2.1961	2.1953	2.1976	2.1898
4iii yeai	0.8	4	8.3518	8.3519	8.3136	8.2480	8.3510	8.3435	8.3496	8.3097	8.2129	8.3527	8.3452	8.3546	8.3097	8.2123	8.2705	8.2339
		16	5.3737	5.3739	4.1635	4.1581	5.3748	5.3741	5.3871	4.1614	4.1466	5.3750	5.3743	5.3838	4.1615	4.1512	4.1180	4.1433
	0.4	4	4.7296	4.7298	4.4640	4.4440	4.7300	4.7302	4.7138	4.4635	4.4388	4.7302	4.7303	4.7164	4.4635	4.4386	4.4428	4.4382
5th year		16	4.0171	4.0171	2.2227	2.2221	4.0185	4.0185	4.0283	2.2224	2.2216	4.0186	4.0186	4.0318	2.2224	2.2216	2.2232	2.2189
Still year	0.8	4	8.4459	8.4460	8.4041	8.3400	8.4451	8.4372	8.4425	8.4005	8.3036	8.4469	8.4390	8.4449	8.4006	8.3035	8.3642	8.3287
		16	5.4261	5.4263	4.2084	4.2034	5.4281	5.4273	5.4429	4.2062	4.1923	5.4284	5.4276	5.4431	4.2063	4.1964	4.1699	4.1907
	0.4	4	4.4698	4.4700	4.2188	4.1992	4.4703	4.4704	4.4576	4.2182	4.1941	4.4704	4.4706	4.4604	4.2182	4.1937	4.2012	4.1945
Avorago		16	3.8840	3.8841	2.1001	2.0994	3.8853	3.8853	3.9001	2.0998	2.0989	3.8854	3.8855	3.9041	2.0998	2.0989	2.1021	2.0974
Average -	0.8	4	7.9787	7.9787	7.9424	7.8796	7.9779	7.9701	7.9994	7.9385	7.8441	7.9794	7.9716	8.0077	7.9384	7.8443	7.8981	7.8696
		16	5.1733	5.1736	3.9755	3.9709	5.1743	5.1737	5.1986	3.9735	3.9614	5.1743	5.1737	5.1940	3.9736	3.9646	3.9417	3.9592

Table 5 - Forecasts RMSE - (N,T)=(50,20), SAR data generating process for φ, W(1,1), 1000 replications

										Estima	ators							
			Pooled	Av. hetero.		DE	Pooled SAR	Av. hete	ro. SAR	FE-SAR	RE-SAR	Pooled SMA	Av. hete	ero. SMA	FE-SMA	RE-SMA	SAR-RE	SMA-RE
	λ	σ_{μ}^{2}	OLS	OLS	FE	RE	MLE	MLE	GM	MLE	MLE	MLE	MLE	GM	MLE	MLE	GM	GM
	0.4	4	3.6699	3.6698	3.3999	3.3863	3.6705	3.6704	3.6873	3.3994	3.3834	3.6706	3.6707	3.6803	3.3994	3.3828	3.3885	3.3517
1st vear		16	3.5573	3.5575	1.6920	1.6915	3.5575	3.5576	3.5501	1.6917	1.6911	3.5574	3.5575	3.5384	1.6917	1.6910	1.7044	1.6870
ist year	0.8	4	3.9856	3.9856	3.7521	3.7370	3.9854	3.9854	3.9823	3.7505	3.7291	3.9860	3.9858	3.9883	3.7502	3.7282	3.7417	3.7272
		16	3.6376	3.6376	1.8802	1.8793	3.6387	3.6387	3.6182	1.8795	1.8778	3.6388	3.6389	3.6226	1.8794	1.8777	1.8990	1.8689
	0.4	4	4.0666	4.0666	3.7765	3.7606	4.0673	4.0673	4.0754	3.7760	3.7571	4.0675	4.0676	4.0721	3.7759	3.7564	3.7690	3.7332
2nd year		16	3.7723	3.7723	1.8861	1.8856	3.7721	3.7724	3.7691	1.8857	1.8851	3.7721	3.7724	3.7580	1.8857	1.8850	1.8916	1.8799
Zhu year	0.8	4	4.4240	4.4241	4.1711	4.1538	4.4242	4.4242	4.4219	4.1698	4.1437	4.4251	4.4251	4.4266	4.1695	4.1428	4.1549	4.1537
		16	3.8756	3.8757	2.0870	2.0860	3.8777	3.8777	3.8650	2.0861	2.0845	3.8779	3.8779	3.8708	2.0860	2.0841	2.0938	2.0732
	0.4	4	4.2243	4.2243	3.9257	3.9098	4.2250	4.2250	4.2282	3.9252	3.9064	4.2252	4.2254	4.2223	3.9251	3.9058	3.9138	3.8949
ard year		16	3.8441	3.8443	1.9592	1.9584	3.8440	3.8443	3.8485	1.9589	1.9579	3.8440	3.8443	3.8393	1.9589	1.9579	1.9619	1.9554
Siù yeai	0.8	4	4.5929	4.5929	4.3315	4.3135	4.5932	4.5932	4.5947	4.3300	4.3043	4.5945	4.5944	4.5975	4.3297	4.3027	4.3126	4.3234
		16	3.9644	3.9645	2.1704	2.1695	3.9665	3.9665	3.9558	2.1696	2.1682	3.9667	3.9667	3.9574	2.1695	2.1678	2.1776	2.1619
	0.4	4	4.3108	4.3109	4.0036	3.9883	4.3114	4.3114	4.3064	4.0030	3.9852	4.3116	4.3117	4.3003	4.0030	3.9846	3.9869	3.9818
4th year		16	3.8849	3.8850	2.0007	1.9999	3.8847	3.8850	3.8836	2.0005	1.9994	3.8848	3.8850	3.8752	2.0004	1.9994	2.0022	1.9994
4ili yeai	0.8	4	4.6780	4.6781	4.4134	4.3953	4.6789	4.6789	4.6829	4.4122	4.3857	4.6802	4.6801	4.6843	4.4120	4.3843	4.4054	4.4078
		16	4.0090	4.0091	2.2137	2.2129	4.0108	4.0108	4.0012	2.2129	2.2117	4.0110	4.0111	4.0032	2.2127	2.2113	2.2213	2.2041
	0.4	4	4.3662	4.3663	4.0529	4.0382	4.3667	4.3666	4.3602	4.0524	4.0357	4.3669	4.3670	4.3537	4.0524	4.0351	4.0411	4.0360
Ethylogr		16	3.9107	3.9109	2.0252	2.0245	3.9107	3.9110	3.9069	2.0249	2.0240	3.9107	3.9110	3.8986	2.0249	2.0240	2.0274	2.0254
Sin year	0.8	4	4.7359	4.7359	4.4707	4.4520	4.7371	4.7370	4.7369	4.4692	4.4429	4.7385	4.7383	4.7396	4.4689	4.4412	4.4623	4.4588
		16	4.0418	4.0419	2.2428	2.2421	4.0434	4.0434	4.0313	2.2420	2.2412	4.0436	4.0436	4.0317	2.2418	2.2407	2.2496	2.2321
	0.4	4	4.1276	4.1276	3.8317	3.8167	4.1282	4.1282	4.1315	3.8312	3.8135	4.1283	4.1285	4.1257	3.8312	3.8129	3.8199	3.7995
Average		16	3.7939	3.7940	1.9127	1.9120	3.7938	3.7940	3.7916	1.9124	1.9115	3.7938	3.7940	3.7819	1.9123	1.9115	1.9175	1.9094
Average	0.8	4	4.4832	4.4833	4.2277	4.2103	4.4838	4.4837	4.4837	4.2263	4.2012	4.4849	4.4847	4.4873	4.2260	4.1998	4.2154	4.2142
		16	3.9057	3.9057	2.1188	2.1180	3.9074	3.9074	3.8943	2.1180	2.1167	3.9076	3.9076	3.8971	2.1179	2.1163	2.1282	2.1081

Table 6 - Forecasts RMSE - (N,T)=(50,20), SMA data generating process for φ, W(1,1), 1000 replications

										Estima	ators							
	true		Pooled	Av. hetero.			Pooled SAR	Av. hete	ro. SAR	FE-SAR	RE-SAR	Pooled SMA	Av. hete	ro. SMA	FE-SMA	RE-SMA	SAR-RE	SMA-RE
	DGP	$\sigma_{_{\!$	OLS	OLS	FE	KE	MLE	MLE	GM	MLE	MLE	MLE	MLE	GM	MLE	MLE	GM	GM
	SAD	4	6.9985	6.9985	7.1226	6.9431	6.9975	6.9935	7.1114	7.1190	6.8901	6.9986	6.9947	7.0931	7.1193	6.8938	7.0744	7.0337
1ct voor	SAR	16	4.6503	4.6503	3.5892	3.5704	4.6506	4.6505	4.6766	3.5874	3.5416	4.6506	4.6505	4.4448	3.5873	3.5529	3.5902	3.5771
ist year	SMA	4	4.0103	4.0102	3.8592	3.8002	4.0100	4.0100	4.0062	3.8581	3.7799	4.0108	4.0107	3.9849	3.8580	3.7780	3.7920	3.7807
	SIVIA	16	3.6578	3.6578	1.9253	1.9221	3.6578	3.6578	3.6689	1.9242	1.9191	3.6578	3.6578	3.6573	1.9238	1.9182	1.9241	1.9189
	SVD	4	7.8090	7.8090	7.9482	7.7505	7.8072	7.8036	7.8799	7.9446	7.6914	7.8083	7.8048	7.8602	7.9449	7.6966	7.8694	7.7965
2nd year	SAN	16	5.1067	5.1067	4.0015	3.9818	5.1072	5.1071	5.1124	3.9998	3.9529	5.1072	5.1071	5.1245	3.9998	3.9641	4.0009	3.9824
znu year	SMA	4	4.4542	4.4542	4.2866	4.2246	4.4542	4.4542	4.4428	4.2851	4.2039	4.4550	4.4550	4.4326	4.2849	4.2010	4.2212	4.2173
	SIVIA	16	3.9015	3.9015	2.1390	2.1358	3.9020	3.9020	3.9049	2.1379	2.1329	3.9019	3.9020	3.9004	2.1376	2.1321	2.1268	2.1282
	SAP	4	8.1109	8.1110	8.2604	8.0514	8.1097	8.1061	8.1691	8.2567	7.9908	8.1108	8.1074	8.1531	8.2570	7.9965	8.1481	8.1222
3rd year	SAR	16	5.2802	5.2802	4.1566	4.1380	5.2810	5.2808	5.2830	4.1548	4.1092	5.2811	5.2809	5.2993	4.1548	4.1202	4.1436	4.1359
Siù yeai	SMA	4	4.6117	4.6118	4.4413	4.3773	4.6119	4.6119	4.6106	4.4396	4.3562	4.6127	4.6127	4.5981	4.4394	4.3533	4.3891	4.3767
	SIVIA	16	3.9868	3.9869	2.2197	2.2165	3.9872	3.9873	3.9899	2.2187	2.2138	3.9872	3.9873	3.9852	2.2185	2.2129	2.2159	2.2105
	SAP	4	8.2880	8.2881	8.4361	8.2247	8.2863	8.2825	8.3401	8.4323	8.1646	8.2872	8.2836	8.3265	8.4325	8.1699	8.3174	8.2757
Ath year	SAR	16	5.3754	5.3754	4.2379	4.2203	5.3764	5.3762	5.3696	4.2361	4.1923	5.3764	5.3763	5.3841	4.2362	4.2025	4.2274	4.2159
4iii yeai	SMA	4	4.7001	4.7001	4.5263	4.4619	4.7002	4.7002	4.6970	4.5246	4.4406	4.7011	4.7012	4.6862	4.5243	4.4375	4.4753	4.4638
	SIVIA	16	4.0305	4.0306	2.2643	2.2611	4.0310	4.0311	4.0350	2.2634	2.2584	4.0310	4.0310	4.0261	2.2632	2.2576	2.2595	2.2564
	SAR	4	8.4030	8.4031	8.5526	8.3403	8.4015	8.3972	8.4337	8.5489	8.2779	8.4025	8.3984	8.4197	8.5491	8.2847	8.4158	8.3736
5th year	OAN	16	5.4232	5.4232	4.2810	4.2636	5.4244	5.4242	5.4204	4.2791	4.2362	5.4243	5.4242	5.4327	4.2792	4.2459	4.2825	4.2742
Stri year	SMA	4	4.7576	4.7576	4.5825	4.5177	4.7578	4.7578	4.7505	4.5806	4.4958	4.7588	4.7588	4.7481	4.5804	4.4929	4.5331	4.5198
	OWA	16	4.0625	4.0625	2.2950	2.2919	4.0630	4.0630	4.0607	2.2940	2.2894	4.0629	4.0630	4.0515	2.2938	2.2886	2.2879	2.2826
	SAR	4	7.9219	7.9219	8.0640	7.8620	7.9204	7.9166	7.9868	8.0603	7.8029	7.9215	7.9178	7.9705	8.0606	7.8083	7.9650	7.9203
	OAK	16	5.1672	5.1672	4.0533	4.0348	5.1679	5.1678	5.1724	4.0514	4.0064	5.1679	5.1678	5.1815	4.0515	4.0171	4.0489	4.0371
Average	SMA	4	4.5068	4.5068	4.3392	4.2764	4.5068	4.5068	4.5014	4.3376	4.2553	4.5077	4.5077	4.4900	4.3374	4.2525	4.2822	4.2717
	SIVIA	16	3.9278	3.9278	2.1687	2.1655	3.9282	3.9282	3.9319	2.1677	2.1627	3.9282	3.9282	3.9241	2.1674	2.1619	2.1629	2.1593

Table 7 - Forecasts RMSE - (N,T)=(100,10), W(1,1), 1000 replications $\rho=\lambda=0.8$ for SAR and SMA data generating processes

										Estim	ators							
	true		Pooled	Av. hetero.			Pooled SAR	Av. hete	ro. SAR	FE-SAR	RE-SAR	Pooled SMA	Av. hete	ro. SMA	FE-SMA	RE-SMA	SAR-RE	SMA-RE
	DGP	σ_{μ}^{2}	OLS	OLS	FE	RE	MLE	MLE	GM	MLE	MLE	MLE	MLE	GM	MLE	MLE	GM	GM
	SVD	4	4.8896	4.8896	4.8105	4.7424	4.8891	4.8842	4.8756	4.8090	4.7109	4.8891	4.8843	4.8631	4.8090	4.7104	4.7175	4.7375
1 et voor	SAN	16	3.9142	3.9142	2.3915	2.3876	3.9141	3.9141	3.9341	2.3908	2.3731	3.9142	3.9142	3.9071	2.3907	2.3804	2.3936	2.3766
ist year	SMA	4	3.6481	3.6481	3.4526	3.4061	3.6482	3.6481	3.6569	3.4522	3.3998	3.6482	3.6481	3.6532	3.4521	3.3997	3.3988	3.4030
	SIVIA	16	3.5641	3.5641	1.7221	1.7201	3.5641	3.5641	3.5640	1.7219	1.7193	3.5641	3.5641	3.5677	1.7219	1.7192	1.7239	1.7176
	SAB	4	5.4287	5.4286	5.3436	5.2695	5.4277	5.4231	5.4291	5.3420	5.2336	5.4277	5.4233	5.4096	5.3420	5.2341	5.2488	5.2694
2nd year	SAN	16	4.2142	4.2142	2.6762	2.6712	4.2141	4.2141	4.2449	2.6756	2.6548	4.2142	4.2141	4.2210	2.6755	2.6633	2.6632	2.6446
znu year	SMA	4	4.0344	4.0344	3.8254	3.7736	4.0346	4.0345	4.0495	3.8250	3.7673	4.0347	4.0346	4.0555	3.8250	3.7669	3.7881	3.7818
	SINIA	16	3.7822	3.7822	1.9174	1.9155	3.7823	3.7823	3.7733	1.9172	1.9148	3.7824	3.7824	3.7828	1.9172	1.9147	1.9158	1.9132
	SAB	4	5.6536	5.6534	5.5673	5.4915	5.6526	5.6479	5.6420	5.5659	5.4550	5.6526	5.6481	5.6318	5.5659	5.4552	5.4511	5.4916
3rd year	SAN	16	4.3316	4.3317	2.7876	2.7825	4.3317	4.3317	4.3494	2.7866	2.7673	4.3318	4.3318	4.3274	2.7866	2.7750	2.7642	2.7511
Siù yeai	SMA	4	4.1902	4.1902	3.9751	3.9225	4.1902	4.1902	4.2048	3.9747	3.9163	4.1903	4.1903	4.2137	3.9747	3.9160	3.9417	3.9364
	SINIA	16	3.8599	3.8599	1.9944	1.9925	3.8600	3.8600	3.8512	1.9942	1.9919	3.8600	3.8600	3.8547	1.9942	1.9918	1.9915	1.9880
	SAB	4	5.7652	5.7651	5.6767	5.6001	5.7643	5.7593	5.7521	5.6755	5.5624	5.7643	5.7593	5.7516	5.6754	5.5632	5.5707	5.6067
Ath year	SAN	16	4.3864	4.3865	2.8394	2.8342	4.3863	4.3863	4.4067	2.8383	2.8181	4.3864	4.3864	4.3796	2.8383	2.8263	2.8198	2.8078
Hillyear	SMA	4	4.2767	4.2767	4.0579	4.0051	4.2768	4.2768	4.2908	4.0575	3.9988	4.2769	4.2769	4.3000	4.0575	3.9984	4.0228	4.0188
	SINIA	16	3.8979	3.8979	2.0318	2.0300	3.8981	3.8981	3.8881	2.0316	2.0294	3.8981	3.8981	3.8923	2.0316	2.0293	2.0310	2.0292
	SAR	4	5.8347	5.8346	5.7504	5.6716	5.8339	5.8290	5.8163	5.7492	5.6306	5.8338	5.8291	5.8164	5.7491	5.6329	5.6386	5.6769
5th year	UAN	16	4.4237	4.4238	2.8759	2.8707	4.4238	4.4238	4.4366	2.8747	2.8544	4.4238	4.4238	4.4162	2.8748	2.8628	2.8540	2.8424
ourycar	SMA	4	4.3286	4.3287	4.1070	4.0542	4.3287	4.3287	4.3442	4.1066	4.0478	4.3288	4.3288	4.3505	4.1066	4.0474	4.0702	4.0662
	OWIA	16	3.9212	3.9212	2.0574	2.0554	3.9213	3.9214	3.9125	2.0572	2.0547	3.9214	3.9214	3.9139	2.0572	2.0546	2.0560	2.0555
	SAR	4	5.5144	5.5142	5.4297	5.3550	5.5135	5.5087	5.5030	5.4283	5.3185	5.5135	5.5088	5.4945	5.4283	5.3192	5.3254	5.3564
	UAN	16	4.2540	4.2541	2.7141	2.7093	4.2540	4.2540	4.2743	2.7132	2.6935	4.2541	4.2541	4.2502	2.7132	2.7016	2.6990	2.6845
Average	SMA	4	4.0956	4.0956	3.8836	3.8323	4.0957	4.0957	4.1093	3.8832	3.8260	4.0958	4.0957	4.1146	3.8832	3.8257	3.8443	3.8412
		16	3.8051	3.8051	1.9446	1.9427	3.8052	3.8052	3.7978	1.9444	1.9420	3.8052	3.8052	3.8023	1.9444	1.9419	1.9436	1.9407

Table 8 - Forecasts RMSE - (N,T)=(100,10), W(5,5), 1000 replications $\rho=\lambda=0.8$ for SAR and SMA data generating processes

										Estima	tors							
			Pooled	Av. hetero.			Pooled SAR	Av. hete	ro. SAR	FE-SAR	RE-SAR	Pooled SMA	Av. hete	ro. SMA	FE-SMA	RE-SMA	SAR-RE	SMA-RE
	ρ	$\sigma_{_{\mu}}^{_{2}}$	OLS	OLS	FE	KE	MLE	MLE	GM	MLE	MLE	MLE	MLE	GM	MLE	MLE	GM	GM
	0.4	4	4.1551	4.1552	4.0358	3.9706	4.1563	4.1564	4.2084	4.0338	3.9550	4.1561	4.1562	4.1772	4.0341	3.9588	3.9718	4.0408
1st vear		16	3.6742	3.6743	2.0307	2.0268	3.6747	3.6746	3.7101	2.0297	2.0240	3.6742	3.6745	3.6869	2.0297	2.0251	2.0188	1.7766
i ot your	0.8	4	9.9810	9.9856	9.3815	9.9337	10.0933	10.1004	9.5173	9.3512	9.9918	10.0797	10.0493	11.0071	9.3561	9.5768	14.1056	11.9221
		16	6.0600	6.0600	5.4403	5.3955	6.0623	6.0588	6.0347	5.4324	5.3177	6.0619	6.0607	5.9938	5.4322	5.3498	5.3437	5.9137
	0.4	4	4.6207	4.6208	4.4863	4.4181	4.6216	4.6217	4.6514	4.4840	4.4016	4.6212	4.6213	4.6323	4.4842	4.4058	4.4093	4.5255
2nd year		16	3.9151	3.9152	2.2464	2.2423	3.9160	3.9158	3.9599	2.2455	2.2397	3.9155	3.9157	3.9448	2.2456	2.2411	2.2401	2.0398
Zna year	0.8	4	11.1618	11.1681	11.4312	11.1342	11.1617	11.1559	10.6685	11.3428	11.4194	11.1554	11.1523	13.2402	11.3491	11.0807	14.8653	12.0044
		16	6.6797	6.6797	5.9939	5.9463	6.6853	6.6811	6.6860	5.9864	5.8707	6.6823	6.6810	6.6818	5.9864	5.9005	5.9473	6.9897
	0.4	4	4.8019	4.8020	4.6608	4.5919	4.8022	4.8024	4.8187	4.6587	4.5759	4.8019	4.8020	4.8093	4.6589	4.5792	4.5873	4.6235
3rd year		16	4.0085	4.0086	2.3379	2.3335	4.0098	4.0096	4.0529	2.3368	2.3305	4.0094	4.0095	4.0385	2.3368	2.3322	2.3302	2.2916
Siù yeai	0.8	4	12.5896	12.5954	13.0581	12.5818	12.6034	12.6015	11.1090	13.0022	12.8293	12.5959	12.5869	13.5612	13.0067	12.6192	14.0236	11.8049
		16	6.9318	6.9319	6.2300	6.1842	6.9396	6.9351	6.9420	6.2227	6.1073	6.9354	6.9344	6.9338	6.2228	6.1366	6.2060	6.5267
	0.4	4	4.9024	4.9025	4.7604	4.6903	4.9028	4.9029	4.9106	4.7582	4.6746	4.9027	4.9029	4.9074	4.7585	4.6782	4.6926	4.8674
4th vear		16	4.0589	4.0591	2.3835	2.3794	4.0601	4.0599	4.0965	2.3825	2.3767	4.0597	4.0598	4.0885	2.3825	2.3782	2.3710	2.4283
Hirycar	0.8	4	13.0269	13.0318	13.3271	13.0157	13.0225	13.0203	12.1625	13.3052	13.1667	13.0220	13.0217	13.7779	13.3076	12.9671	14.0499	12.3835
		16	7.0667	7.0667	6.3577	6.3117	7.0745	7.0700	7.0763	6.3503	6.2328	7.0698	7.0689	7.0685	6.3505	6.2637	6.3310	6.3103
	0.4	4	4.9614	4.9616	4.8170	4.7469	4.9618	4.9620	4.9720	4.8148	4.7308	4.9617	4.9619	4.9667	4.8151	4.7346	4.7497	4.8157
5th year		16	4.0865	4.0866	2.4125	2.4084	4.0876	4.0875	4.1232	2.4114	2.4058	4.0872	4.0874	4.1151	2.4114	2.4071	2.4017	2.5225
our year	0.8	4	13.4890	13.4947	13.7509	13.4767	13.4764	13.4771	13.2666	13.7220	13.5829	13.4730	13.4675	14.1849	13.7242	13.4160	13.9525	12.2767
		16	7.1538	7.1538	6.4418	6.3950	7.1625	7.1581	7.1581	6.4340	6.3148	7.1575	7.1566	7.1510	6.4342	6.3459	6.4185	6.3724
	0.4	4	4.6883	4.6884	4.5521	4.4836	4.6889	4.6891	4.7122	4.5499	4.4676	4.6887	4.6889	4.6986	4.5502	4.4713	4.4821	4.5746
Average		16	3.9486	3.9488	2.2822	2.2781	3.9496	3.9495	3.9885	2.2812	2.2753	3.9492	3.9494	3.9748	2.2812	2.2768	2.2724	2.2118
Average	0.8	4	12.0497	12.0551	12.1897	12.0284	12.0715	12.0710	11.3448	12.1447	12.1980	12.0652	12.0556	13.1542	12.1487	11.9319	14.1994	12.0783
		16	6.7784	6.7784	6.0928	6.0465	6.7848	6.7806	6.7794	6.0851	5.9687	6.7814	6.7803	6.7658	6.0852	5.9993	6.0493	6.4226

Table 9 - Forecasts RMSE - (N,T)=(50,10), SAR data generating process for ϕ , W(1,1) asymmetric weight matrix of French administrative communes, 1000 replications

										Estima	tors							
			Pooled	Av. hetero.		DE	Pooled SAR	Av. hete	ro. SAR	FE-SAR	RE-SAR	Pooled SMA	Av. hete	ero. SMA	FE-SMA	RE-SMA	SAR-RE	SMA-RE
	λ	σ_{μ}^{2}	OLS	OLS	FE	RE	MLE	MLE	GM	MLE	MLE	MLE	MLE	GM	MLE	MLE	GM	GM
	0.4	4	3.7809	3.7808	3.5814	3.5379	3.7822	3.7821	3.8043	3.5791	3.5262	3.7831	3.7829	3.7746	3.5791	3.5277	3.5663	3.5515
1st vear		16	3.5887	3.5886	1.8019	1.8003	3.5895	3.5894	3.6027	1.8010	1.7993	3.5901	3.5897	3.5790	1.8010	1.7993	1.8095	1.8088
iot your	0.8	4	4.3674	4.3676	4.2789	4.1978	4.3665	4.3666	4.3992	4.2743	4.1715	4.3663	4.3664	4.3773	4.2727	4.1660	4.2319	4.2213
		16	3.7471	3.7470	2.1479	2.1440	3.7489	3.7487	3.7494	2.1466	2.1407	3.7492	3.7484	3.7325	2.1463	2.1395	2.1355	2.1329
	0.4	4	4.1973	4.1973	3.9972	3.9465	4.1982	4.1985	4.2122	3.9952	3.9346	4.1988	4.1989	4.1837	3.9952	3.9355	3.9610	3.9549
2nd year		16	3.8053	3.8053	1.9898	1.9884	3.8064	3.8065	3.8219	1.9892	1.9874	3.8072	3.8066	3.8084	1.9892	1.9874	2.0021	2.0062
Zild year	0.8	4	4.8689	4.8690	4.7680	4.6864	4.8694	4.8693	4.8883	4.7639	4.6591	4.8690	4.8690	4.8525	4.7622	4.6534	4.7003	4.6877
		16	4.0193	4.0193	2.3854	2.3817	4.0204	4.0204	4.0244	2.3835	2.3781	4.0206	4.0201	4.0001	2.3832	2.3770	2.3821	2.3708
	0.4	4	4.3641	4.3642	4.1645	4.1114	4.3649	4.3653	4.3629	4.1625	4.0999	4.3657	4.3657	4.3599	4.1624	4.1003	4.1042	4.1069
3rd year		16	3.8871	3.8872	2.0752	2.0738	3.8877	3.8879	3.8996	2.0745	2.0727	3.8884	3.8880	3.8903	2.0744	2.0725	2.0785	2.0819
Siù yeai	0.8	4	5.0644	5.0644	4.9605	4.8799	5.0650	5.0649	5.0699	4.9569	4.8509	5.0647	5.0647	5.0435	4.9556	4.8451	4.8845	4.8738
		16	4.1141	4.1141	2.4801	2.4761	4.1157	4.1156	4.1222	2.4781	2.4720	4.1159	4.1155	4.0955	2.4777	2.4708	2.4713	2.4660
	0.4	4	4.4458	4.4459	4.2439	4.1899	4.4471	4.4474	4.4539	4.2419	4.1787	4.4478	4.4477	4.4435	4.2418	4.1791	4.1901	4.1942
Ath year		16	3.9249	3.9250	2.1216	2.1196	3.9255	3.9258	3.9445	2.1209	2.1183	3.9262	3.9259	3.9344	2.1208	2.1182	2.1195	2.1222
4iii yeai	0.8	4	5.1625	5.1626	5.0576	4.9749	5.1629	5.1628	5.1742	5.0533	4.9453	5.1631	5.1631	5.1466	5.0519	4.9399	4.9794	4.9755
		16	4.1643	4.1644	2.5301	2.5259	4.1661	4.1661	4.1721	2.5278	2.5212	4.1664	4.1660	4.1469	2.5272	2.5200	2.5214	2.5224
	0.4	4	4.4976	4.4977	4.2964	4.2411	4.4986	4.4989	4.5052	4.2943	4.2302	4.4992	4.4992	4.5008	4.2943	4.2305	4.2456	4.2445
5th year		16	3.9472	3.9473	2.1446	2.1425	3.9478	3.9480	3.9653	2.1438	2.1411	3.9484	3.9481	3.9564	2.1438	2.1411	2.1458	2.1490
Jul year	0.8	4	5.2171	5.2172	5.1101	5.0272	5.2179	5.2177	5.2409	5.1059	4.9979	5.2184	5.2184	5.2129	5.1047	4.9916	5.0386	5.0420
		16	4.1913	4.1913	2.5593	2.5551	4.1929	4.1929	4.2007	2.5570	2.5505	4.1932	4.1928	4.1792	2.5565	2.5493	2.5579	2.5532
	0.4	4	4.2571	4.2572	4.0567	4.0053	4.2582	4.2584	4.2677	4.0546	3.9939	4.2589	4.2589	4.2525	4.0546	3.9946	4.0135	4.0104
Average		16	3.8307	3.8307	2.0266	2.0249	3.8314	3.8315	3.8468	2.0259	2.0238	3.8321	3.8317	3.8337	2.0258	2.0237	2.0311	2.0336
Average	0.8	4	4.9361	4.9362	4.8350	4.7532	4.9363	4.9362	4.9545	4.8309	4.7249	4.9363	4.9363	4.9266	4.8294	4.7192	4.7669	4.7601
		16	4.0472	4.0472	2.4206	2.4166	4.0488	4.0488	4.0538	2.4186	2.4125	4.0491	4.0486	4.0308	2.4182	2.4113	2.4137	2.4090

Table 10 - Forecasts RMSE - (N,T)=(50,10), SMA data generating process for φ, W(1,1) asymmetric weight matrix of French administrative communes, 1000 replications

										Estim	ators							
			Pooled	Av. hetero.		DE	Pooled SAR	Av. hete	ro. SAR	FE-SAR	RE-SAR	Pooled SMA	Av. hete	ro. SMA	FE-SMA	RE-SMA	SAR-RE	SMA-RE
	ρ	σ_{μ}^{2}	OLS	OLS	FE	RE	MLE	MLE	GM	MLE	MLE	MLE	MLE	GM	MLE	MLE	GM	GM
	0.4	4	3.7378	3.7381	3.5505	3.5077	3.7375	3.7376	3.7345	3.5494	3.5024	3.7375	3.7377	3.7351	3.5494	3.5008	3.5002	3.7998
1st vear		16	3.5580	3.5581	1.7728	1.7699	3.5603	3.5595	3.5726	1.7726	1.7690	3.5589	3.5593	3.5620	1.7726	1.7692	1.7776	1.9072
ist year	0.8	4	5.7212	5.7210	5.6986	5.6317	5.7122	5.6908	5.7728	5.6943	5.6694	5.7170	5.7194	5.6710	5.6944	5.5835	5.7257	6.5819
		16	4.1824	4.1824	2.8355	2.8319	4.1857	4.1826	4.1996	2.8337	2.8071	4.1829	4.1819	4.2006	2.8338	2.8200	2.8318	3.3513
	0.4	4	4.1233	4.1237	3.9270	3.8750	4.1234	4.1239	4.1296	3.9263	3.8712	4.1232	4.1236	4.1446	3.9264	3.8691	3.8801	3.9050
2nd year		16	3.7876	3.7877	1.9771	1.9748	3.7894	3.7887	3.7993	1.9769	1.9741	3.7883	3.7887	3.7848	1.9769	1.9745	1.9719	2.0943
Zhu you	0.8	4	6.3676	6.3677	6.3403	6.2672	6.3531	6.3335	6.3887	6.3362	6.3238	6.3609	6.3589	6.3212	6.3364	6.2137	6.3270	5.7960
		16	4.5451	4.5452	3.1584	3.1544	4.5496	4.5451	4.5415	3.1565	3.1296	4.5461	4.5448	4.5448	3.1566	3.1418	3.1552	3.6714
	0.4	4	4.2904	4.2906	4.0876	4.0360	4.2905	4.2908	4.2843	4.0868	4.0321	4.2904	4.2905	4.3064	4.0869	4.0301	4.0363	4.1751
3rd yoar		16	3.8611	3.8612	2.0518	2.0494	3.8630	3.8622	3.8770	2.0516	2.0485	3.8618	3.8622	3.8666	2.0516	2.0490	2.0452	2.0835
Siù yeai	0.8	4	6.6165	6.6168	6.5966	6.5164	6.6021	6.5814	6.6730	6.5927	6.5671	6.6102	6.6083	6.5702	6.5928	6.4634	6.5194	6.3848
		16	4.6793	4.6792	3.2918	3.2870	4.6832	4.6791	4.6847	3.2897	3.2589	4.6803	4.6788	4.6760	3.2898	3.2735	3.2866	3.3368
	0.4	4	4.3835	4.3838	4.1786	4.1261	4.3837	4.3840	4.3727	4.1777	4.1220	4.3836	4.3838	4.3891	4.1778	4.1205	4.1130	4.0951
Ath year		16	3.9001	3.9002	2.0885	2.0863	3.9021	3.9014	3.9155	2.0883	2.0854	3.9009	3.9014	3.9083	2.0883	2.0859	2.0873	2.2285
4iii yeai	0.8	4	6.7475	6.7479	6.7260	6.6453	6.7361	6.7144	6.7928	6.7211	6.6999	6.7415	6.7384	6.7232	6.7213	6.5918	6.6467	6.6437
		16	4.7503	4.7503	3.3568	3.3523	4.7549	4.7504	4.7473	3.3547	3.3262	4.7519	4.7503	4.7426	3.3548	3.3394	3.3510	3.2717
	0.4	4	4.4334	4.4337	4.2277	4.1747	4.4336	4.4338	4.4326	4.2268	4.1704	4.4335	4.4336	4.4434	4.2268	4.1689	4.1664	4.2339
5th year		16	3.9201	3.9203	2.1117	2.1092	3.9221	3.9214	3.9399	2.1115	2.1084	3.9210	3.9214	3.9338	2.1115	2.1089	2.1097	2.2687
Jul year	0.8	4	6.8084	6.8087	6.7852	6.7061	6.7956	6.7748	6.8683	6.7804	6.7546	6.8021	6.7968	6.8067	6.7807	6.6514	6.7120	6.7112
		16	4.7890	4.7889	3.3940	3.3894	4.7937	4.7892	4.7822	3.3921	3.3628	4.7907	4.7889	4.7875	3.3922	3.3763	3.3955	3.3104
	0.4	4	4.1937	4.1940	3.9942	3.9439	4.1937	4.1940	4.1907	3.9934	3.9396	4.1936	4.1938	4.2037	3.9934	3.9379	3.9392	4.0418
Avorago		16	3.8054	3.8055	2.0004	1.9979	3.8074	3.8066	3.8208	2.0002	1.9971	3.8062	3.8066	3.8111	2.0002	1.9975	1.9983	2.1165
Average	0.8	4	6.4522	6.4524	6.4293	6.3534	6.4398	6.4190	6.4991	6.4250	6.4030	6.4463	6.4444	6.4185	6.4251	6.3008	6.3862	6.4235
		16	4.5892	4.5892	3.2073	3.2030	4.5934	4.5893	4.5911	3.2053	3.1769	4.5904	4.5889	4.5903	3.2054	3.1902	3.2040	3.3883

Table 11 - Forecasts RMSE - (N,T)=(50,10), SAR data generating process for φ, W(5,5) asymmetric weight matrix of French administrative communes, 1000 replications

	Estimators																	
			Pooled	Av. hetero.		DE	Pooled SAR	Av. hete	ro. SAR	FE-SAR	RE-SAR	Pooled SMA	Av. hete	ero. SMA	FE-SMA	RE-SMA	SAR-RE	SMA-RE
	λ	$\sigma_{_{\mu}}^{_{2}}$	OLS	OLS	FE	RE	MLE	ML	GM	MLE	MLE	MLE	ML	GM	MLE	MLE	GM	GM
1st vear	0.4	4	3.5992	3.5992	3.4169	3.3717	3.5997	3.5999	3.6059	3.4164	3.3749	3.5996	3.5996	3.6014	3.4164	3.3691	3.3497	3.4269
		16	3.5402	3.5402	1.6946	1.6936	3.5413	3.5414	3.5182	1.6944	1.6935	3.5413	3.5415	3.5471	1.6944	1.6943	1.7046	1.6667
i ot your	0.8	4	3.7404	3.7406	3.5619	3.5171	3.7423	3.7402	3.7387	3.5615	3.5079	3.7397	3.7395	3.7386	3.5614	3.5051	3.4983	3.7018
		16	3.5763	3.5763	1.7736	1.7724	3.5784	3.5768	3.5765	1.7733	1.7717	3.5761	3.5766	3.5518	1.7733	1.7719	1.7671	1.9863
	0.4	4	3.9882	3.9883	3.7807	3.7315	3.9886	3.9885	3.9929	3.7803	3.7346	3.9885	3.9885	3.9787	3.7804	3.7292	3.7141	3.7752
2nd year		16	3.7325	3.7326	1.8843	1.8820	3.7337	3.7335	3.7255	1.8841	1.8820	3.7333	3.7336	3.7482	1.8841	1.8824	1.8924	1.8990
Znu you	0.8	4	4.1487	4.1488	3.9466	3.8985	4.1508	4.1490	4.1595	3.9463	3.8911	4.1484	4.1482	4.1572	3.9463	3.8876	3.9060	4.0544
		16	3.7945	3.7945	1.9649	1.9636	3.7965	3.7950	3.7956	1.9645	1.9631	3.7946	3.7949	3.7681	1.9645	1.9632	1.9608	2.1374
ard year	0.4	4	4.1439	4.1440	3.9258	3.8766	4.1441	4.1442	4.1404	3.9253	3.8806	4.1441	4.1442	4.1386	3.9253	3.8742	3.8689	3.9639
		16	3.8086	3.8087	1.9598	1.9574	3.8097	3.8094	3.8019	1.9596	1.9574	3.8092	3.8094	3.8260	1.9596	1.9580	1.9655	1.9680
ord year	0.8	4	4.3099	4.3100	4.1039	4.0546	4.3127	4.3109	4.3117	4.1034	4.0469	4.3097	4.3096	4.3163	4.1034	4.0436	4.0568	4.0984
		16	3.8692	3.8693	2.0447	2.0431	3.8721	3.8702	3.8734	2.0443	2.0424	3.8699	3.8701	3.8470	2.0442	2.0425	2.0438	2.1830
	0.4	4	4.2235	4.2236	3.9988	3.9501	4.2239	4.2240	4.2289	3.9984	3.9547	4.2238	4.2240	4.2267	3.9984	3.9480	3.9478	4.0507
4th vear		16	3.8489	3.8490	2.0028	2.0005	3.8499	3.8498	3.8420	2.0026	2.0006	3.8495	3.8498	3.8650	2.0026	2.0011	2.0032	1.9883
Hirycar	0.8	4	4.3863	4.3865	4.1806	4.1298	4.3888	4.3869	4.3979	4.1799	4.1220	4.3861	4.3860	4.4070	4.1799	4.1188	4.1376	4.3898
		16	3.9086	3.9086	2.0920	2.0902	3.9114	3.9097	3.9116	2.0916	2.0894	3.9094	3.9096	3.8932	2.0916	2.0895	2.0859	2.1698
	0.4	4	4.2738	4.2739	4.0483	3.9994	4.2743	4.2744	4.2805	4.0478	4.0040	4.2742	4.2744	4.2790	4.0478	3.9974	3.9973	4.0307
5th year		16	3.8698	3.8699	2.0246	2.0222	3.8708	3.8706	3.8620	2.0244	2.0223	3.8704	3.8707	3.8878	2.0244	2.0228	2.0260	1.9919
ouryour	0.8	4	4.4411	4.4412	4.2345	4.1830	4.4432	4.4420	4.4491	4.2337	4.1754	4.4410	4.4410	4.4584	4.2337	4.1723	4.1896	4.3298
		16	3.9337	3.9338	2.1190	2.1171	3.9367	3.9350	3.9372	2.1186	2.1163	3.9347	3.9348	3.9176	2.1185	2.1163	2.1123	2.2562
	0.4	4	4.0457	4.0458	3.8341	3.7859	4.0461	4.0462	4.0497	3.8337	3.7898	4.0461	4.0461	4.0449	3.8337	3.7836	3.7756	3.8495
Average		16	3.7600	3.7601	1.9132	1.9112	3.7611	3.7609	3.7499	1.9130	1.9112	3.7607	3.7610	3.7748	1.9130	1.9117	1.9183	1.9028
	0.8	4	4.2053	4.2054	4.0055	3.9566	4.2076	4.2058	4.2114	4.0050	3.9486	4.2050	4.2049	4.2155	4.0050	3.9455	3.9577	4.1148
		16	3.8164	3.8165	1.9989	1.9973	3.8190	3.8174	3.8188	1.9984	1.9966	3.8169	3.8172	3.7955	1.9984	1.9967	1.9940	2.1465

Table 12 - Forecasts RMSE - (N,T)=(50,10), SMA data generating process for φ, W(5,5) asymmetric weight matrix of French administrative communes, 1000 replications

	Estimators																	
		Pooled Av. hetero. Pooled SA		Pooled SAR	Av. hete	ero. SAR	FE-SAR	RE-SAR	Pooled SMA	Av. hete	ero. SMA	FE-SMA	RE-SMA	SAR-RE	SMA-RE			
	ρ	σ_{μ}^{2}	OLS	OLS	FE	RE	MLE	MLE	GM	MLE	MLE	MLE	MLE	GM	MLE	MLE	GM	GM
1st year	0.4	4	3.9843	3.9843	3.8382	3.7875	3.9851	3.9851	3.9382	3.8371	3.7811	3.9858	3.9853	3.9646	3.8372	3.7793	3.7658	3.7748
		16	3.6053	3.6055	1.9078	1.9063	3.6062	3.6064	3.5985	1.9068	1.9047	3.6062	3.6064	3.5523	1.9068	1.9047	1.9053	1.8890
	0.8	4	7.1129	7.1136	7.2553	7.0794	7.1079	7.0996	7.0918	7.2510	7.0585	7.1095	7.1052	7.0710	7.2515	7.0284	7.0274	7.0651
		16	4.6280	4.6280	3.5936	3.5781	4.6288	4.6276	4.6144	3.5912	3.5555	4.6285	4.6285	4.6275	3.5912	3.5636	3.6072	3.5700
	0.4	4	4.3846	4.3848	4.2351	4.1769	4.3853	4.3854	4.3615	4.2332	4.1700	4.3862	4.3855	4.3940	4.2332	4.1694	4.1758	4.1855
2nd year		16	3.8386	3.8386	2.1195	2.1177	3.8400	3.8400	3.8379	2.1185	2.1161	3.8400	3.8401	3.7906	2.1186	2.1163	2.1200	2.1139
znu year	0.8	4	7.8694	7.8698	8.0282	7.8327	7.8657	7.8516	7.8364	8.0208	7.8165	7.8670	7.8612	7.8633	8.0211	7.7782	7.8009	7.7973
		16	5.0879	5.0877	4.0125	3.9963	5.0906	5.0888	5.0560	4.0091	3.9693	5.0907	5.0902	5.0778	4.0092	3.9789	3.9965	3.9857
3rd year	0.4	4	4.5505	4.5506	4.4045	4.3422	4.5510	4.5512	4.5436	4.4023	4.3345	4.5518	4.5513	4.5538	4.4023	4.3349	4.3386	4.3471
		16	3.9260	3.9261	2.2016	2.1996	3.9269	3.9270	3.9281	2.2007	2.1984	3.9270	3.9272	3.8853	2.2008	2.1985	2.2016	2.1962
	0.8	4	8.1983	8.1986	8.3547	8.1578	8.1953	8.1803	8.1728	8.3467	8.1401	8.1968	8.1916	8.1733	8.3468	8.1020	8.1314	8.0947
		16	5.2566	5.2564	4.1800	4.1625	5.2589	5.2569	5.2433	4.1764	4.1337	5.2589	5.2583	5.2468	4.1765	4.1442	4.1660	4.1409
	0.4	4	4.6440	4.6441	4.4982	4.4345	4.6446	4.6447	4.6292	4.4962	4.4266	4.6452	4.6449	4.6391	4.4963	4.4275	4.4165	4.4388
4th year		16	3.9691	3.9692	2.2455	2.2435	3.9699	3.9700	3.9702	2.2447	2.2422	3.9700	3.9703	3.9296	2.2447	2.2423	2.2405	2.2421
401 year	0.8	4	8.3769	8.3771	8.5328	8.3333	8.3737	8.3591	8.3368	8.5245	8.3186	8.3752	8.3709	8.3320	8.5247	8.2773	8.2932	8.2741
		16	5.3387	5.3386	4.2529	4.2362	5.3417	5.3396	5.3215	4.2493	4.2083	5.3417	5.3410	5.3354	4.2496	4.2180	4.2461	4.2309
	0.4	4	4.6925	4.6926	4.5465	4.4825	4.6933	4.6933	4.6871	4.5444	4.4746	4.6940	4.6935	4.6986	4.5444	4.4754	4.4726	4.4953
Ethe searce		16	3.9955	3.9957	2.2725	2.2704	3.9963	3.9964	3.9971	2.2717	2.2692	3.9964	3.9966	3.9525	2.2718	2.2693	2.2666	2.2691
Juliyear	0.8	4	8.4803	8.4805	8.6351	8.4361	8.4766	8.4619	8.4395	8.6269	8.4259	8.4780	8.4740	8.4222	8.6272	8.3795	8.4074	8.3813
		16	5.3877	5.3876	4.2969	4.2799	5.3903	5.3883	5.3718	4.2931	4.2520	5.3903	5.3897	5.3847	4.2933	4.2615	4.2954	4.2892
	0.4	4	4.4512	4.4513	4.3045	4.2447	4.4519	4.4519	4.4319	4.3026	4.2374	4.4526	4.4521	4.4500	4.3027	4.2373	4.2339	4.2483
Average		16	3.8669	3.8670	2.1494	2.1475	3.8679	3.8680	3.8664	2.1485	2.1461	3.8679	3.8681	3.8221	2.1485	2.1462	2.1468	2.1421
Average	0.8	4	8.0075	8.0079	8.1612	7.9679	8.0038	7.9905	7.9755	8.1540	7.9519	8.0053	8.0006	7.9723	8.1543	7.9131	7.9321	7.9225
		16	5.1398	5.1397	4.0672	4.0506	5.1421	5.1403	5.1214	4.0638	4.0237	5.1420	5.1415	5.1344	4.0640	4.0333	4.0622	4.0434

Table 13 - Forecasts RMSE - (N,T)=(50,10), SAR data generating process for ϕ , W(1,1), 1000 replications under non-normality of individual effects

	Estimators																	
			Pooled	Av. hetero.			Pooled SAR	Av. hete	ro. SAR	FE-SAR	RE-SAR	Pooled SMA	Av. hete	ero. SMA	FE-SMA	RE-SMA	SAR-RE	SMA-RE
	ρ	σ_{μ}^{2}	OLS	OLS	FE	RE	MLE	MLE	GM	MLE	MLE	MLE	MLE	GM	MLE	MLE	GM	GM
	0.4	4	3.6733	3.6734	3.4998	3.4549	3.6746	3.6745	3.6271	3.4987	3.4473	3.6748	3.6750	3.6574	3.4986	3.4485	3.4357	3.4539
1st voar		16	3.4929	3.4932	1.7556	1.7525	3.4940	3.4941	3.4804	1.7551	1.7510	3.4939	3.4940	3.4827	1.7550	1.7511	1.7336	1.7350
ist year	0.8	4	3.9969	3.9972	3.8617	3.8078	3.9973	3.9966	3.9817	3.8584	3.7885	3.9997	3.9992	3.9887	3.8579	3.7860	3.7983	3.8020
		16	3.5497	3.5495	1.9274	1.9242	3.5524	3.5514	3.5657	1.9249	1.9203	3.5522	3.5516	3.5819	1.9244	1.9192	1.9219	1.9204
	0.4	4	4.0535	4.0534	3.8677	3.8189	4.0541	4.0540	4.0285	3.8663	3.8117	4.0545	4.0544	4.0544	3.8662	3.8127	3.8111	3.8231
2nd year		16	3.7113	3.7116	1.9382	1.9358	3.7119	3.7121	3.7025	1.9375	1.9344	3.7118	3.7121	3.7042	1.9375	1.9345	1.9271	1.9294
2nu year	0.8	4	4.4167	4.4168	4.2835	4.2212	4.4170	4.4164	4.3894	4.2799	4.1969	4.4194	4.4188	4.4118	4.2792	4.1940	4.2167	4.2318
		16	3.7920	3.7921	2.1416	2.1385	3.7944	3.7939	3.8054	2.1394	2.1354	3.7942	3.7940	3.8197	2.1391	2.1343	2.1341	2.1358
3rd year	0.4	4	4.2007	4.2007	4.0168	3.9653	4.2009	4.2010	4.1888	4.0154	3.9581	4.2012	4.2013	4.2063	4.0153	3.9594	3.9635	3.9718
		16	3.7873	3.7876	2.0075	2.0051	3.7879	3.7883	3.7774	2.0067	2.0037	3.7879	3.7883	3.7772	2.0067	2.0038	2.0066	2.0078
	0.8	4	4.5866	4.5866	4.4500	4.3852	4.5876	4.5866	4.5760	4.4459	4.3620	4.5895	4.5891	4.5831	4.4454	4.3584	4.3836	4.3842
		16	3.8824	3.8826	2.2276	2.2245	3.8841	3.8836	3.8915	2.2251	2.2213	3.8837	3.8838	3.9154	2.2249	2.2202	2.2155	2.2156
	0.4	4	4.2871	4.2872	4.1026	4.0499	4.2877	4.2878	4.2745	4.1012	4.0429	4.2879	4.2881	4.2814	4.1010	4.0440	4.0378	4.0524
4th year		16	3.8243	3.8246	2.0460	2.0434	3.8246	3.8249	3.8150	2.0452	2.0420	3.8246	3.8250	3.8178	2.0452	2.0421	2.0451	2.0480
4th year	0.8	4	4.6830	4.6830	4.5438	4.4782	4.6831	4.6822	4.6643	4.5391	4.4548	4.6849	4.6846	4.6652	4.5384	4.4507	4.4714	4.4656
		16	3.9245	3.9247	2.2678	2.2647	3.9262	3.9258	3.9301	2.2654	2.2615	3.9258	3.9259	3.9626	2.2650	2.2602	2.2625	2.2616
	0.4	4	4.3336	4.3337	4.1502	4.0964	4.3340	4.3341	4.3277	4.1489	4.0897	4.3343	4.3345	4.3378	4.1488	4.0908	4.0843	4.0988
5th yoar		16	3.8487	3.8489	2.0716	2.0691	3.8490	3.8493	3.8374	2.0710	2.0677	3.8489	3.8493	3.8416	2.0710	2.0678	2.0739	2.0698
Sili yeai	0.8	4	4.7412	4.7413	4.6001	4.5343	4.7420	4.7408	4.7269	4.5957	4.5119	4.7436	4.7434	4.7220	4.5950	4.5073	4.5249	4.5125
		16	3.9535	3.9537	2.2953	2.2922	3.9548	3.9545	3.9573	2.2928	2.2887	3.9545	3.9546	3.9903	2.2923	2.2874	2.2895	2.2925
	0.4	4	4.1096	4.1097	3.9274	3.8771	4.1103	4.1103	4.0893	3.9261	3.8699	4.1106	4.1107	4.1075	3.9260	3.8711	3.8665	3.8800
Avoraça		16	3.7329	3.7332	1.9638	1.9612	3.7335	3.7337	3.7225	1.9631	1.9597	3.7334	3.7338	3.7247	1.9631	1.9598	1.9572	1.9580
Average	0.8	4	4.4849	4.4850	4.3478	4.2853	4.4854	4.4845	4.4677	4.3438	4.2628	4.4874	4.4870	4.4742	4.3432	4.2593	4.2790	4.2792
		16	3.8204	3.8205	2.1719	2.1688	3.8224	3.8218	3.8300	2.1695	2.1654	3.8221	3.8220	3.8540	2.1692	2.1643	2.1647	2.1652

Table 14 - Forecasts RMSE - (N,T)=(50,10), SMA data generating process for φ, W(1,1), 1000 replications under non-normality of individual effects