

IBS WORKING PAPER 04/2019
APRIL 2019

**TECHNOLOGY, SKILLS, AND
GLOBALISATION: EXPLAINING
INTERNATIONAL DIFFERENCES IN
ROUTINE AND NONROUTINE WORK
USING SURVEY DATA**

Piotr Lewandowski
Albert Park
Wojciech Hardy
Du Yang



TECHNOLOGY, SKILLS, AND GLOBALIZATION: EXPLAINING INTERNATIONAL DIFFERENCES IN ROUTINE AND NONROUTINE WORK USING SURVEY DATA[•]

Piotr Lewandowski[♦]

Albert Park[♥]

Wojciech Hardy[▲]

Du Yang[♣]

Abstract

The shift away from manual and routine cognitive work, and towards non-routine cognitive work is a key feature of labor markets. There is no evidence, however, if the relative importance of various tasks differs between workers performing seemingly similar jobs in different countries. We develop worker-level, survey-based measures of task content of jobs – non-routine cognitive analytical and personal, routine cognitive and manual – that are consistent with widely-used occupation-specific measures based on O*NET database. We apply them to representative surveys conducted in 42 countries at different stages of development. We find substantial cross-country differences in the content of work within occupations. Routine task intensity (RTI) of jobs decreases significantly with GDP per capita for high-skill occupations but not for middle- and low-skill occupations. We estimate the determinants of workers' RTI as a function of technology (computer use), globalization (specialization in global value chains), structural change, and supply of skills, and decompose their role in accounting for the variation in RTI across countries. Computer use, better education, and higher literacy skills are related to lower RTI. Globalization (as measured by sector foreign value-added share) increases RTI in poorer countries but reduces RTI in richer countries. Differences in technology endowments and in skills' supply matter most for cross-country differences in RTI, with globalization also important. Technology contributes the most to the differences in RTI among workers in high-skilled occupations and non-off-shorable occupations; globalization contributes the most to differences among workers in low-skilled occupations and offshorable occupations.

Keywords: task content of jobs, deroutinisation, global division of labor, PIAAC, STEP, CULS.

JEL: J21, J23, J24

• We thank Saier Wu, Peng Jia, Cheng Jie and Zeyang Yu for their excellent research assistance, Emily Pawlowski of the American Institutes for Research for help with using the US PIAAC restricted data. We thank participants of ASSA2019 in Atlanta, EALE2018 in Lyon, Jobs and Development Conference 2018 in Bogotá, TASKS V in Bonn, seminars at UCSD, The World Bank, and Princeton University, and workshops in Beijing, Geneva, Luxembourg and Warsaw for their insightful comments. We acknowledge support from the Hong Kong Research Grants Council General Research Fund [16504217]. The usual disclaimers apply. All errors are our own.

♦ Institute for Structural Research, Warsaw, and IZA, Bonn. E-mail: piotr.lewandowski@ibs.org.pl. Corresponding author.

♥ HKUST Institute for Emerging Market Studies, Hong Kong. E-mail: albertpark@ust.hk.

▲ Institute for Structural Research (IBS), Warsaw, and Faculty of Economics, University of Warsaw. E-mail: wojciech.hardy@ibs.org.pl.

♣ Chinese Academy of Social Science (CASS), Beijing. E-mail: duyang@cass.org.cn.

1. Introduction

In recent years, there has been an explosion of interest in how the future of work will be shaped by new technologies. Much speculation focuses on whether and how quickly specific job tasks can be performed by robots or artificial intelligence (Arntz et al., 2017; Frey and Osborne, 2013; McKinsey, 2017). Economists have found analysis of changes in job tasks to be a fruitful way to understand how labor market outcomes are influenced by the changing nature of work (Autor, 2013; Acemoglu and Autor, 2011; Firpo et al., 2011). In the US and other advanced countries, the relative share of routine jobs – both cognitive and manual – has declined over time, presumably because such jobs are easily replaced by computers or automation, or can be outsourced to other countries (Autor et al., 2003; Goos et al., 2014; Jensen and Kletzer, 2010; 2008; Michaels et al., 2014; Spitz-Oener, 2006). Because these jobs tend to be middle-skill jobs, the decline of routine jobs has contributed to wage polarization by hollowing out the middle of the wage distribution. However, in developing countries and emerging markets, evidence on how the nature of work is changing is decidedly mixed (World Bank, 2019). For example, there is evidence that in China and some Central Eastern European countries, routine-intensive occupations actually increased in recent decades (Du and Park, 2018; Hardy et al., 2018).

In this paper, we study how four fundamental forces predict the task content of jobs across countries: technology, globalization, structural change, and supply of skills.¹ Large labor productivity differences across countries as well as significant differences in information and communication technologies (ICT) and robot adoption suggest that large technological gaps remain across countries (Hsieh and Klenow, 2010; World Bank, 2019). Globalization is expected to lead to the outsourcing of routine-intensive tasks from high-wage countries to low-wage countries (Grossman and Rossi-Hansberg, 2008; Hummels et al., 2018). Structural changes such as industrialization and the growth of services alter the demand for goods and services which alter the demand for different types of jobs (Bárány and Siegel, 2018). Finally, the labor force in poorer countries often is much less educated, which could influence the optimal assignment of routine and non-routine tasks (World Bank, 2019).

The analysis of task demand in the US and globally has been greatly facilitated by the codification of the task content of different occupations in the US by the Department of Labor, first through the Dictionary of Occupational Titles (DOT) dating back to 1939, and since 2003 through the Occupation Information Network (O*NET). These databases provide detailed and periodically updated descriptions of the specific tasks associated with each occupation in the US. Acemoglu and Autor (2011) used O*NET data to construct what have now become standard indices of different types of job tasks: non-routine cognitive analytical, non-routine cognitive interpersonal, routine cognitive, routine manual, and non-routine manual. Because other countries have not systematically collected similar information on occupational job tasks, analyses of task demand in other countries has frequently used the US O*NET task data, requiring the assumption that the task content of occupations in those countries is identical to the US (Arias et al., 2014, Goos et al., 2014, Lewandowski et al., 2017, Hardy et al., 2018). This is almost certain to be problematic for less developed countries, given significant heterogeneity in the four fundamental forces described above.

¹ We focus on factors that directly influence prices of outputs and factors, and firms' technology. We do not consider institutional factors, although we recognize their importance in shaping technology, globalization, structural change, and skills, as well as the organization of work (firm size, management structure, etc.) all of which may influence job tasks.

In this paper, we present new evidence on the global distribution of tasks, focusing on the determinants of cross-country differences in the nature of work. We use micro-data on job tasks collected from large-scale surveys of workers in 42 countries around the world spanning developed and developing countries. The survey data come from three sources: the OECD's Programme for the International Assessment of Adult Competencies (PIAAC), the World Bank's Skills toward Employment and Productivity (STEP) surveys, conducted in middle- and low-income countries, and the China Urban Labor Survey (CULS) which included a module based on STEP. We develop harmonized survey-based measurements of non-routine cognitive analytical, non-routine cognitive interpersonal, routine cognitive, and manual tasks which closely mirror widely used task measurements for occupations proposed by Acemoglu and Autor (2011) based on the US O*NET data. However, our measures are worker-specific, enabling us to capture within-occupation differences in job tasks both within and across countries. Even in the US, research has shown considerable variation in tasks among workers within the same occupation (Autor and Handel, 2013). Construction of worker-specific task measures that are consistent with O*NET and cover low-, middle-, and high-income countries is a first main contribution of this study.²

Our second contribution is to document new stylized facts about cross-country differences in the task content of jobs. On average, workers in the more developed economies perform more non-routine cognitive tasks, both analytical and interpersonal, and less routine tasks. The relationship between relative routine task intensity and country GDP per capita differs quite markedly for different occupation groups. In high-skilled occupations (e.g., managers, professionals, technicians), there is a sharp gradient with respect to GDP per capita, with work being more routine-intensive in poorer countries. However, for middle-skill occupations like clerical workers, and low-skill occupations like plant and machine operators and assemblers, the relative routine-intensity of tasks varies considerably across countries but is *not* systematically related to the country's level of development. Overall, cross-country differences in task content within the same occupations are sizable. Understanding the extent and nature of these differences is of both scientific and policy relevance, as they reflect differences in the nature of work that can inform future labor market challenges, such as the share of jobs that can be automated (Arntz et al., 2017, Frey and Osborne, 2017).

Our third and most important contribution is to quantify for the first time how four fundamental forces – technology, globalization, supply of skills, and structural change – are associated with cross-country differences in the task content of jobs. Previous research has documented associations between specific factors for subsets of countries, but this study is the first to examine the impact of all of these factors in a comprehensive framework, and for countries that span low-, middle-, and high-income countries, and using survey-based measures.³ All of the previous studies assume that tasks within occupations are identical across countries.

² Other studies of tasks that use international survey data define measures of "de-routinisation" (de la Rica and Gortazar, 2016), group occupations into non-, low-, medium- and high-routine intensive (Marcolin et al., 2016), or combine manual and cognitive tasks (Dicarlo et al., 2016).

³ For example, earlier research documents the importance of ICT technology for the demand for tasks in the OECD countries (Akerman et al. 2015, Autor et al., 2003, 2006; Deming, 2017; Spitz-Oener, 2006), but no studies document the relationship between ICT and tasks in a cross-country setting that includes low-income countries. Evidence also exists that offshoring contributes to the shift away from routine work in the OECD countries (Goos et al., 2014; Hummels et al., 2018; Oldenski, 2012) and that participation in global value chains leads to a higher share of routine-intensive occupations in some

Our assessment of the relative importance of the four fundamental factors in predicting cross-country task differences starts with a regression in which workers' relative routine task intensity (a summary measure that combines routine cognitive and non-routine cognitive tasks) is regressed on individual, sector, and country-level variables that capture the four factors, for the pooled, global sample of all workers. Technology is captured by country-sector computer use calculated from the survey data, globalization is measured by the foreign share of value added in the country-sector plus this share interacted with GDP per capita, structural change is captured by 18 sector indicator variables and GDP per capita, and skills are captured by individual education, demographics (age, sex), and a direct test of literacy proficiency which allows us to measure skill more accurately than studies which only use data on education attainment (which cannot capture education quality differences). Given our finding that GDP per capita predicts relative routineness of tasks differently for different occupation groups, we run separate task regressions for workers in high-skilled occupations (managers, professionals and technicians), middle-skilled occupations (clerks, sales and services workers) and low-skilled occupations (craft and related trades workers, plant and machine operators and assemblers, elementary occupations). To investigate further the importance of globalization, we look separately at workers in offshorable versus non-offshorable occupations.

Using the coefficients from these regressions, we decompose gaps in mean routine task intensity across countries into the contributions associated with the four fundamental forces. We decompose the cross-country variance in mean routine task intensity, and compare each country to the US, summarizing the importance of different factors in predicting gaps between the US and groups of countries sorted by GDP per capita.

Our regression-based decompositions show that technology, the supply of skills and globalization are all strongly associated with cross-country differences in routine task intensity (RTI). International differences in technology use are especially important in accounting for cross-country variation in RTI in high-skilled occupations that are typically rich in non-routine tasks, highlighting the complementarity between non-routine tasks and ICT (Autor et al., 2003). Globalization contributes the most to cross-country differences in RTI among workers in low-skilled occupations, and in offshorable occupations. This finding is in line with the view that offshoring enables countries to specialize, within industries, according to their abundant factors (Grossman and Rossi-Hansberg, 2008). We also find that the supply of skills contributes to the cross-country differences in tasks mainly by shaping the employment shares of high-, middle-, and low-skilled occupations. Moreover, in the low- and middle income countries, lower supply of skills accounts for a large share of the difference in RTI compared to the US. The supply of skills is often overlooked in the studies of tasks that are focused on the most developed countries. We provide evidence that it in the poorer countries it should be accounted for as it may help to understand not only why the shares of high-skilled occupations are low, but also why the tasks performed by workers in given occupations are more routine-intensive. Finally, we show differences in occupational structure across countries accounts for a relatively small share of cross-country differences in task content. This highlights the importance of using comparable survey data to accurately estimate the extent of cross-country differences in task content and the determinants of those differences.

developing countries (Reijnders and de Vries, 2018). Regarding skill supply, a positive relationship between the supply of tertiary educated workers and non-routine tasks has been documented by studies using O*NET data (Hardy et al., 2018, Montresor, 2018; Salvatori, 2018). Structural change has been identified as relevant for polarization and shifts in tasks over time, both theoretically (Bárány and Siegel, 2018) and empirically (Du et al., 2017; Hardy et al., 2018).

In the second section we outline our methodology of creating the task content measures using the US PIAAC and O*NET data, and applying them to 42 countries covered by the PIAAC, STEP and CULS surveys. In the third section we present the cross-country differences in task structures, and in the fourth section we examine the determinants of these differences. The fifth section concludes.

2. Data and Task Measurement

2.1 Data

Our aim is to create task content measures based on PIAAC and STEP surveys which are worker-specific but are as consistent as possible with well-established measures of job tasks. To accomplish this objective, we first use the US PIAAC dataset to create measures that maximize consistency with US O*NET occupation-specific task measures popularized by Acemoglu and Autor (2011).

We use data from three comparable surveys: OECD's Programme for the International Assessment of Adult Competencies (PIAAC) the World Bank's Skills Measurement Program (STEP), and the third wave of the China Urban Labor Survey (CULS) conducted by the Institute of Population and Labor Economics of the Chinese Academy of Social Science (CASS). Our sample covers 42 countries in total.

In two rounds of PIAAC surveys (in 2011-12 and 2014-15), data was collected in 32 countries that made their data publicly available.⁴ The countries covered by PIAAC are high- or middle-income countries (see Appendix A for the list of countries). The survey respondents were aged 16-65, with sample sizes ranging from about 4000 in Russia to 26000 in Canada.⁵ STEP surveys are available for 12 low- or middle-income countries, out of which we use nine (Appendix A).⁶ The surveys were conducted between 2012 and 2014 of urban residents aged 15-64, with sample sizes ranging from about 2400 (in Ukraine) to approx. 4 000 (in Kenya), of urban residents aged 15-64.⁷ We also use the third wave of CULS which included the "skill use at work" questionnaire of STEP and therefore it is directly comparable to STEP. The survey was conducted in 2016 in six large cities in China (Guangzhou, Shanghai, and Fuzhou on the coast, Shenyang in the northeast, Xian in the northwest, and Wuhan in central China) and has a sample of 15 448 individuals.⁸ We refer to CULS as one of the STEP countries.⁹

⁴ In the US, PIAAC was supplemented by an additional wave aimed at enhancing the sample size, while retaining representativeness. We use this sample which is available from the US National Center for Education Statistics (NCES).

⁵ Individuals aged 15-year were also surveyed in Australia and Chile. Individuals aged 66-74 were surveyed in Australia.

⁶ We decided against using three available STEP datasets: Yunnan (China), Sri Lanka, and Vietnam. For China, we use the CULS data instead the STEP survey for the Chinese Yunnan province, as the former contains far more observations (almost 15 500) and covers a more comprehensive area. Yunnan is one of the poorer and more rural provinces in China so it might not reflect the dominant patterns of work in Chinese urban areas. Dicarlo et al. (2016) also omitted the Yunnan dataset. The survey of Sri Lanka includes too few observations in urban areas (about 650 workers), the Vietnam survey has low quality of data on skill use at work.

⁷ Because nearly all STEP surveys were urban only, for Laos which surveyed both urban and rural residents we drop the rural part of sample in order to ensure consistency.

⁸ The survey sampled 260 neighborhoods, 2 581 migrant households and 3 897 local households.

2.2 Task measures' definitions based on the US data

To construct survey-based task measures consistent with those based on O*NET, we first identify harmonized survey questions available in both PIAAC and STEP surveys whose content is similar to the questions used to construct the O*NET-based task measures (Autor, 2013). Then we systematically search for combinations of appropriate survey questions (and best groupings of answers) for which US PIAAC survey-based measures (averaged for each occupation) are most correlated with US O*NET-based occupation measures. Because PIAAC and STEP include only one question on physical tasks, we apply our procedures to the cognitive tasks measures only. For methodological details, see Appendix B.

Our procedure results in the following survey-based task definitions. The non-routine cognitive analytical task measure is based on questions on solving problems, reading news, reading professional journals, solving problems and programming. The non-routine cognitive interpersonal task measure is based on supervising others and making presentations. The routine cognitive task measure is based on the ability to change the order of tasks (reversed, so not being able to change the order of tasks), filling out forms, and making speeches or giving presentations (reversed, so making no speeches and giving no presentations). The manual task measure is based on the item describing if a job usually involves working physically for a long period. The cutoffs for each item are presented in Table 1.

In the US, our survey-based measures follow closely the task measures based on O*NET. At the 3-digit occupation level, the correlations between the survey measures (occupation-level averages) and the Acemoglu and Autor (2011) measures range from 55% (routine cognitive) to 77% (non-routine cognitive analytical, Table 1).¹⁰ The occupation-level averages of survey measures vary less between occupations than measures based on O*NET (Figure B1 in Appendix B). At the 3-digit occupation level, the standard deviations of tasks range from 0.50 (routine cognitive) to 0.67 (non-routine cognitive analytical) while the standard deviations of the O*NET-based tasks range from 1.02 (non-routine cognitive personal) to 1.23 (routine cognitive).¹¹ This shouldn't come as surprise as the O*NET measures are defined for occupations while the survey measures allow heterogeneity within occupations: at the 3-digit ISCO level, the within-group variance contributes from 65% (non-routine cognitive analytical) to 70% (non-routine cognitive personal, routine cognitive) to 83% (manual) of overall variance of the survey measures in the US.

⁹ We reweight the STEP and Indonesian⁹ data in order to achieve representativeness of the occupational structures in urban areas. To this aim, we retain the original shares of workers in agriculture and elementary occupations and adjust the distribution of other 1-digit ISCO occupations in line with occupational distributions reported in the International Labour Organization Database (ILOSTAT). In the case of China, we use the urban occupational distribution from the 2015 Census to reweight the CULS data.

¹⁰ The highest correlations obtained at the 4-digit occupation level range from 62% to 79%.

¹¹ The high standard deviation of routine cognitive tasks based on O*NET is driven by negative outliers: occupations 521 (Street and Market Salespersons), 951 (Street and Related Services Workers) and 952 (Street Vendors, excluding food). If these outliers are ignored, the standard deviation of routine cognitive tasks turns out the lowest among the O*NET based measures (0.97), similarly to our measures.

Table 1. The task items selected to calculate task content measures with the US PIAAC data

Task content	Non-routine cognitive analytical	Non-routine cognitive interpersonal	Routine cognitive	Manual
Task items	Solving problems Reading news (at least once a month – answers 3,4,5) Reading professional journals (at least once a month – answers 3,4,5) Programming (any frequency – answers 2,3,4,5)	Supervising others Making speeches or giving presentations (any frequency - answers 2,3,4,5)	Changing order of tasks - reversed (not able) Filling out forms (at least once a month – answers 3,4,5) Making speeches or giving presentations - reversed (never)	Physical tasks
Correlation with Acemoglu and Autor (2011) measures	0.77	0.72	0.55	0.74

*Note: The cutoffs for the “yes” dummy in brackets. The full wording of questions and definitions of cutoff are presented in Appendix C. Source: own elaboration based on US PIAAC and O*NET data.*

The survey measures also exhibit pair-wise correlations that are consistent with those exhibited by the Acemoglu and Autor (2011) measures (Table 2). The non-routine cognitive measures are strongly and positively correlated with each other, and negatively correlated with the routine cognitive and manual measures.¹² The moderate positive correlation between the routine cognitive and manual measures is also very close to those calculated using the O*NET-based measures. Overall, in the US our survey based measures proxy well the occupational patterns exhibited by the Acemoglu and Autor (2011) measures.

Table 2. Pair-wise correlations between particular task content measures across 3-digit ISCO occupations in the US

	Non-routine cognitive analytical	Non-routine cognitive personal	Routine cognitive	Manual
Acemoglu and Autor (2011) measures based on O*NET				
Non-routine cognitive analytical	1			
Non-routine cognitive personal	0.71	1		
Routine cognitive	-0.35	-0.54	1	
Manual	-0.64	-0.55	0.32	1
Survey measures based on PIAAC				
Non-routine cognitive analytical	1			
Non-routine cognitive personal	0.64	1		
Routine cognitive	-0.49	-0.57	1	
Manual	-0.57	-0.58	0.42	1

*Note: correlations between occupation-level averages in the case of survey measures. Weighted by employment level at the 3-digit ISCO level. Source: own calculations based on PIAAC and O*NET data.*

¹² This should alleviate the concerns related to the use of “Making speeches or giving presentations” variable in both the non-routine cognitive personal measure and the routine cognitive measure. The negative correlation (across occupations) between them is virtually identical to the one implicitly implied by the Acemoglu and Autor (2011) measures.

2.3 Country-specific task measurement

We use the definitions presented in Table 1 to calculate worker job task content measures for all countries studied.¹³ We also merge O*NET with PIAAC, STEP and CULS in order to calculate the Acemoglu and Autor (2011) task measures for each country. For both measures, we standardize the measure using the relevant mean and standard deviation in the US. Hence, for each task measure, zero reflects the US average and 1 reflects standard deviations in the US. As the STEP surveys are urban surveys, we omit skilled agricultural workers (ISCO 6) in all countries to improve comparability.

We create a synthetic measure of relative routine task intensity (RTI) at a worker-level, using the formula:

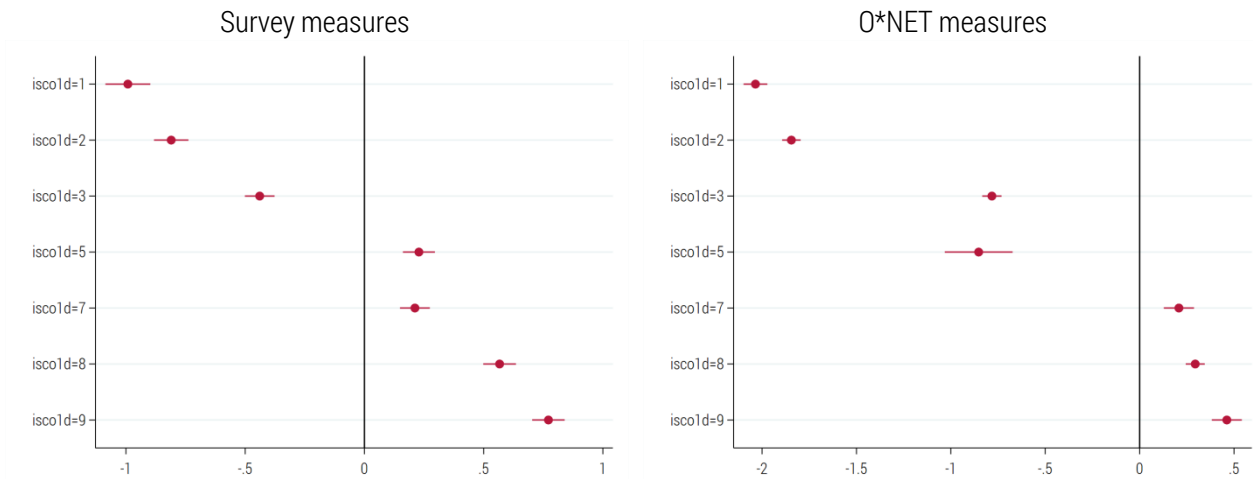
$$RTI = \ln(r_{cog}) - \ln\left(\frac{nr_{analytical} + nr_{personal}}{2}\right), \quad (1)$$

whereby r_{cog} , $nr_{analytical}$ and $nr_{personal}$ are routine cognitive, non-routine cognitive analytical and non-routine cognitive personal task levels, respectively. For each task, we add the lowest score in the sample to the scores of all individuals, plus 0.1, to avoid non-positive values in the logarithm. This definition follows the literature (Autor and Dorn, 2009, 2013; Goos et al., 2014) but we omit the manual tasks for two reasons: first, we cannot distinguish between routine and non-routine manual tasks. Second, the manual measure is less comparable across countries than the other measures – we provide evidence in the next section. For consistency, we standardize the RTI using its mean and standard deviation in the US.

The occupational patterns of the survey task content measures are consistent with those using O*NET measures, although they differ in important aspects. They are consistent as both measures show that workers in high-skill occupations (ISCO 1-3) perform, on average, less routine-intensive tasks, while workers in middle- and low-skill occupations (ISCO 4-5, ISCO 7-9) perform more routine-intensive tasks (Figure 1). Note that although we don't use the manual task variable in the survey-based RTI, it successfully captures the general routine aspect of work, not only the cognitive aspect. In particular, the routine task intensity among plant and machine operators and assemblers (ISCO 8), who perform highly routine jobs according to the RTI based on O*NET that accounts for manual routine tasks, is also high according to the survey measure (Figure 1). On the other hand, the RTI differences between occupations are lower according to the survey measures than according to the O*NET measures. The survey measure also shows that tasks performed by sales and services workers (ISCO 5) around the world are on average slightly more routine than tasks than the tasks performed by clerical support workers (ISCO 4, a reference group), contrary to O*NET measures (Figure 1).

¹³ The Ukrainian STEP does not include the question about reading of professional items so we are unable to calculate the non-routine cognitive analytical measure for Ukraine. We use only the non-routine cognitive personal measure in the RTI.

Figure 1. The differences in RTI across 1-digit ISCO occupations according to survey- and O*NET measures.

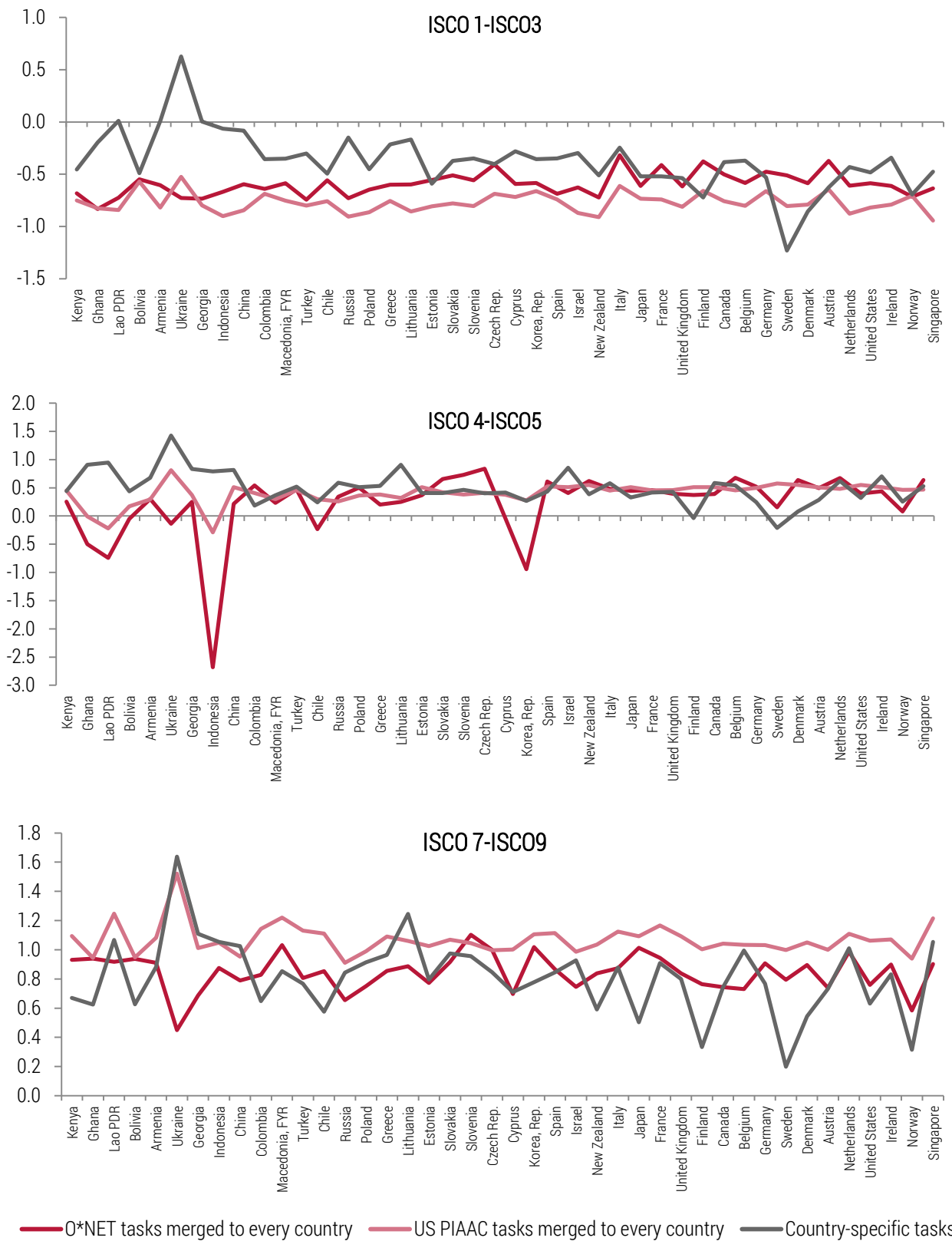


*Note: coefficients pertaining to occupation fixed effects (1-digit ISCO) estimated in a worker-level model on RTI against occupation fixed effects and country fixed effects. Manual tasks are included in the RTI based on O*NET. Sample size 151,625. Reference groups: Clerical support workers (ISCO 4), the United States. Source: own calculations based on PIAAC, STEP, CULS and O*NET data.*

The key difference between our survey measures and the measures based on O*NET is that the cross-country differences in task content within broad occupation groups are noticeably higher using the survey-based measures. The cross-country standard deviation of RTI using the survey-based measure is generally higher than that using the O*NET-based measure: 0.95 vs. 0.79 among workers in high-skill occupations, 1.00 vs. 1.10 among workers in middle-skill occupations, and 0.95 vs. 0.73 among workers in low-skill occupations. Using the survey-based measures, differences may result from country differences in work even within the most narrowly defined occupations. Using O*NET-based measures, cross-country differences are driven entirely by cross-country differences in occupational structures at finer ISCO levels. To highlight this point, we use the US PIAAC survey data to construct mean task content measures at the same occupational level that we use for O*NET measures. We then apply these to all workers assuming that job tasks in a specific occupation are identical across the world (Figure 2). The resulting task measures based on the assumption that occupations are identical across the world (but using our survey-based measure) follow the measures based on O*NET much more closely than they follow country-specific measures based on the survey data. Hence, it's the country-specific measurement that matters.

The cross-country differences in O*NET-based measures may also result from inconsistencies in coding of occupations. Indeed, the O*NET measures suggest that in many low- or middle-income countries (e.g. Bolivia, Ghana, Indonesia) workers in middle-skilled occupations perform highly non-routine work that on average is less routine-intensive than the work in high-skilled occupations (Figure 2). This implausible conclusion results from the fact that large shares of workers in these countries are classified as street sellers or services workers. These occupations require a lot of interpersonal tasks in the US but may not require as many of them in poorer countries. Indeed, the survey measures show that work in the middle-skilled occupations in poorer countries is quite intensive in routine tasks.

Figure 2. The average values of relative routine intensity (RTI) according to different methodologies, by occupational categories.



Note: Countries are ranked according to the GDP per capita level. For brevity, we aggregate occupations into three categories: high-skilled (ISCO 1-3), middle-skilled (ISCO 4-5) and low-skilled (ISCO 7-9). Results for particular occupations are available upon request.

Source: own calculations based on PIAAC, STEP, CULS, O*NET (tasks), and World Bank data (GDP).

3. Cross-country differences in the task content of jobs

We find substantial cross-country differences in the average values of particular task content measures. In general, the more developed countries exhibit higher average values of non-routine tasks than the less developed countries (Figure 3). The Nordic countries (Denmark, Sweden, Norway, Finland), most of the English-speaking countries (Canada, New Zealand, the UK and the US) and Singapore stand out with the highest levels of non-routine cognitive tasks. Perhaps not surprisingly, the less developed countries – Georgia, Ghana, Laos, Colombia, Turkey, Indonesia, but also Lithuania and Greece – have the lowest average values of non-routine cognitive tasks. The average value of non-routine cognitive tasks, especially of analytical tasks, is also low in China. The differences between the average values of non-routine tasks in the highest-scoring and the lowest-scoring countries are of a magnitude comparable to a one standard deviation of particular task content values among the US workers.

The relationship between routine cognitive tasks and the level of development is inverse U-shaped (Figure 3). The least developed countries and the Nordic countries exhibit the lowest values of routine cognitive tasks. On the other hand, Central and Eastern European countries (Ukraine, Lithuania, Czechia, Russia, Slovakia, Slovenia) have the highest average values of routine cognitive tasks. The values of routine cognitive tasks are also high in Southern European countries (Greece, Italy), as well as in the United Kingdom and Ireland.

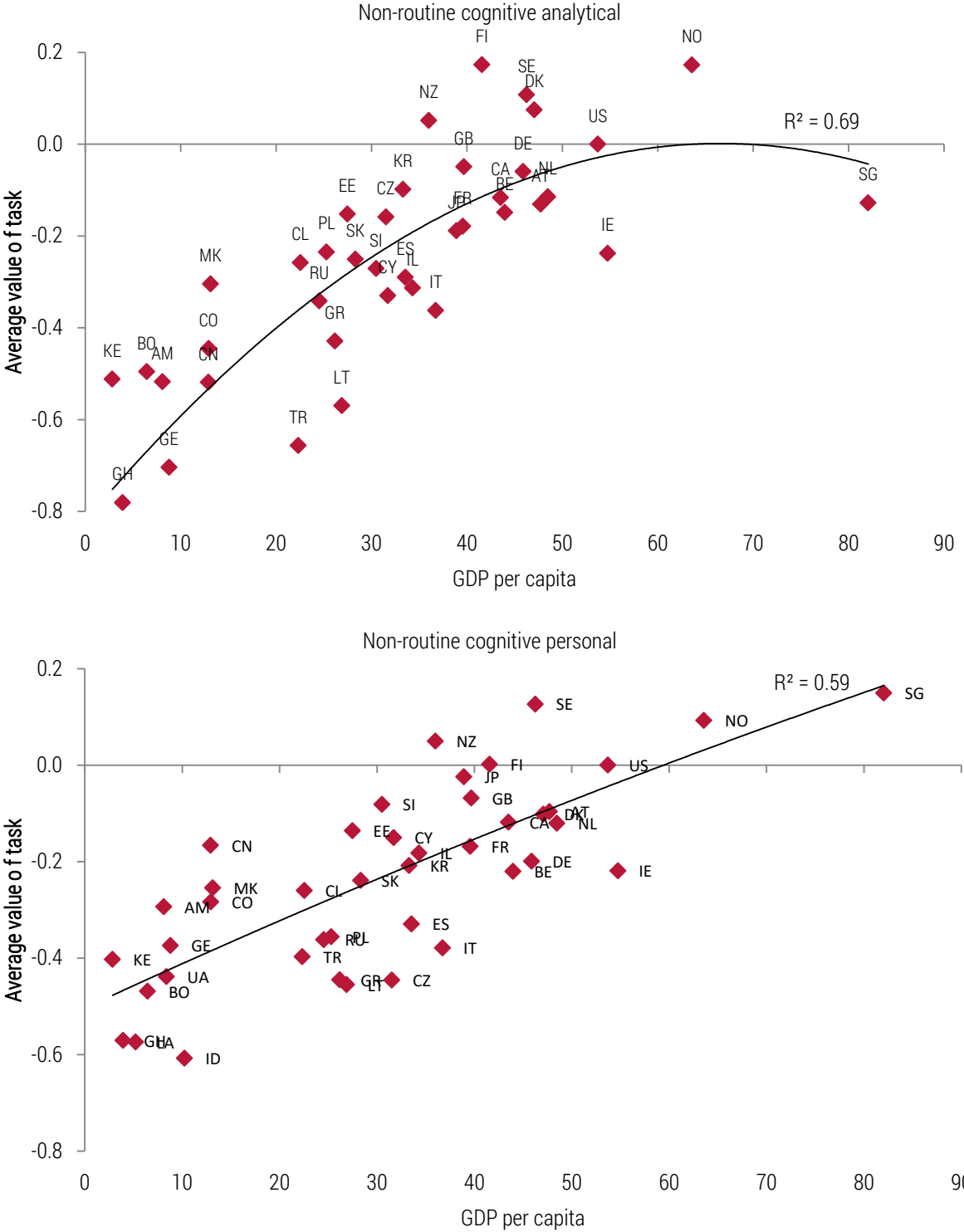
The average values of manual tasks do not show any clear-cut relationship with the level of development. For instance, Indonesia and Turkey exhibit the highest manual task levels, but Armenia, Macedonia, Ukraine and Georgia are among the countries with the lowest levels. United States and New Zealand are among the countries with highest manual levels, while Japan, Finland and Belgium among those with the lowest. However, these differences should be interpreted with caution as we are able to use only one task item for the manual task content measure. In all further analyses, we will focus on the other three measures and the RTI.

Importantly, our survey-based measures show large cross-country differences in the relative routine-intensity of tasks in particular occupations (Figure 4). Among workers in high-skill occupations (ISCO1 – managers, ISCO 2 – professionals, ISCO 3 – technicians) individuals in the more developed countries consistently perform less routine-intensive tasks than those in poorer countries. For those in middle- or low-skill occupations, the relationship between GDP per capita and relative routine intensity is inconsistent. Among sales and services workers (ISCO 5) and to a lesser extent craft and related trades workers (ISCO 7), those in richer countries do less routine work. However, among clerical workers (ISCO 4) and workers in the low-skilled occupations (ISCO8 – plant and machine operators, ISCO9 – elementary occupations), the cross-country differences are highly variable but are not correlated with the level of GDP per capita.¹⁴

Overall, our results show that the higher is the GDP per capita of a country, the higher is the relative role of non-routine content of jobs, in particular among the high-skilled occupations.

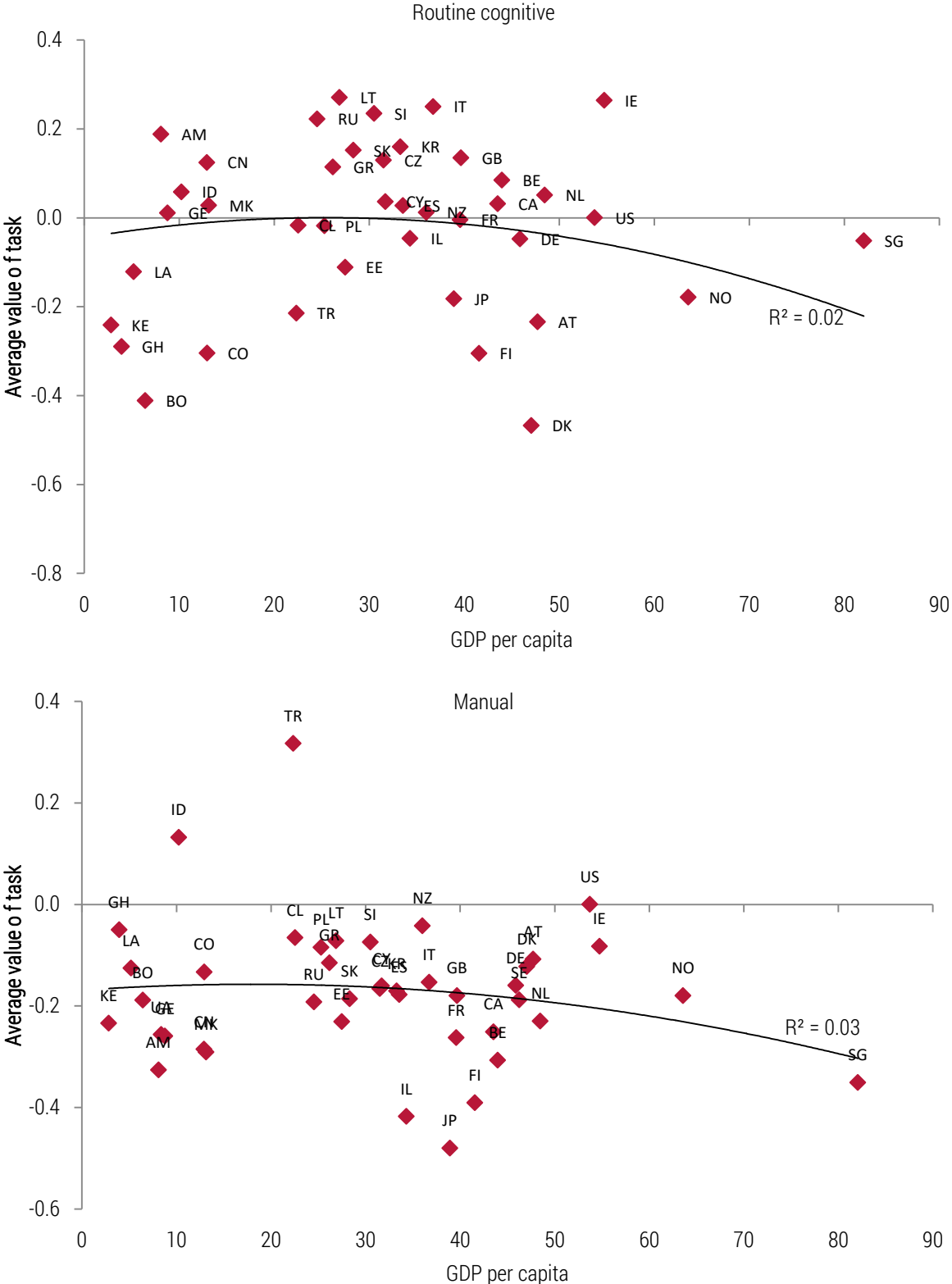
¹⁴ The standard deviation of country-specific average routine task intensity ranges from 0.24 (elementary occupations) to 0.32 (professionals).

Figure 3. The average values of tasks against GDP per capita.



Note: for each task content, the 0 is set at the US average value and 1 corresponds to one standard deviation of this particular task content value in the US. GDP per capita in PPP, current international \$, country averages for 2011-2016. Source: own calculations based on PIAAC, STEP, CULS (tasks), and World Bank data (GDP).

Figure 3 (cont'd). The average values of tasks against GDP per capita.



Note: for each task content, the 0 is set at the US average value and 1 corresponds to one standard deviation of this particular task content value in the US. GDP per capita in PPP, current international \$, country averages for 2011-2016. Source: own calculations based on PIAAC, STEP, CULS (tasks), and World Bank data (GDP).

Figure 4. Average values of routine intensity of tasks (RTI) by 1-digit occupations against GDP per capita.

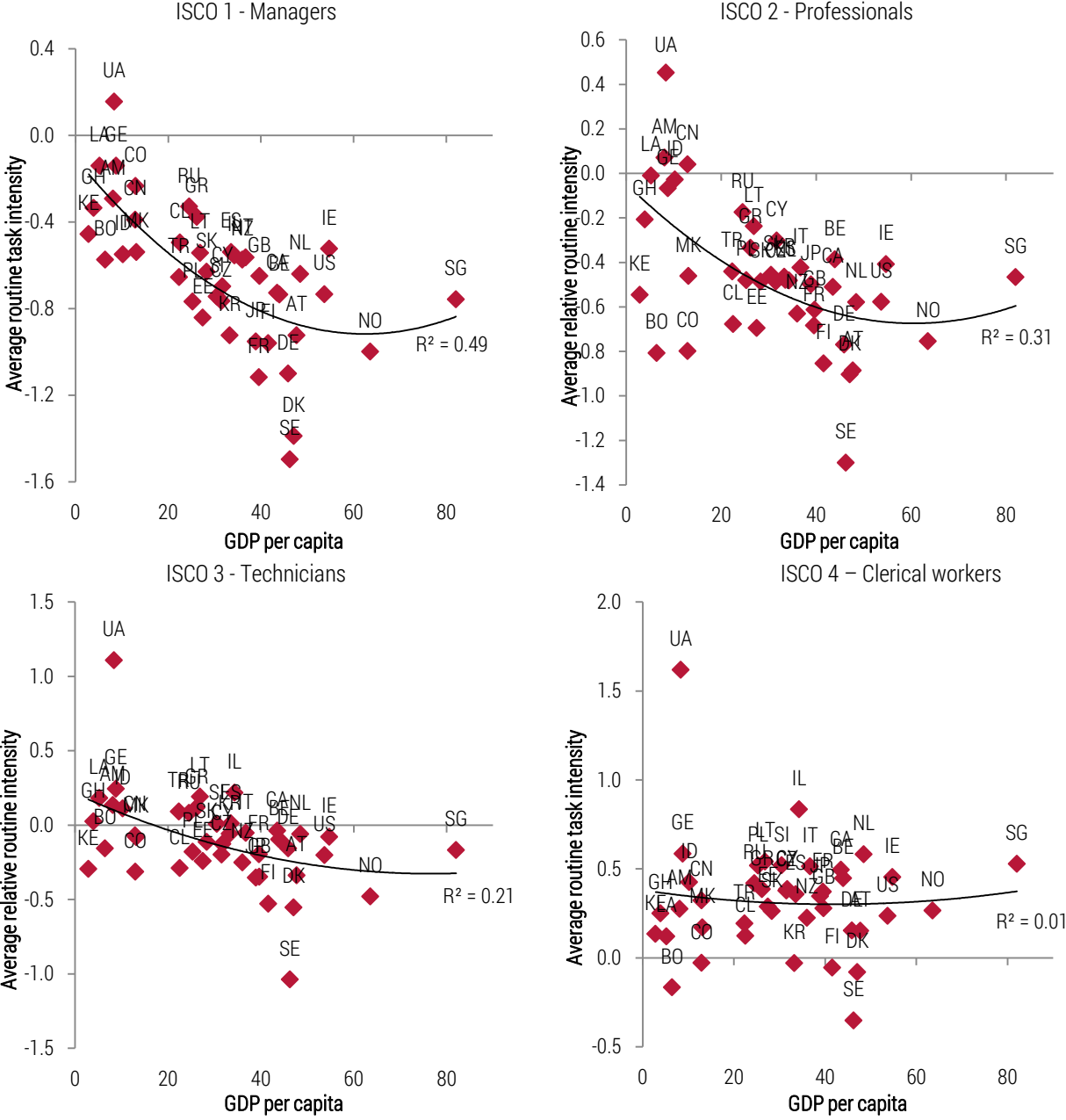
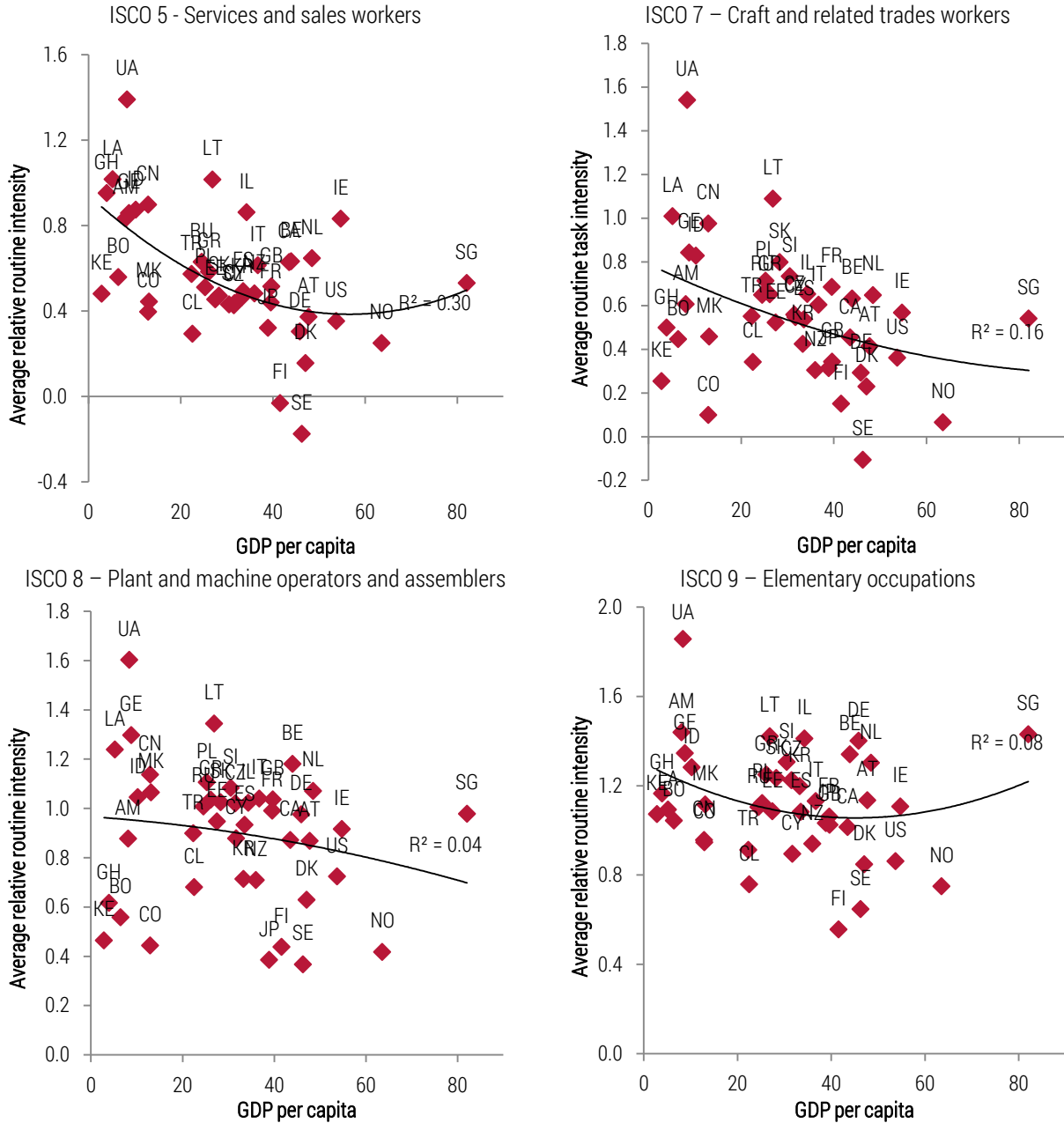


Figure 4 (cont'd). Average values of routine intensity of tasks (RTI) by 1-digit occupations against GDP per capita.



Note: the horizontal axis denotes GDP per capita, PPP (international \$, country averages for 2011-2016). We omit the occupational group ISCO 6 (Skilled agricultural workers) because of small sample sizes, especially in countries where surveys covered only urban areas.

Source: own calculations based on PIAAC, STEP, CULS, O*NET and World Bank data.

4. Determinants of Task Differences Across Countries

4.1 Methodology

To shed light on factors associated with cross-country differences in routine task intensity, we estimate pooled OLS regressions of the form:

$$RTI_{ijsc} = \beta_0 + \beta_1 Z_{sc} + \beta_2 G_{sc} + \beta_3 E_{ijsc} + \lambda_s + \varepsilon_{ijsc}, \quad (2)$$

Here, RTI_{ijsc} is the routine task intensity of individual i in occupation j in sector s in country c , Z_{sc} is technology used in sector s in country c , G_{sc} measures globalization in sector s in country c , E_{ijsc} are the individual skills of workers, and λ_s are sector fixed effects. Because the regressions are cross-sectional, they are best thought of as characterizing equilibrium allocations of tasks rather than being interpreted causally. The technology, globalization, and structural change measures are all country-sector level measures, which are plausibly exogenous to the decisions of individual firms and workers. We measure skills at the individual level, given that education and literacy are mostly pre-determined before entering the labor market. We have also conducted the analysis defining skill levels at the sector-level and discuss how this alters the results below.

Turning to measurements, the main technology variable is the share of workers in sector s in country c who use computers at work. PIAAC and STEP surveys include a question on individual computer use, and we aggregate this variable to the sector level due to concerns that decisions about computer use and tasks are made simultaneously. Separately, we also test the impact of robot stock per worker by sector (International Federation of Robotics), and country-level ICT capital stock per worker (Eden and Gaggl, 2015). Adding these variables turns out not to alter the main findings in an important way, but these data are available for only 31 countries so we exclude them from our preferred specification.

We employ two variables to measure globalization – participation in global value chains (henceforth GVC participation, Wang et al., 2017), and FDI stock as a share of GDP.¹⁵ Our basic GVC participation variable is the backward linkage-based measure defined as the foreign value added share in production of final goods and services (FVA share). For robustness we also use the forward linkage-based measure (domestic value added from production of intermediate exports or domestic factor content in intermediate exports, Wang et al., 2017). We allow for different effects of GVC participation in developed and developing countries by interacting the GVC participation with GDP per capita (log, demeaned). This captures the prediction that globalization reduces routine tasks in rich countries and increases them in poor countries.

To measure worker skills, we include a test-based measure of literacy skills (four proficiency levels), education level (primary, secondary, tertiary), age (measured by 10-year age groups), and gender. The literacy test is comprehensive and quantifies individuals' skills to understand, evaluate, use and engage with written texts in personal, work-related, societal and educational contexts (PIAAC Literacy Expert Group, 2009).¹⁶

¹⁵ Data sources and precise definitions of the technology and globalization variables are provided in Appendix E.

¹⁶ We account for the fact that PIAAC and STEP include multiple plausible values of the literacy proficiency variables. To this aim, we use the “pv” package in Stata that implements the Rubin (1987) combination methods which are standard in skill assessment literature. China and Laos did not collect literacy data, so we impute the literacy scores for those two countries

To capture the impact of structural change, we add indicator variables for 18 of 19 sectors based on the one-digit codes of International Standard Industrial Classification (ISIC Rev.4), as well as their interactions with GDP per capita (log, demeaned).¹⁷

We estimate the regressions for all workers, and for subgroups of workers.¹⁸ Given the evidence presented above that the cross-country patterns in routine-intensity vary for occupations of different skill levels, we estimate separate regressions for workers in high- (ISCO 1-3), middle- (ISCO 4-5) and low-skilled (ISCO 7-9) occupations. Second, to examine how globalization alters the determination of job tasks, we distinguish between workers in offshorable and non-offshorable occupations.¹⁹ To do so, we use the Blinder and Krueger (2013) classification of occupations into offshorable and non-offshorable jobs (see Appendix F for details).

In order to assess the relative importance of the four fundamental factors in predicting cross-country differences in tasks, we use the estimated coefficients to calculate a linear prediction of routine task intensity at the country level \widehat{RTI}_c . We decompose the variance of RTI using the covariance-based decomposition proposed by Morduch and Sicular (2002). Formally, the contribution of a variable group, k , to the variance of RTI is defined as follows:

$$\sigma_k = \frac{cov(\beta_k \bar{X}_c^k, \widehat{RTI}_c)}{var(\widehat{RTI}_c)}, \quad (3)$$

We use the average worker characteristics in each country (denoted by dashed symbols) to decompose the difference in the linear prediction of RTI in country c , \widehat{RTI}_c , and the US, \widehat{RTI}_{US} , to the contributions of various factors:

$$\widehat{RTI}_c - \widehat{RTI}_{US} = \beta_1(\overline{Z}_{Ijsc} - \overline{Z}_{IjsUS}) + \beta_2(\overline{G}_{sc} - \overline{G}_{sU}) + \lambda(\overline{S}_{sc} - \overline{S}_{sUS}) + \beta_3(\overline{E}_{Ijsc} - \overline{E}_{IjsUS}), \quad (4)$$

For presentation purposes, we aggregate countries to three groups based on their development level, and use unweighted averages of differences in RTI, all explanatory variables and all contributions (Table 3).

Table 3. Allocation of countries to groups based on GDP per capita

	Low and Middle Income Countries	Bottom High Income Countries	Top High Income Countries	Reference country
Countries	Kenya, Ghana, Lao, PDR, Ukraine, Bolivia, Indonesia, China, Armenia, Georgia, Colombia, Russia, Turkey	Chile, Poland, Lithuania, Slovakia, Cyprus, Estonia, Greece, Czech Rep., Slovenia, Spain, Korea, Rep., Italy	France, Israel, Japan, New Zealand, United Kingdom, Belgium, Germany, Canada, Finland, Austria, Netherlands, Ireland, Sweden, Denmark, Norway, Singapore	United States

Source: own elaboration based on World Bank data

using a regression estimated for other countries, controlling for education, demographic characteristics, occupation and sector of employment, computer use at work as well as macroeconomic variables (GDP per capita, export, FDI).

¹⁷ In order to achieve consistent sector definition across all countries, we merge sectors D (Electricity, gas, steam and air conditioning supply) with E (Water supply; sewerage, waste management and remediation activities), and M (Professional, scientific and technical activities) with N (Administrative and support service activities).

¹⁸ We drop Macedonia from our sample due to lack of data on globalization variables, and estimate our models on a sample of 41 countries.

¹⁹ We thank Gordon Hanson for this suggestion.

4.2 Determinants of task differences among all workers and by occupational groups

Results of a benchmark regression estimated for all workers, as well as for workers in high (ISCO 1-3), middle (ISCO 4-5) and low-skilled (ISCO 7-9) occupations are presented in Table 4.

Better access to technology is associated with lower routine intensity of tasks performed by workers. The higher is the probability of computer use in a sector, the less routine-intensive are tasks performed by workers in this sector. A 25 pp. higher share of computer use in the sector, which is equivalent to the difference between the US (75%) and China (50%) would translate into routine task intensity being lower by 0.1 standard deviations of RTI in the US, which is equivalent to 15% of the difference between average RTI in the US and China. We also find evidence that the routine-replacing function of computers is relevant especially for jobs that require higher skills, as the relationship between probability of computer use and RTI is by far the strongest among workers in high-skilled occupations, and not significant among workers in low-skilled occupations. We also examine the impact of sector robot stock and ICT stock (both expressed in per worker terms) for a subsample of countries and find no significant relationship between either of these variables and the routine task intensity (Table G3). This shows that the probability of using computer is a key technology variable related to the routine content of work.

Globalization also plays an important role. A higher foreign value added (FVA) share in domestic production is associated with a higher routine task-intensity (evaluated at mean GDP per capita in our sample). Thus, workers in country-sectors that specialize in smaller segments of global value chains (e.g., assemblers of final products) tend to perform more routine tasks. This effect is particularly strong for low-skilled workers, but not significant for high- or middle-skilled workers. We also find that the coefficient on the interaction term between FVA share and $\ln(\text{GDP per capita})$ is large and negative. This suggests that for a country with GDP per capita twice the mean, the FVA share has no effect on RTI, and that the positive impact of FVA share on RTI is nearly twice as great in countries with half of mean GDP per capita. The second globalization measure, FDI share of GDP, is barely significant in the regression for all workers, but this masks heterogeneity across occupations of different skill levels. FDI is positively associated with routine-intensity of high-skill workers, but negatively associated with RTI of low-skill workers. However, the magnitude of effects pertaining to FVA share is much larger than that of FDI. For instance, a 25 pp. higher FVA share, which is a difference between the US and countries most specialized in smaller segments of global value chains (e.g. small Central Eastern European countries) is associated with a routine task intensity being higher by 0.1 US standard deviations, which is equivalent to about 40% of the RTI difference between the US and these small CEE countries. But the 30pp. difference in FDI share between the US and these would translate into the RTI in these countries being lower by only 0.003 of the US standard deviation. Overall, the results are consistent with theories arguing that routine jobs are easier to offshore and so poorer countries may specialize in them (Grossman and Rossi-Hansberg, 2008). Our findings are robust to the choice of the GVC participation measure – the results of the estimation with forward linkage-based measure of participation in global value chains used instead of the backward linkage-based measure are presented in Table G2 in Appendix G.

Next, we turn to the skill variables. Workers with higher education levels and higher literacy are more likely to perform less routine tasks, overall and within particular occupational groups. A worker with the highest literacy proficiency (level 4-5) is expected to perform 0.029 (of the US standard deviation) less routine-intensive task than an otherwise identical worker with a lower medium literacy proficiency (level 2). We also find that the workers performing more routine-intensive jobs are more likely to be female and young (aged 16-24). The relationship

between age and routine task intensity varies among occupational groups. In high-skilled occupations (ISCO 1-3), older individuals perform significantly less routine-intensive tasks, but in middle- (ISCO 4-5) and low-skilled occupations (ISCO 7-9), older workers perform more routine-intensive tasks, especially if aged over 55. This difference may suggest that experience and firm- or sector-specific knowledge can play a more important role for allocation of workers to tasks among high-skilled occupations than among middle- and low-skilled occupations.

The sector of employment also matters for the routine intensity of tasks performed by workers. In particular, workers in service sectors, such as information and communication; financial and insurance activities; education; or arts, entertainment and recreation perform less routine-intensive jobs than workers in the reference sector (wholesale and retail trade). On the other hand, workers in transportation and storage perform more routine-intensive tasks than workers in trade in repairs. However, there are notable heterogeneities of these sector effects across occupation skill groups. Low-skilled workers in many services sectors in which work is largely low-end support services (e.g. information and communication, financial and insurance activities, education, and human health and social work activities) perform significantly more routine-intensive jobs than low-skilled workers in trade and repairs. Similarly, we find that high-skilled workers in real estate activities, financial and insurance activities, or health exhibit higher routine-intensity of tasks than high-skilled workers in trade and repairs. Workers in manufacturing perform more routine intensive tasks, but this effect is driven by the dominant demand for middle- or low-skilled jobs in manufacturing and is not present within particular occupational groups.

As an additional robustness check, we re-estimate our benchmark specification using averages at the sector level (Table G4 in Appendix G). The results confirm the negative relationship between the probability of computer use and skills, and RTI. However, the coefficients pertaining to the employment shares of educational groups are not significant at the sector level. This suggests that the significance of education in the worker-level regressions reflects the allocation of less routine tasks to better educated workers within sectors. The globalization variables are barely significant in the sector-level regression, which suggests that the heterogeneous effects we find for workers in high- and low-skilled occupations cancel each other out at the sector level.

Table 4. The correlates of routine task intensity (RTI) at the worker level

	All workers	High-skilled occupations (ISCO 1-3)	Middle-skilled occupations (ISCO 4-5)	Low-skilled occupations (ISCO 7-9)
Computer use	-0.501** (0.203)	-0.690*** (0.171)	-0.353 (0.314)	-0.240 (0.234)
Foreign Value Added (FVA) share	0.266* (0.151)	-0.057 (0.140)	0.189 (0.191)	0.796*** (0.173)
Ln(GDP per capita) – mean(Ln(GDP per capita))	0.057 (0.079)	-0.038 (0.068)	0.013 (0.105)	0.052 (0.078)
FVA share * [Ln(GDP pc) – mean(Ln(GDP pc))]	-0.424** (0.197)	-0.216 (0.171)	-0.239 (0.214)	-0.347 (0.233)
FDI / GDP	0.009* (0.005)	0.023*** (0.007)	0.010 (0.008)	-0.016*** (0.006)
Education: primary	0.246*** (0.020)	0.135*** (0.030)	0.223*** (0.022)	0.135*** (0.027)
Education: tertiary	-0.486*** (0.017)	-0.267*** (0.018)	-0.198*** (0.021)	-0.142*** (0.034)

Literacy skills level: 1 or lower	0.077*** (0.018)	0.032 (0.028)	0.051** (0.026)	0.057** (0.025)
Literacy skills level: 3	-0.138*** (0.013)	-0.086*** (0.016)	-0.062*** (0.022)	-0.048** (0.023)
Literacy skills level: 4 and 5	-0.293*** (0.019)	-0.190*** (0.019)	-0.064** (0.029)	-0.174*** (0.044)
Female	0.249*** (0.012)	0.239*** (0.013)	0.203*** (0.018)	0.346*** (0.025)
Age: 16-24	0.227*** (0.017)	0.220*** (0.028)	0.207*** (0.025)	0.147*** (0.025)
Age: 35-44	-0.054*** (0.011)	-0.062*** (0.014)	-0.020 (0.018)	-0.038* (0.021)
Age: 45-54	-0.012 (0.014)	-0.062*** (0.016)	0.017 (0.023)	0.043* (0.022)
Age: 55-65	0.020 (0.016)	-0.052*** (0.019)	0.110*** (0.026)	0.078*** (0.025)
Agriculture [A]	0.034 (0.075)	-0.101 (0.097)	-0.143 (0.143)	0.042 (0.096)
Mining [B]	-0.048 (0.071)	-0.080 (0.073)	-0.206 (0.131)	-0.063 (0.104)
Manufacturing [C]	0.049 (0.060)	-0.026 (0.066)	-0.136** (0.068)	-0.054 (0.069)
Electricity & Water supply [D+E]	0.089 (0.064)	0.101 (0.075)	-0.121 (0.131)	0.169*** (0.065)
Construction [F]	-0.074 (0.058)	-0.207*** (0.062)	-0.133 (0.081)	-0.147** (0.069)
Transportation and storage [H]	0.225*** (0.061)	-0.035 (0.083)	0.088 (0.084)	0.129* (0.070)
Accommodation and food service [I]	0.027 (0.062)	-0.177** (0.078)	0.047 (0.083)	0.136* (0.079)
Information and communication [J]	-0.307*** (0.097)	-0.086 (0.089)	-0.124 (0.136)	-0.019 (0.118)
Financial and insurance [K]	-0.097 (0.102)	0.162* (0.084)	-0.130 (0.141)	0.337** (0.150)
Real estate & Professional [L]	-0.070 (0.077)	0.136 (0.112)	0.005 (0.099)	-0.023 (0.123)
Administrative [M+N]	-0.071 (0.065)	-0.017 (0.069)	-0.024 (0.075)	0.177** (0.070)
Public administration [O]	-0.075 (0.083)	0.081 (0.069)	-0.111 (0.112)	0.151 (0.097)
Education [P]	-0.277*** (0.078)	-0.074 (0.064)	-0.089 (0.118)	0.199** (0.088)
Human health [Q]	0.018 (0.070)	0.229*** (0.062)	0.037 (0.076)	0.329*** (0.091)

Arts [R]	-0.245*** (0.069)	-0.123* (0.068)	-0.032 (0.071)	-0.069 (0.112)
Other service [S]	-0.227*** (0.068)	-0.187*** (0.072)	-0.262*** (0.078)	-0.016 (0.074)
Activities of household [T]	0.130 (0.113)	-0.657** (0.328)	0.040 (0.161)	0.068 (0.151)
Extraterritorial organizations [U]	-0.036 (0.140)	0.027 (0.146)	-0.323 (0.209)	0.564 (0.348)
No. of observations	148,569	62,907	47,373	38,289
R-squared	0.220	0.126	0.090	0.083

Note: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$. Standard errors in parentheses. We use standardized weights that give each country equal weight. The reference levels are: age 25-34, secondary education, wholesale and retail trade; repair of motor vehicles and motorcycles (ISIC G), lower medium literacy skills (level 2). The coefficients for interactions between sector fixed effects and $\ln(\text{GDP per capita})$ are presented in Table G1 in Appendix G.

Source: own estimations based on PIAAC, STEP, CULS World Bank, and RIGVC UIBE (2016) data.

For all workers, the regression model whose estimation results are reported in the second column of Table 4 accounts for more than half of the cross-country variance in RTI (Table 5). About 23% of the variance can be attributed to differences in technology (computer use), followed in importance by globalization (20%) and skills (18%), while structural change has a small negative contribution (-6%). When we conduct the analysis separately for high-, medium-, and low-skill occupations (regression results in columns 3-5 of Table 4), we find interesting differences in the relative importance of different factors. For high skill workers, the explained variance is similar to the one for all workers, and the contribution of technology to cross-country variance is the greatest (26%), followed in importance by globalization (10%), structural change (10%), and skills (7%). The explained variance for middle- and low-skill workers (25% and 23%) is much less than for all workers (57%), and both structural change and skills account for very little variance. For those in middle-skill occupations, just as for all workers, technology accounts for the most variance (13%), compared to globalization's 8%. But for low-skill occupations, the most important factor is globalization, which accounts for 21% of cross-country variance, with technology accounting for only 6%.²⁰ The contribution of structural change for all workers and for workers in low-skilled occupations is negative because the employment shares of some typically non-routine sectors (e.g. education) are virtually the same in all country groups, and the shares of some typically routine sectors (e.g. manufacturing) are in some low- and middle-income countries are lower than in the high-income countries.²¹ The fact that skills account for much more of the cross-country variance when looking at all workers compared to variation within particular occupational skill groups suggests that a main influence of skills is its effect on occupational structure (we explore this further in subsection 4.4).²²

²⁰ These results hold if we control for more technology variables (robot and ICT stocks per worker) and calculate decompositions based on regression results presented in Table G2 in Appendix G. These results are available upon request.

²¹ The within-sector differences in RTI between less and more developed countries are substantial but this effect is of course attributed to other factors.

²² We can conduct a similar variance decomposition analysis for individual-level RTI. We find that explained variance is much smaller (20% for all workers, 8% to 12% for different occupation skill groups). Skills account for the lion's share of explained variance, which is expected given that it is the only category for which we employ individual data. The relative importance of technology, globalization, and structural change are similar to the results for country differences.

Table 5. Decomposition of cross-country variance of RTI by fundamental factors, (% of total variance)

	Technology	Globalization	Structural Change	Supply of skills	Total
All workers	23.4	20.5	-5.4	18.2	56.7
High-skilled occupations (ISCO 1-3)	25.6	9.9	10.4	6.9	52.8
Middle-skilled occupations (ISCO 4-5)	13.5	8.2	0.9	2.5	25.1
Low-skilled occupations (ISCO 7-9)	6.2	21.2	-5.3	1.1	23.3

Note: the contributions of particular factors to RTI variance, σ_k , calculated in line with equation (3) using the model presented in Table 4.

Source: own estimations based on PIAAC, STEP, CULS, World Bank and RIGVC UIBE (2016) data.

Next we report the results of the decomposition analysis of gaps between average RTI in different countries and the US, which we take as the benchmark for advanced countries (Figure 5). We group countries into three types: low- and middle-income countries, the bottom high-income countries (including those in Southern, Central and Eastern Europe, as well as Chile and South Korea), and top high-income countries (mainly in North America, Western Europe, and Australasia, plus Singapore, Japan, and Israel). Decomposition results for gaps between every country and the US are reported in Appendix H.

As can be seen in Table 6, the average RTI scores are much higher for low- and middle-income countries (0.54 US standard deviations) and bottom high-income countries (0.28) than for the top high income group and the US (-0.01 and 0). Low- and middle-income countries have much less computer use (35%) compared to the top high income countries (76%) and the US (75%). In terms of skill, the low- and middle-income countries have fewer older workers and double the share of workers with education level of primary school and below. Notably, 45% of those in low- and middle-income countries are at the lowest literacy level, compared to just 13% and 14% in top high income countries and the US; while 56% and 55% of workers have at least upper medium literacy scores (3-5) in the top high-income countries and the US, compared to 44% in the bottom high-income countries and just 19% in low- and middle-income countries. Finally, integration into global value chains as captured by foreign value added share is highest in the bottom high income group (0.24) compared to just 0.15 in low- and middle-income countries, 0.19 in top high income countries, and 0.08 in the US. The gap in GDP per capita is about 2.5 points in log scale, implying that GDP per capita is more than 250% greater in the top high-income countries compared to the low- and middle-income countries.

The most important factors that contribute to a much higher RTI in low- and middle-income countries (0.55) than in the US are technology, which accounts for more than a third of the gap, and skills, which account for another third (Figure 7). Globalization also plays a notable role, with structural differences relatively unimportant. For bottom high-income countries, the RTI gaps with the US are about half as large on average compared to low- and middle-income countries. However, globalization is most important, followed by technology and skills, with structural change lagging far behind. For the top high-income countries, RTI gaps with the US are negligible.

Table 6. Average levels of RTI and explanatory variables by country groups

	Low and Middle Income Countries	Bottom High Income Countries	Top High Income Countries	US
RTI	0.54	0.28	0.01	0.00
Computer use	0.35	0.60	0.76	0.75
Log of GDP per capita (demeaned)	-1.48	0.12	1.02	1.23
FDI stock/GDP	0.42	1.24	0.79	0.35
FVA Share	0.15	0.24	0.19	0.08
Education: primary	0.32	0.17	0.15	0.10
Education: tertiary	0.34	0.34	0.42	0.42
Literacy skills level: 1 or lower	0.45	0.18	0.13	0.14
Literacy skills level: 3	0.17	0.36	0.41	0.40
Literacy skills level: 4 and 5	0.02	0.08	0.15	0.15
Female	0.47	0.46	0.48	0.49
Age: 16-24	0.16	0.08	0.12	0.15
Age: 35-44	0.24	0.27	0.25	0.22
Age: 45-54	0.20	0.25	0.25	0.23
Age: 55-65	0.10	0.13	0.16	0.18
Agriculture [A]	0.021	0.021	0.008	0.008
Mining [B]	0.014	0.006	0.004	0.005
Manufacturing [C]	0.167	0.191	0.140	0.112
Electricity & Water supply [D+E]	0.018	0.018	0.013	0.010
Construction [F]	0.063	0.086	0.069	0.066
Transportation and storage [H]	0.064	0.057	0.054	0.043
Accommodation and food service [I]	0.053	0.056	0.050	0.072
Information and communication [J]	0.023	0.029	0.040	0.043
Financial and insurance [K]	0.022	0.029	0.037	0.048
Real estate & Professional [L]	0.006	0.008	0.011	0.015
Administrative [M+N]	0.066	0.078	0.098	0.101
Public administration [O]	0.037	0.063	0.059	0.060
Education [P]	0.087	0.089	0.088	0.090
Human health [Q]	0.044	0.062	0.140	0.137
Arts [R]	0.016	0.018	0.021	0.026
Other service [S]	0.051	0.024	0.025	0.032
Activities of household [T]	0.023	0.012	0.004	0.015
Extraterritorial organizations [U]	0.001	0.001	0.000	0.000

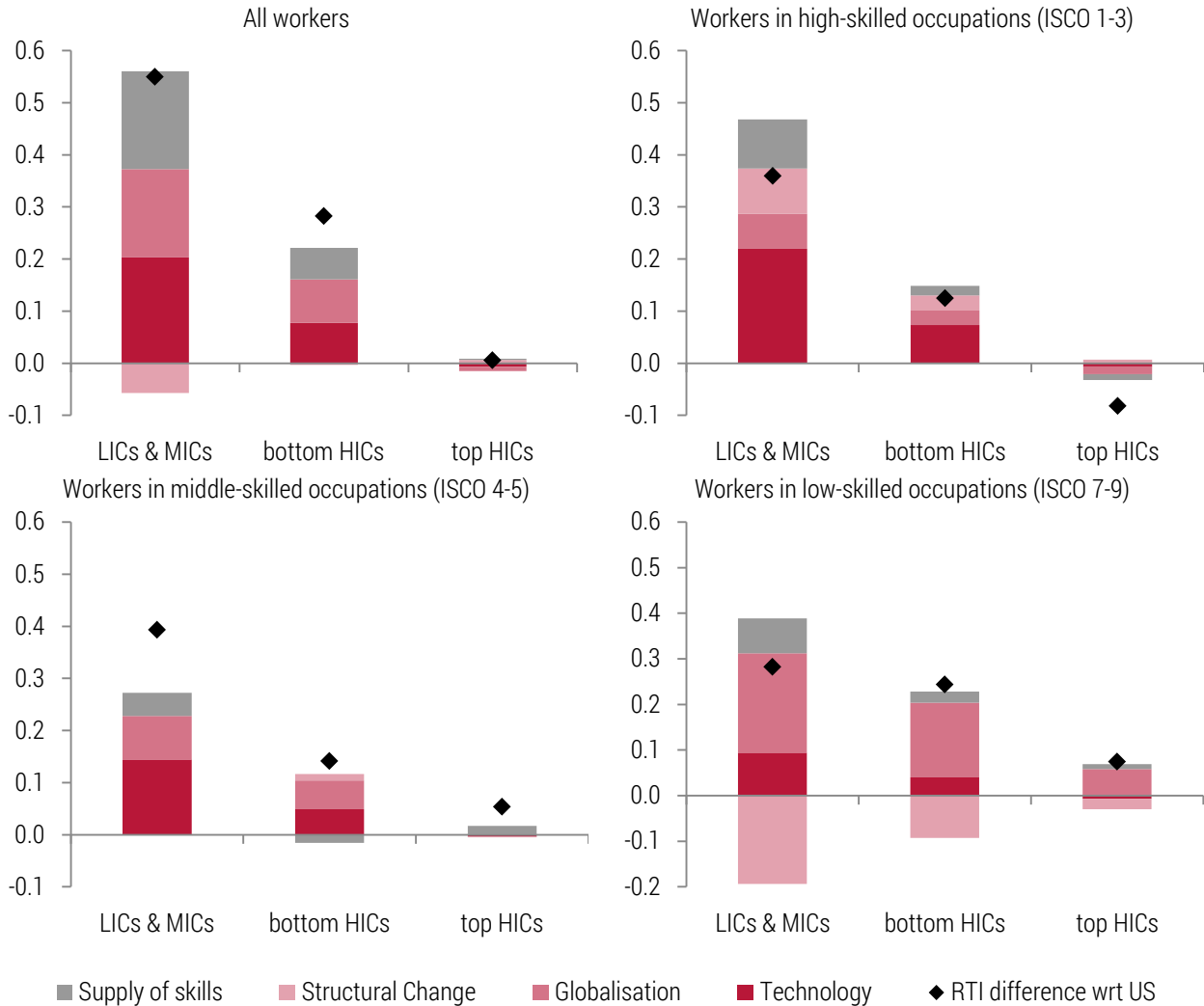
Source: own calculations based on PIAAC, STEP, CULS World Bank, and RIGVC UIBE (2016) data

Next we conduct that decomposition separately for high-, middle-, and low-skilled occupations. Compared to the results for all workers, for those in high-skill occupations, RTI gaps with the US are slightly smaller and technology explains a larger share of the gaps. For low- and middle-income countries, skills and economic structure become more important, consistent with the negative gradient of RTI with GDP per capita for high-skill occupations. For middle-skill occupations, gaps with the US are substantially smaller than for high-skill

occupations. Technology and globalization remain the most- and second-most important factors, while skills and structural change matter relatively little in explaining gaps with the US. Finally, in low-skill occupations, gaps with the US are greater for high-income countries than for high- and middle-skill occupations. Globalization is by far the most important factor, with technology and skills playing minor roles in accounting for gaps with the US.

The relatively greater importance of technology to differences in RTI of high skill occupations is consistent with technology being complementary to non-routine cognitive tasks. The importance of globalization in accounting for gaps in RTI of low-skill occupations is consistent with low-skill occupations involving routine tasks that are more easily outsourced from richer countries to poorer countries.

Figure 5. Regression-based decomposition of differences in RTI between particular countries and the US, by country groups.



*Note: Results of decomposition (3) based on the estimates presented in Table 4, and averaged for country groups defined in Table 3. 0 is set at the US average value and 1 corresponds to one standard deviation of RTI in the US.
Source: own estimations based on PIAAC, STEP, CULS, World Bank and RIGVC UIBE (2016) data.*

4.3 Determinants of task differences in offshorable and non-offshorable occupations

To further investigate how globalization affects the distribution of tasks across countries, in this subsection we study whether the determinants of task differences differ between workers in offshorable and non-offshorable occupations. The share of offshorable jobs ranges from 6% in Canada to 26% in the Czech Republic, and in the majority of countries the routine task intensity is higher in the offshorable occupations (Table F1 in Appendix F). We re-estimate the benchmark specification for these two sub-samples.²³ Results pertaining to regressions of RTI on technology and globalization are presented in Table 7, and complete estimation results are presented in Table G5 in Appendix G.

The results show striking differences in the importance of different factors for offshorable and non-offshorable occupations. First, consistent with expectations, among workers in the non-offshorable occupations, the coefficients on the globalization variable (FVA share) and its interaction with GDP per capita are insignificant, while for workers in offshorable occupations the coefficient on FVA share is much larger and significant (and the coefficient on the interaction with GDP per capita also is larger). Another major difference between offshorable and non-offshorable occupations is the role of technology, which is large and highly significant for non-offshorable occupations but insignificant and close to zero for offshorable occupations.

Table 7. The effects of technology and globalization on routine task intensity (RTI) among workers in offshorable and non-offshorable occupations

	All workers	Workers in non-offshorable occupations	Workers in offshorable occupations
Computer use	-0.508** (0.204)	-0.555*** (0.204)	-0.012 (0.309)
Foreign Value Added (FVA) share	0.269* (0.151)	0.171 (0.147)	0.762*** (0.236)
Ln(GDP per capita) – mean(Ln(GDP per capita))	0.060 (0.040)	0.062 (0.042)	0.015 (0.050)
FVA share * [Ln(GDP pc) – mean(Ln(GDP pc))]	-0.424** (0.197)	-0.396** (0.185)	-0.530* (0.300)
FDI / GDP	0.009* (0.005)	0.012** (0.005)	-0.006 (0.007)
Skills and demographic characteristics	Yes	Yes	Yes
Sector fixed effects	Yes	Yes	Yes
No. of observations	148,120	129,965	18,155
R-Squared	0.220	0.222	0.245

Note: *** p<0.01, ** p<0.05, * p<0.1. Standard errors in parentheses. We use standardized weights that give each country equal weight. The reference levels are: age 25-34, secondary education, wholesale and retail trade; repair of motor vehicles and motorcycles (ISIC G), lower medium literacy skills (level 2). Complete estimation results are presented in Table G5 in Appendix G.

Source: own estimations based on PIAAC, STEP, CULS World Bank, and RIGVC UIBE (2016) data.

²³ We removed from the sample 449 individuals for which we cannot measure offshorability as their occupations were not covered in PDII.

To understand why computer use is unimportant for explaining cross-country differences in RTI of offshorable occupations, it is instructive to look at who are the workers in offshorable occupations. The first large group (42% in our sample) are in manufacturing, mostly consisting of low-skill workers. The other large group are high- and middle-skilled workers in information and communication, and financial and insurance activities. In manufacturing, the probability of computer use is relatively low (50% in our sample compared to the overall average of 59%) and the role of automation and other technologies may be more important. In the other two above mentioned sectors, computer use is very high (over 90%) but hardly varies at all across countries (the coefficient of variation across countries is 1%). This analysis thus suggests that computer use does not help predict task content of offshorable jobs because such jobs are either concentrated in sectors in which computer use is not critical for production or in sectors in which virtually all workers are using them. The importance of variables other than those related to technology and globalization are similar for both groups of workers (see Table G5 in Appendix G).

As before, we can decompose the cross-country variance in RTI as well as the gaps with the US in average RTI separately by occupation group (offshorable versus non-offshorable). The variance decomposition results confirm the different importance of fundamental factors in offshorable and non-offshorable occupations apparent in the RTI regression results. Among workers in offshorable occupations, globalization contributes the most to the cross-country variation in RTI (17%), followed by structural change (12%) while the contribution of technology is negligible (Table 8). Skills are also important (14%). On the other hand, among workers in non-offshorable occupations, the contributions of technology and skills explain the most cross-country variation in RTI (26% and 19%) while the contribution of globalization is relatively smaller (18%), but still noticeable.

Similarly, technology accounts for the largest share of RTI gaps with the US for those in non-offshorable occupations, while globalization accounts for the largest share of the gaps in offshorable occupations (Figure 7). Because offshorability should matter the most for tradeable goods, we also separately analyze the determinants of gaps with the US for workers in the manufacturing sector, and indeed find the differences between offshorable and non-offshorable occupations to be even more pronounced in manufacturing, especially for gaps between the US and low- and middle-income countries (Figure 7). Nearly all of the gaps in RTI with the US are attributable to globalization, but most of the differences in non-offshorable occupations are due to skills (Figure 7).²⁴

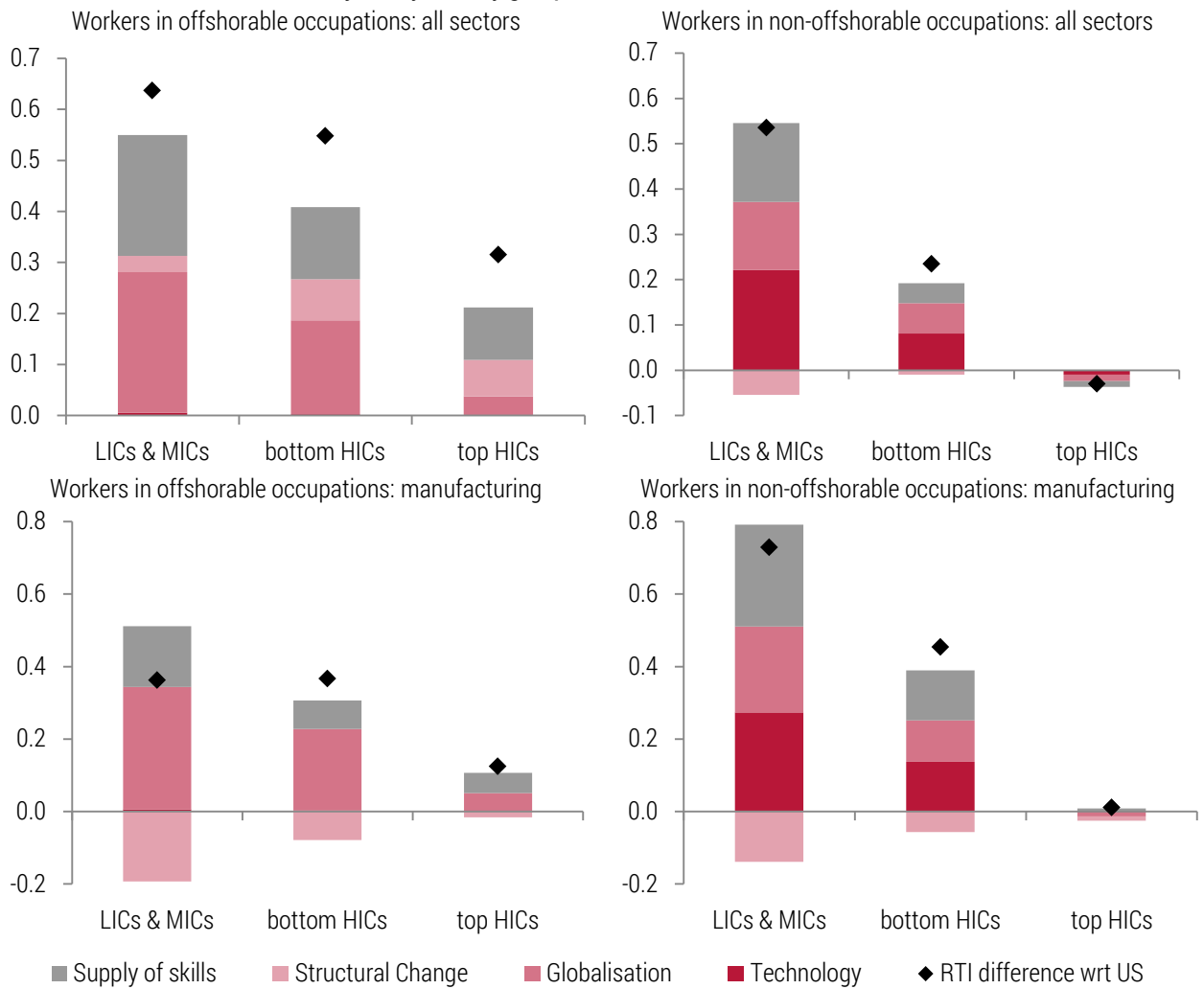
Table 8. Decomposition of cross-country RTI variance in offshorable and non-offshorable occupations (% of total variance)

	Technology	Globalization	Structural Change	Supply of skills	Total
Non-offshorable occupations	26.2	18.5	-4.6	18.9	58.9
Offshorable occupations	0.2	17.4	12.0	14.2	43.9

Note: the contributions of particular factors to RTI variance, σ_k , calculated in line with equation (3) using the model presented in Table 7. Source: own estimations based on PIAAC, STEP, CULS, World Bank and RIGVC UIBE (2016) data.

²⁴ In sectors other than manufacturing, the contribution of globalization is much smaller and differences in skills play a larger role even for offshorable occupations. These results are available upon request.

Figure 6. Regression-based decomposition of differences in RTI between particular countries and the US among workers in offshorable and non-offshorable jobs, by country groups.



Note: Results of decomposition (3) based on the estimates presented in Table 7, and averaged for country groups defined in Table 3. 0 is set at the US average value and 1 corresponds to one standard deviation of RTI in the US.

Source: own estimations based on PIAAC, STEP, CULS, World Bank and RIGVC UIBE (2016) data.

4.4 Assessing the role of occupations

The results thus far have shown that most of the cross-country differences in job tasks are associated with differences in technology, globalization, supply of skills and sectoral structure. Having individual survey data on job tasks enables us to study the correlates of RTI without making any assumptions about the nature of work in different occupations. Nonetheless, given that much research on the nature of work focuses on occupations, it is of interest to investigate how much of the above-documented relationship between routine-intensity and the four fundamental forces is explained by differences in occupational structure and how much is due to differences in tasks within occupations. To examine this question, we expand the benchmark specification of task regression by adding occupation fixed effects, τ_o :

$$RTI_{ijsc} = \beta_0 + \beta_1 Z_{sc} + \beta_2 G_{sc} + \lambda_s + \beta_3 E_{ijsc} + \tau_o + \varepsilon_{ijsc}, \quad (5)$$

In this specification, the coefficients on the variables for the four main factors capture their influence among workers in the same occupation. Thus, by comparing the coefficients with those estimated using the baseline specification without occupation fixed effects, we can infer how much of the relationship between the four factors is captured by their impact on occupational structure and how much is a within-occupation association.

The coefficients on the occupation dummies are in line with intuition: workers in high-skilled occupations (ISCO 1-3) perform less routine-intensive tasks than clerical workers (ISCO 4), while workers in low-skilled occupations (ISCO 7-9) and sales and services workers (ISCO 5) perform more routine-intensive tasks (Table 9). However, the four fundamental factors still strongly predict differences in routine intensity even after controlling for occupation. Although the absolute sizes of the coefficients pertaining to education, literacy, and computer use are somewhat smaller than in the benchmark specification (Table 4), none of them loses statistical significance. The coefficients pertaining to globalization variables change little and remain significant.

Table 9. The correlates of routine task intensity (RTI) at the worker level, including occupations

	All workers	High-skilled occupations (ISCO 1-3)	Middle-skilled occupations (ISCO 4-5)	Low-skilled occupations (ISCO 7-9)	Offshorable occupations	Non-offshorable occupations
Computer use	-0.403** (0.192)	-0.739*** (0.174)	-0.363 (0.311)	-0.216 (0.221)	-0.266 (0.270)	-0.446** (0.187)
Foreign Value Added share	0.359** (0.141)	-0.020 (0.140)	0.188 (0.187)	0.783*** (0.171)	0.783*** (0.212)	0.292** (0.139)
Ln(GDP per capita) – mean(Ln(GDP per capita))	0.017 (0.073)	-0.056 (0.076)	0.016 (0.104)	0.014 (0.071)	0.043 (0.093)	0.027 (0.070)
FVA share * [Ln(GDP pc) – mean(Ln(GDP pc))]	-0.332* (0.183)	-0.181 (0.169)	-0.228 (0.211)	-0.280 (0.227)	-0.411 (0.252)	-0.328* (0.178)
FDI / GDP	0.008 (0.005)	0.021*** (0.007)	0.010 (0.008)	-0.014*** (0.005)	0.002 (0.006)	0.009* (0.005)
Education: primary	0.147*** (0.018)	0.147*** (0.030)	0.216*** (0.022)	0.113*** (0.025)	0.173*** (0.036)	0.144*** (0.018)
Education: tertiary	-0.186*** (0.015)	-0.204*** (0.017)	-0.190*** (0.021)	-0.120*** (0.034)	-0.133*** (0.035)	-0.191*** (0.015)
Literacy skills level: 1 or lower	0.036** (0.017)	0.032 (0.028)	0.046* (0.025)	0.056** (0.024)	0.089** (0.038)	0.028 (0.018)
Literacy skills level: 3	-0.066*** (0.012)	-0.075*** (0.016)	-0.057*** (0.022)	-0.035 (0.023)	-0.063** (0.030)	-0.067*** (0.013)
Literacy skills level: 4 and 5	-0.155*** (0.015)	-0.165*** (0.020)	-0.058** (0.029)	-0.151*** (0.041)	-0.176*** (0.039)	-0.151*** (0.017)
Female	0.230*** (0.011)	0.215*** (0.012)	0.213*** (0.018)	0.274*** (0.024)	0.320*** (0.026)	0.218*** (0.012)
Age: 16-24	0.177*** (0.015)	0.198*** (0.028)	0.204*** (0.026)	0.127*** (0.023)	0.169*** (0.040)	0.179*** (0.016)
Age: 35-44	-0.030*** (0.010)	-0.038*** (0.013)	-0.021 (0.018)	-0.029 (0.021)	-0.063** (0.025)	-0.028*** (0.011)

Age: 45-54	0.010 (0.012)	-0.032** (0.016)	0.017 (0.023)	0.049** (0.021)	-0.017 (0.025)	0.012 (0.013)
Age: 55-65	0.056*** (0.013)	-0.015 (0.019)	0.111*** (0.026)	0.090*** (0.023)	0.087*** (0.032)	0.047*** (0.014)
ISCO 1	-0.786*** (0.023)	-	-	-	-0.852*** (0.073)	-0.792*** (0.026)
ISCO 2	-0.599*** (0.023)	0.195*** (0.018)	-	-	-0.629*** (0.038)	-0.589*** (0.027)
ISCO 3	-0.351*** (0.020)	0.431*** (0.019)	-	-	-0.233*** (0.042)	-0.358*** (0.024)
ISCO 5	0.112*** (0.022)	-	0.119*** (0.022)	-		0.119*** (0.025)
ISCO 7	0.208*** (0.026)	-	-	-0.442*** (0.030)	0.317*** (0.043)	0.193*** (0.031)
ISCO 8	0.546*** (0.027)	-	-	-0.105*** (0.028)	0.546*** (0.038)	0.559*** (0.040)
ISCO 9	0.621*** (0.025)	-	-	-	-	0.627*** (0.028)
Sector fixed effects	Yes	Yes	Yes	Yes	Yes	Yes
Observations	148,569	62,907	47,373	38,289	18,155	129,965
R-squared	0.317	0.151	0.092	0.114	0.321	0.319

Note: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$. Standard errors in parentheses. We use standardized weights that give each country equal weight. The reference levels are: age 25-34, secondary education, wholesale and retail trade; repair of motor vehicles and motorcycles (ISIC G), lower medium literacy skills (level 2). Clerical support workers (ISCO 4) are the reference group in the regressions for all workers, for middle-skilled occupations and for offshorable and non-offshorable occupations. Managers (ISCO 1) and Elementary occupations (ISCO 9) are the reference groups in regressions for high-skilled and low-skilled occupations, respectively.

Source: own estimations based on PIAAC, STEP, CULS, World Bank, and RIGVC UIBE (2016) data.

We can also conduct the cross-country variance decomposition and decomposition of gaps with the US adding the occupation structure as an additional factor. We find that occupations have a noticeable contribution to the cross-country variation in RTI for all workers (17%). However, they explain less than one third of the explained variation of RTI across countries (56%), which is virtually the same as in the specification with no occupational fixed effects (57%) (Table 10). The contributions attributed to other factors, especially to the supply of skills, are somewhat lower than for the benchmark specification (Table 5). Still, the contribution of technology remains larger than the contribution of occupations. When we analyze the importance of occupations separately for high-, middle-, and low-skill occupation groups (not reported here), we find that the occupation dummies have very little explanatory power, suggesting that only differences in broad occupation group categories is meaningful for explaining task content differences across countries.

Table 10. Decomposition of cross-country variance of RTI by fundamental factors, controlling for occupations (% of total variance)

	Technology	Globalization	Structural Change	Supply of skills	Occupations	Total
With occupation	18.9	16.0	-3.0	8.6	16.8	57.2
Without occupation dummies	23.4	20.5	-5.4	18.2	-	56.7

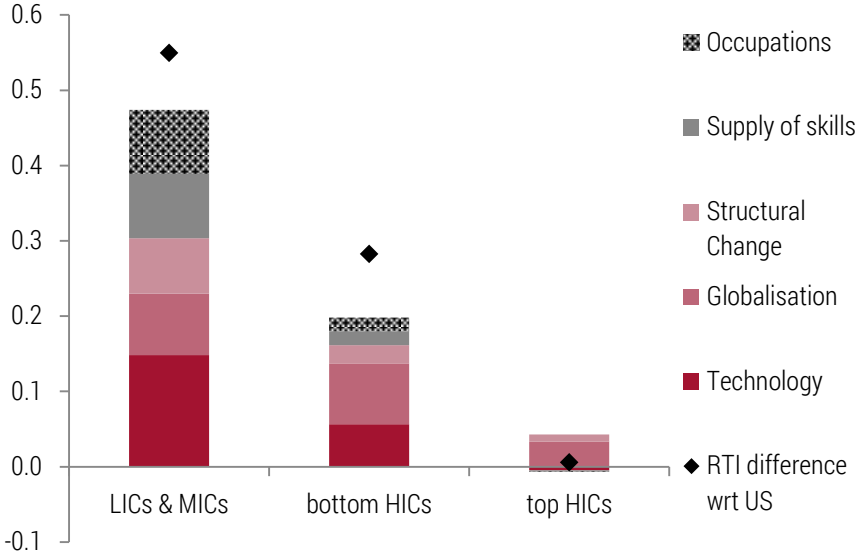
Note: the contributions of particular factors to RTI variance, σ_k , calculated in line with equation (3) using the model presented in Table 9.

Source: own estimations based on PIAAC, STEP, CULS, World Bank and RIGVC UIBE (2016) data.

In explaining gaps in RTI between low- and middle-income countries and the US, we find that occupational structure on average explains only about one fifth of the explained gap, with four fifths of the gap explained by the association of fundamental factors and within-occupation differences in RTI across countries.

Overall, we find, perhaps surprisingly, that occupations are of limited importance in explaining cross-country differences in routine task intensity, and that most of the association between fundamental factors and RTI appear in differences within occupation groups. This highlights the importance of collecting survey-based task data to understand the nature of work in specific countries.

Figure 7. Regression-based decomposition of differences in RTI between particular countries and the US, controlling for occupations, by country groups.



*Note: Results of decomposition (4) based on the estimates presented in Table 9, and averaged for country groups defined as in Table 4. 0 is set at the US average value and 1 corresponds to one standard deviation of RTI in the US.
Source: own estimations based on PIAAC, STEP, CULS, World Bank and RIGVC UIBE (2016) data.*

5. Summary and conclusions

We have developed a novel dataset that measures the task content of jobs at the individual worker level for a large number of countries at different stages of economic development. The new survey-based measures are validated to be consistent with US O*NET-based task content measures that have been widely used in the existing literature on job tasks. A key advantage of the new measures is that they can distinguish between differences in task content among workers who have the same occupation but live in different country environments.

Our results show that there are substantial cross-country differences in the routine-intensity of job tasks, both at the national level and within specific occupations. The differences in tasks across countries at different stages of development are much greater than could be explained by differences in occupational structure. Not surprisingly, work in the most developed countries involves the most non-routine cognitive analytical and non-routine cognitive interpersonal tasks, and often have the least manual tasks, while the opposite is true for developing and emerging economies. Routine cognitive tasks are lowest in the least and most developed countries, and highest

in Eastern and Southern European countries, suggesting an inverse U-shaped relationship between the role of routine cognitive work and development level. Moreover, cross-country differences in routine task intensity are most strongly related to the differences in GDP per capita for high-skilled occupations, with no systematic correlation for middle- and low-skill occupations.

We have estimated a regression that captures the association between the relative routine task intensity (RTI) of jobs and four fundamental forces: technology, globalization, structural change, and supply of skills. We have used these results to decompose the extent to which cross-country differences in relative routine task intensity are statistically associated with these different factors, both in terms of cross-country variance in mean RTI and RTI gaps between the US and groups of countries sorted by GDP per capita. Consistent with much recent literature emphasizing the influence of technology on the nature of work, we find that technology plays the largest role in explaining cross-country differences in RTI, followed closely by skills and globalization. Structural change has the least explanatory power. However, we have also find interesting heterogeneities in the impact of these factors for different types of occupations. Technology matters the most for high-skill occupations, consistent with the complementarity between technology and non-routine cognitive tasks, while globalization matters the most for low-skill occupations which are more likely to involve routine tasks that are more easily outsourced from rich countries to poor countries. Similarly, technology matters most for non-offshorable jobs, while globalization matters most for offshorable jobs.

Our work stresses the need to quantify the country-specific task content of jobs and identify differences between occupational task content in countries at different stages of development. It paves the way for a more comprehensive research on the distribution of tasks around the world that can account for the within-occupation and between-country variation in task demand.

References

- Acemoglu, D., Autor, D. H. (2011). Skills, Tasks and Technologies: Implications for Employment and Earnings. In: Card, D. and Ashenfelter, O. (eds). *Handbook of Labor Economics*. Amsterdam: Elsevier, 1043–1171.
- Akerman, A., Gaarder I., Mogstad M. (2015). The Skill Complementarity of Broadband Internet. *The Quarterly Journal of Economics*, 130, (4), 1781-1824.
- Arias, O.S., Sánchez-Páramo, C., Dávalos, M.E., Santos, I., Tiongson, E.R., Gruen, C., de Andrade Falcão, N., Saiovici, G., Cancho, C.A. (2014). Back to work: growing with jobs in Eastern Europe and Central Asia, *Europe and Central Asia reports*. Washington, DC.: The World Bank.
- Arntz, M., Gregory, T., Zierahn, U. (2017). Revisiting the risk of automation, *Economics Letters* 159 (C), 157-160.
- Autor, D. (2013) "The "task approach" to labor markets : an overview, *Journal for Labour Market Research*, 46(3), 185-199.
- Autor, D., Dorn, D. (2009). This Job is "Getting Old": Measuring Changes in Job Opportunities using Occupational Age Structure. *American Economic Review: Papers and Proceedings*, 99 (2), 45-51.
- Autor, D. Dorn, D. (2013). The Growth of Low-Skill Service Jobs and the Polarization of the US Labor Market, *American Economic Review*, 103 (5), 1553–97.
- Autor, D., Katz, L. F., Kearney, M.S. (2006). The Polarization of the U.S. Labor Market. *American Economic Review Papers and Proceedings* 96, 2.
- Autor, D., Handel, M. (2013). Putting Tasks to the Test: Human Capital, Job Tasks, and Wages. *Journal of Labor Economics*, 31(2), 59-96.
- Autor, D. H., Levy, F., Murnane, R. (2003). The Skill Content of Recent Technological Change: An Empirical Exploration. *Quarterly Journal of Economics* 118, 4.
- Bárány, Z. L., Siegel C. 2018. Job Polarization and Structural Change. *American Economic Journal: Macroeconomics*, 10 (1): 57-89.
- de la Rica, S., Gortazar, L. (2016). Differences in Job De-Routinization in OECD Countries: Evidence from PIAAC, IZA Discussion Paper No. 9736.
- Deming, D. (2017). The Growing Importance of Social Skills in the Labor Market, *Quarterly Journal of Economics* 2017;132 (4) :1593-1640.
- Dicarlo, E., Bello, S. L., Monroy-Taborda, S., Oviedo A. M., Sanchez-Puerta, M. L., Santos, I. (2016). The Skill Content of Occupations across Low and Middle Income Countries: Evidence from Harmonized Data. *IZA Discussion Paper* No. 10224.
- Du, Y., Jia, P. and Cheng, J. (2017) Structural Changes in Labor Market, Job Tasks and Skill Demand. *Studies in Labor Economics* 5(3), 30-49.
- Du Y., Park A. (2018), Changing Demand for Tasks and Skills in China, Background Report for the World Bank Group (WBG)-Development Research Center under the State Council (DRC) report on New Drivers of Growth in China.
- Eden, M., Gaggl, P. (2015). Do poor countries really need more IT? The role of relative prices and industrial composition. Policy Research Working Paper Series 7352, The World Bank.
- Firpo, S., Fortin, N. M., Lemieux, T., (2011). Occupational Tasks and Changes in the Wage Structure, IZA Discussion Papers 5542, Institute for the Study of Labor (IZA)

- Frey, C. B., Osborne, M. A. (2017). The future of employment: How susceptible are jobs to computerisation?, *Technological Forecasting and Social Change* 114(C), 254-280.
- Goos, M., Manning, A., Salomons, A. (2014). Explaining Job Polarization: Routine-Biased Technological Change and Offshoring. *American Economic Review* 104, 2509–2526.
- Grossman, G., Rossi-Hansberg E. (2008). Trading Tasks: A Simple Theory of Offshoring, *American Economic Review* 98(5): 1978-1997.
- Hardy, W., Keister, R., Lewandowski, P. (2018). Educational upgrading, structural change and the task composition of jobs in Europe. *Economics of Transition*, 26 (2), 201–231.
- Hsieh CT., Klenow P. J., (2010). Development Accounting, *American Economic Journal: Macroeconomics* 2(1), 207-223.
- Hummels, D., Munch J. R., Xiang C. (2018). Offshoring and Labor Markets, *Journal of Economic Literature*, 56 (3): 981-1028.
- Jensen, J. and Kletzer, L. G. (2010). Measuring Tradable Services and the Task Content of Offshorable Services Jobs, p. 309-335 in: *Labor in the New Economy*, National Bureau of Economic Research, Inc
- Lewandowski, P. Keister, R., Hardy, W., Górka S. (2017). Routine and Ageing? The Intergenerational Divide in the Deroutinisation of Jobs in Europe. IBS Working Paper 01/2017.
- Macdonald, K. (2014). PV: Stata module to perform estimation with plausible values. *Statistical Software Components* from Boston College Department of Economics.
- McKinsey, (2017), Future that works: AI, automation, employment, and productivity, McKinsey Global Institute
- Marcolin, L., Miroudot, S., Squicciarini, M. (2016). The Routine Content of Occupations: New Cross-country Measures Based on PIAAC. OECD Science, Technology and Industry Working Papers, 2016/02, OECD Publishing, Paris.
- Michaels, G., Natraj, A., Van Reenen, J. (2014) Has ICT polarized skill demand? Evidence from eleven countries over twenty-five years. *Review of Economics and Statistics* 96, 60–77.
- Montesor G. (2018) Job polarization and labour supply changes in the UK, *Labour Economics* (forthcoming), <https://doi.org/10.1016/j.labeco.2018.05.009>
- Morduch, J., Sicular, T. (2002). Rethinking Inequality Decomposition, With Evidence from Rural China. *The Economic Journal*, 112: 93-106
- Oldenski, L. (2012). Export Versus FDI and the Communication of Complex Information, *Journal of International Economics* 87(2): 312-322
- OECD (2016). Technical Report of the Survey of Adult Skills (PIAAC). 2nd Edition.
- PIAAC Literacy Expert Group, (2009), PIAAC Literacy: A Conceptual Framework, No 34, OECD Education Working Papers, OECD Publishing, <https://EconPapers.repec.org/RePEc:oea:eduab:34-en>.
- Reijnders, S.M., de Vries, G.J., (2018). Technology, offshoring and the rise of non-routine jobs, *Journal of Development Economics* 135(C), 412-432.
- RIGVC UIBE (2016). UIBE GVC Index. Research Institute for Global Value Chains, University of International Business and Economics
- Rodrik D. (2018) New Technologies, Global Value Chains, and the Developing Economies. 2018. University of Oxford, Pathways for Prosperity Commission, September 2018

- Rubin, D.B. (1987) Multiple Imputation for Nonresponse in Surveys. John Wiley and Sons Inc., New York.
- Salvatori, A. (2018). The Anatomy of Job Polarisation in the UK. *Journal for Labour Market Research*, 52:8
- Spitz-Oener, A. (2006). Technical Change, Job Tasks, and Rising Educational Demands: Looking outside the Wage Structure. *Journal of Labor Economics* 24, 235–270.
- Wang, Z., Wei, S., Yu, X., Zhu, K. (2017) Measures of Participation in Global Value Chains and Global Business Cycles. *NBER Working Paper* No. 23222, National Bureau of Economic Research, Inc.
- World Bank (2019) World Development Report 2019: Digital Dividends. Washington, DC

Appendices

Appendix A. List of countries in PIAAC, STEP and CULS

PIAAC surveys include publically available data representative of 32 countries. 23 in Round I: Austria, Belgium (Flanders), Canada, Cyprus (the area under the effective control of the Government of the Republic of Cyprus), Czechia, Denmark, Estonia, Finland, France, Germany, Ireland, Italy, Japan, South Korea, Netherlands, Norway, Poland, Russia (w/o Moscow municipal area), Slovakia, Spain, Sweden, UK (England and Northern Ireland), United States. 9 in Round II: Chile, Greece, Indonesia (Jakarta), Israel, Lithuania, New Zealand, Singapore (only permanent residents), Slovenia and Turkey. Moreover, a dataset with supplementary 2nd round is available for the United States via the US National Center for Education Statistics (NCES).

We use STEP surveys for 9 countries: Armenia, Bolivia (four main capital cities – La Paz, El Alto, Cochabamba and Santa Cruz de la Sierra), Colombia (13 main metropolitan areas), Georgia (w/o Abkhazia and South Ossetia), Ghana, Kenya, Lao PDR (both urban and rural areas), Macedonia, Ukraine.

The 3rd wave of the CULS survey includes data on individuals in six large cities in China: Guangzhou, Shanghai, and Fuzhou on the coast, Shenyang in the northeast, Xian in the northwest, and Wuhan in central China.

Appendix B. Construction of task content measures based on US PIAAC and US O*NET

To construct the reference task content measures proposed by Acemoglu and Autor (2011), we use the Occupational Information Network (O*NET) database which contains extensive information on the occupations in the US. We merge the O*NET data with the US PIAAC data using the occupational crosswalks prepared by the O*NET Resource Center, the U.S. Bureau of Labor Statistics and the National Crosswalk Service Center, and adapted to the ISCO classification of occupations by Hardy et al. (2018).²⁵ ISCO is used in PIAAC, and 3-digit or 4-digit codes are available in the US PIAAC.²⁶ We apply our procedure at each level.

To calculate the task content of occupations, we follow Acemoglu and Autor (2011). First, we standardize the values $t_{j^o,i}$ of each task item j^o in the set of O*NET task items J^o , using the means ($\bar{t}_{j^o}^{US}$) and standard deviations ($\delta_{j^o}^{US}$) in the US PIAAC:

²⁵ See: www.ibs.org.pl/resources [accessed: 2017-05-04].

²⁶ The dataset with 3-digit ISCO codes is available for researchers from National Center for Education Statistics. The 4-digit ISCO codes are included in the restricted dataset at the American Institutes for Research who have kindly run our code.

$$\forall_i \forall_{j^o \in J^o} t_{i,j^o}^{std} = \frac{t_{i,j^o} - \bar{t}_{j^o}^{US}}{\delta_{j^o}^{US}}, \quad (B1)$$

whereby i is a worker-level observation in the US PIAAC data. The set of O*NET task items, J^o , is presented in Appendix B. Second, we construct four task content measures: non-routine cognitive analytical, non-routine cognitive interpersonal, routine cognitive, and manual. Each task content measure is calculated as a sum of constituent task items (Table B1), except for the manual measure which is the sum of all items that define routine and non-routine manual task content measures in Acemoglu and Autor (2011). Each of these sums is then standardized to have a mean 0 and standard deviation 1 in the US PIAAC sample. Using one measure of manual tasks is not a limitation because the correlation between the non-routine and routine manual tasks in the US PIAAC is very high (85% across 3-digit ISCO occupations and 88% across 2-digit occupations).²⁷

Table B1. Set of O*NET items, J^o , used in Acemoglu and Autor (2011) task contents measures

Task content measure (T)	Task items (J)
Non-routine cognitive analytical	Analysing data/information Thinking creatively Interpreting information for others
Non-routine cognitive interpersonal	Establishing and maintaining personal relationships Guiding, directing and motivating subordinates Coaching/developing others
Routine cognitive	The importance of repeating the same tasks The importance of being exact or accurate Structured vs. unstructured work
Routine manual	Pace determined by the speed of equipment Controlling machines and processes Spending time making repetitive motions
Non-routine manual physical	Operating vehicles, mechanized devices, or equipment Spending time using hands to handle, control or feel objects, tools or controls Manual dexterity Spatial orientation

Source: Own elaboration based on Acemoglu and Autor (2011).

PIAAC and STEP surveys provide data on the job tasks performed by workers. In first step, we identified the set of potential items, $J^P = \{J^{P,NRCA}, J^{P,NRCP}, J^{P,RC}, J^{P,M}\}$, that are available in both surveys in the identical or almost identical form, and which could potentially be used to derive particular task content measures (see Table B2, and Appendix C for the full wording of questions and allowed answers). We chose between three and eight potential items for particular task content measures, except the manual content for which only one item (“working physically”) is available in both STEP and PIAAC. We reverse the values of three variables considered for the routine cognitive measure (“changing order of tasks”, “solving problems”, “giving presentations”), so the higher is the value, the less common is a given phenomenon. To ensure comparability between STEP and PIAAC data, we rescale the answers to achieve the same value ranges. In particular, for PIAAC questions with five possible answers, except for “changing order of tasks”, “solving problems”, and “giving presentations”, we consider four variants of binary variables, based on cutoffs available in the original answers (see Appendix C for details).

²⁷ Studies on the US (Autor and Price, 2013) or European countries (Lewandowski et al., 2017) found that routine and non-routine manual tasks are also highly correlated over time and follow similar trends.

Our selection of questions is based on the similarities between PIAAC / STEP items and the O*NET items, and attributes of particular type of work (Autor, 2013). It is also consistent with the selections of Dicarolo et al. (2016), de la Rica and Gortazar (2016) or Marcolin et al. (2016), shown in Appendix D. However, contrary to these authors, we don't assign task items to task contents in an arbitrary way, but search for subsets of questions and cutoffs which provide the best proxy for the O*NET tasks in the US.

Formally, we consider every subset of different questions allowed for a particular task measure:

$$t \in T = \{x_{c_1}^{j_1}, \dots, x_{c_l}^{j_l} : 2 \leq k \leq r_R, j_1 \neq j_2 \neq \dots \neq j_k, j_i \in J^{P,R}\}, \quad (B2)$$

Where r_R is the number of questions considered for particular task content $R \in \{NRCA, NRCP, RC, M\}$ (Table B2), l is the number of variable variants c available for each question (1 or 4), and x are the values. Note that we don't allow two variants of the same question in a given subset, and consider only subsets with at least two variables. The total number of subsets considered for particular task content measures is shown in Table B2.

In the next step, we adapt the Acemoglu and Autor (2011) methodology to the PIAAC items. We standardize the worker-level values $x_{c_l,i}^{j_k}$ using the means ($\bar{x}_{c_l}^{j_k,US}$) and standard deviations ($\delta_{c_l}^{j_k,US}$) in the US:

$$\forall_i \forall_{j_k \in J^{P,R}} x_{c_l,i}^{j_k, std} = \frac{x_{c_l,i}^{j_k} - \bar{x}_{c_l}^{j_k,US}}{\delta_{c_l}^{j_k,US}}, \quad (B3)$$

For each subset, we sum these standardized values and standardize those sums again within the US dataset. Then, we calculate (weighted) averages of these subset-specific values at the level of 3-digit and 4-digit ISCO occupations. Finally, we calculate the correlations between these occupation-specific averages and the relevant O*NET-based task content measures across 3-digit and 4-digit ISCO occupations in the US.

Table B2. PIAAC and STEP questions considered for the measurement of particular task content measures, with number of variable variants (in brackets)

Task content	Non-routine cognitive analytical ($J^{P,NRCA}$)	Non-routine cognitive personal ($J^{P,NRCP}$)	Routine cognitive ($J^{P,RC}$)	Manual ($J^{P,M}$)
Task items	Solving problems (1) Reading bills (4) Reading news (4) Reading professional journals (4) Advanced math (4) Calculating prices (4) Calculating fractions (4) Programming (4)	Supervising (1) Collaborating (1) Making speeches or giving presentations (4)	Changing order of tasks - reversed (1) Reading bills (4) Filling forms (4) Calculating fractions (4) Solving problems - reversed (1) Making speeches or giving presentations - reversed (4)	Physical tasks (1)
No. of subsets	156 221	18	4 982	1

Note: 1 and 4 identify variables for which we use original questions (1), or four variants of binary variables based on cutoffs available in original question (4). For each task content measure except the manual measure, we consider only combinations that include at least two questions. Last row shows the number of subsets of variables considered for given task content measure.

Source: Own elaboration.

For each task content measure, we use the following criteria to select the best subset of PIAAC items:

- We consider five subsets with the highest correlations with the relevant O*NET-based measure at the 3-digit, or at the 4-digit level of ISCO.
- A particular subset can be preferred over a subset with the higher correlation at the 4-digit level only if it has a higher correlation at the 3-digit level.
- The reversed version of variables used in the measure of routine cognitive tasks should use the same cutoffs as the original variables used in the measures of non-routine cognitive tasks.
- We allowed changes in the cutoffs if it increased the correlation at a 3-digit occupation level without a meaningful drop in the correlation at a 4-digit level, and if it mitigated any systematic differences between the task content measures calculated in PIAAC and STEP surveys.

Finally, in order to verify whether the values of task contents do not depend on the data source (PIAAC or STEP), we estimate a range of OLS regressions. In the base model, we regress (OLS) each task content measure against individual characteristics (gender, 10-year age groups, education, 1-digit occupations, sectors) and the STEP survey fixed effect which turns out negative and significant for all tasks except non-routine cognitive personal (Table B3). When we control for the level of literacy skills and GDP per capita,²⁸ the difference between STEP and PIAAC remains significant only in the case of manual tasks. This shows our survey measures of cognitive tasks are consistent and comparable between the two surveys. However, the STEP fixed effect remains significant even in the most elaborate specification. Therefore, we correct the values of manual task scores in STEP by this fixed effect (we add 0.17 to the manual task score of each individual in STEP sample).

Table B3. OLS regressions of task measures on sets of control variables and a STEP dummy

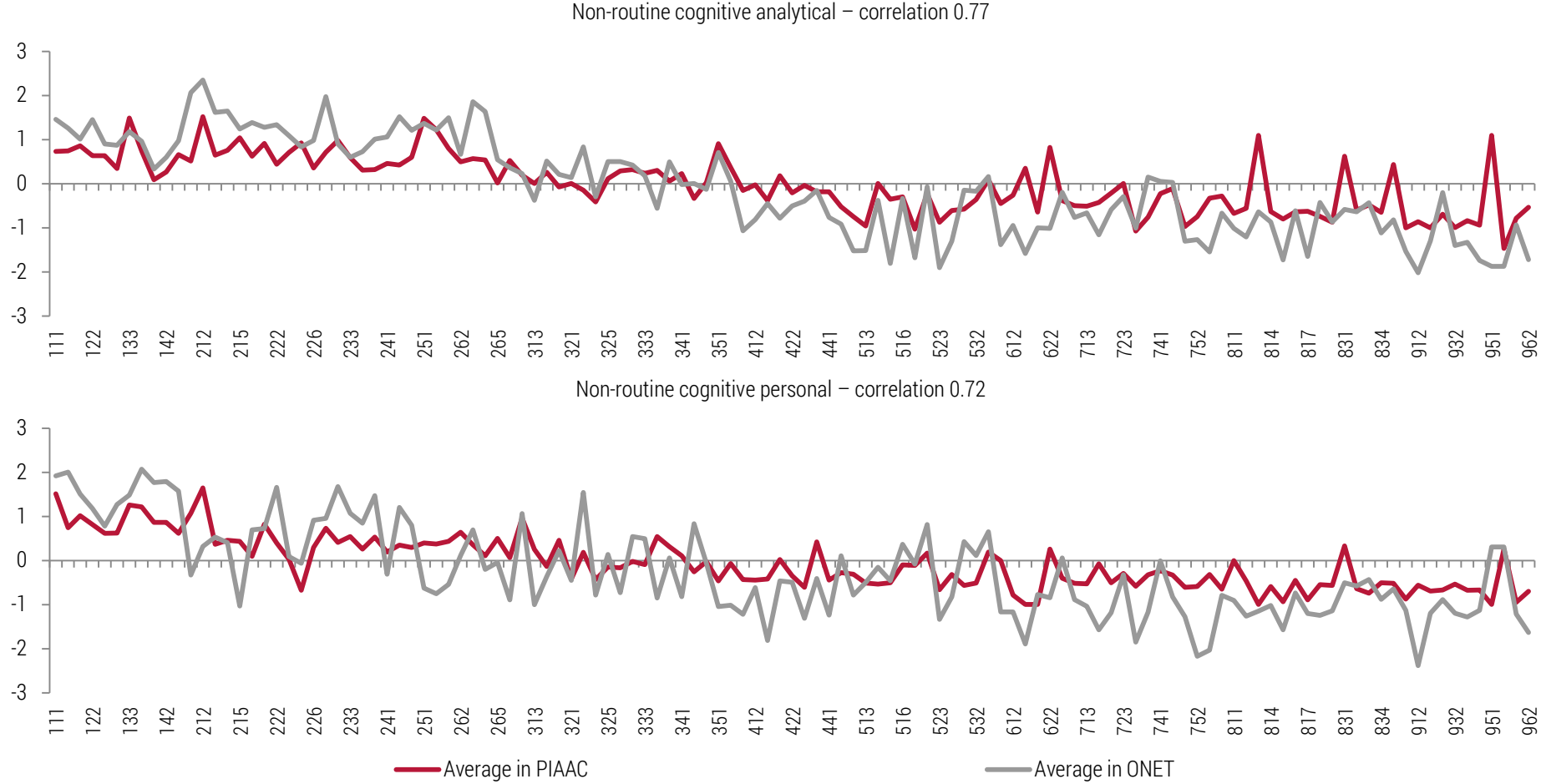
	Non-routine cognitive analytical	Non-routine cognitive personal	Routine cognitive	Manual
Base model, total sample of 42 countries				
STEP dummy	-0.22***	-0.03	-0.05	-0.38***
Base model, subsample of 39 countries with literacy assessment data				
STEP dummy	-0.17**	-0.08	-0.17	-0.39***
Base model + controls for literacy skills and for GDP per capita, subsample of 39 countries with literacy assessment data				
Literacy skills level: 0 and 1	-0.10***	-0.04***	-0.02	0.02
Literacy skills level: 3	0.08***	0.05***	-0.09***	-0.14***
Literacy skills level: 4 and 5	0.16***	0.11***	-0.22***	-0.30***
GDP per capita	-0.95	-1.51***	1.41	0.27
GDP per capita squared	0.05	0.08***	-0.07	-0.01
STEP dummy	-0.00	0.06	-0.07	-0.18***

Note: the base regressions include dummies for gender, 10-year age groups, education, 1-digit occupations and sectors. To save space, we report only the coefficients for the STEP dummy, literacy skills and GDP per capita (in 1000s, in PPP, current international \$, country averages for 2011-2016). The regressions with literacy scores exclude China (CULS), Laos and Macedonia due to lack of literacy skills assessment in these countries. The total number of observations equals around 155,500 for the base model regression with all countries and around 144,500 for the specifications without China (CULS), Laos and Macedonia. The standard errors are clustered at a country level.

Source: own estimations based on PIAAC, STEP, CULS and World Bank data.

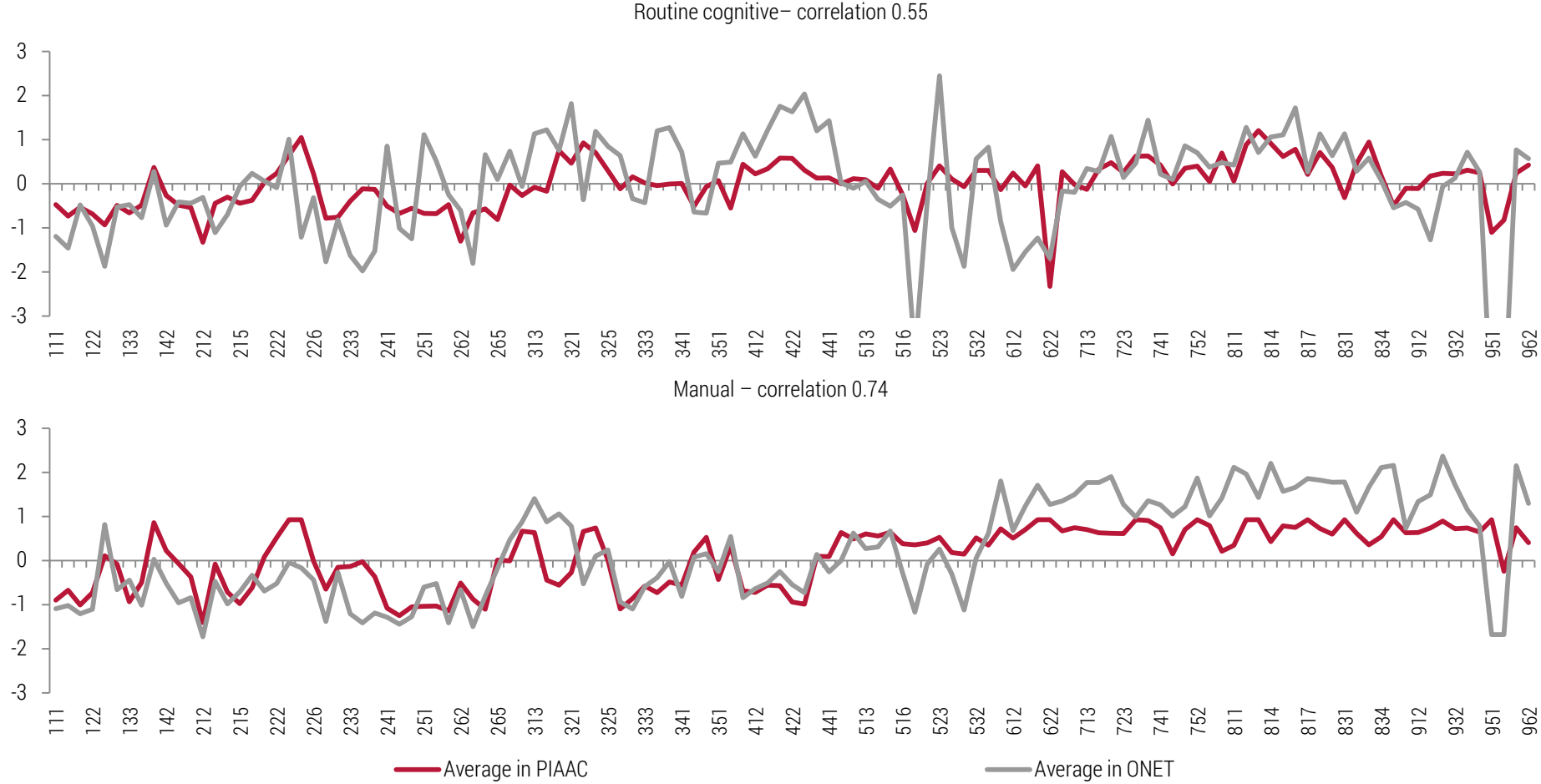
²⁸ The literacy skills tests in STEP and PIAAC follow the same methodology and are comparable.

Figure B1. Values of task contents across 3-digit ISCO occupations in the United States.



Note: The horizontal axis shows selected 3-digit ISCO occupation codes.
 Source: Own calculations using O*NET and PIAAC data.

Figure B1. Values of task contents across 3-digit ISCO occupations in the United States (cont'd).



Note: The horizontal axis shows selected 3-digit ISCO occupation codes. In order to use the same range for all tasks, the negative outliers in the O*NET routine cognitive tasks are truncated at -3: occupation 521 (Street and Market Salespersons) which has the value of -3.86, and occupation 951 (Street and Related Services Workers) and 952 (Street Vendors, excluding food) which both have the value of -5.29.

Source: Own calculations using O*NET and US PIAAC data.

Appendix C. Relevant task items in PIAAC and STEP surveys

Table B1. The considered task items, their exact wordings and possible answers in PIAAC and STEP surveys.

Task item name	PIAAC		STEP	
	Question	Answers	Question	Answers
	In your job, how often do you usually...	1. Never 2. Less than once a month 3. Less than once a week but at least once a month 4. At least once a week but not every day 5. Every day	As a regular part of this work, do you have to read the following...?	Yes / No
Reading bills	- Read bills, invoices, bank statements or other financial statements?		- Bills or financial statements	
Reading news	- Read articles in newspapers, magazines or newsletters?		- Newspapers or magazines	
Reading professional journals	- Read articles in professional journals or scholarly publications?		- Reports	
Reading manuals	- Read manuals or reference materials?		- Instruction manuals/operating manuals	
Filling forms	- Fill in forms?		As part of this work, do you fill out bills or forms?	
	In your job, how often do you usually...	As above	As a normal part of this work, do you do any of the following...?	As above
Advanced math	- Use more advanced math or statistics such as calculus, complex algebra, trigonometry or use of regression techniques?		- Use more advanced math, such as algebra, geometry, trigonometry, etc.	
Calculating prices	- Calculate prices, costs or budgets?		- Calculate prices or costs	
Calculating fractions	- Use or calculate fractions, decimals or percentages?		- Use or calculate fractions, decimals or percentages	
Programming	In your job, how often do you usually use a programming language to program or write computer code?	As above	Does your work as [OCCUPATION] require the use of software programming?	As above
Making speeches or giving presentations	How often does your job usually involve making speeches or presentations in front of five or more people?	As above	As part of this work, do you have to make formal presentations to clients or colleagues to provide information or persuade them of your point of view?	As above
Solving problems	And how often are you usually confronted with more complex problems that take at least 30 minutes to find a good solution? The 30 minutes only refers to the time needed to THINK of a solution, not the time needed to carry it out.	As above	Some tasks are pretty easy and can be done right away or after getting a little help from others. Other tasks require more thinking to figure out how they should be done. As part of this work as [OCCUPATION], how often do you have to undertake tasks that require at least 30 minutes of thinking (examples: mechanic figuring out a car problem, budgeting for a business, teacher making a lesson plan, restaurant owner creating a new menu/dish for restaurant, dress maker designing a new dress)	1. Never 2. Less than once a month 3. Less than once a week but at least once a month 4. At least once a week but not every day 5. Every day
Physical tasks	How often does your job usually involve working physically for a long period?	As above	Using any number from 1 to 10 where 1 is not at all physically demanding (such as sitting at a desk answering a telephone) and 10 is extremely	1-10

			physically demanding (such as carrying heavy loads, construction worker, etc.), what number would you use to rate how physically demanding your work is?	
Supervising	Do you manage or supervise other employees?	Yes / No	As a normal part of this work do you direct and check the work of other workers (supervise)?	Yes / No
Collaborating	In your job what proportion of your time do you usually spend cooperating or collaborating with co-workers?	1. None of the time 2. Up to a quarter of the time 3. Up to half of the time 4. More than half of the time 5. All the time	As part of this work, how frequently do you spend time co-operating or collaborating with co-workers?	1. Never 2. Less than once a month 3. Less than once a week but at least once a month 4. At least once a week but not every day 5. Every day
Changing order of tasks	The next few questions are about the amount of flexibility you have in deciding how you do your job: To what extent can you choose or change the sequence of your tasks?	1. Not at all 2. Very little 3. To some extent 4. To a high extent 5. To a very high extent	Still thinking of your work as [OCCUPATION] how much freedom do you have to decide how to do your work in your own way, rather than following a fixed procedure or a supervisor's instructions? Use any number from 1 to 10 where 1 is no freedom and 10 is complete freedom.	1-10

Note: the PIAAC questions wordings in this table come from the International Master Questionnaire, available at the OECD website.²⁹ The STEP questions wordings in this table come from the English version of the Armenia STEP Skills Measurement Survey, available at the World Bank's microdata website.³⁰

To ensure comparability between STEP and PIAAC data, we rescale the answers to achieve common answer scales in both surveys. The PIAAC questions typically refer to the frequency of performing a task (five levels ranging from 'never' to 'every day'), while many STEP questions refer to whether the responders normally perform a specific task as part of their job or not. Out of 16 questions we consider, two have five available answers in both PIAAC and STEP, and two have 'Yes/No' answers in both PIAAC and STEP. For these questions, we use original variables. For 10 questions which have five available answers in PIAAC but a 'Yes/No' answer in STEP, we convert PIAAC variables into four variants of dummy variables based on the cutoffs in original answers. For two questions which have five available answers in PIAAC and 10 available answers in STEP, we recode the STEP answers into a 1 to 5 scale (1 and 2 to 1, 3 and 4 to 2,..., 9 and 10 to 5). We also correct the item indicating supervising other workers in the STEP data so that only individuals with co-workers are allowed to supervise others.³¹ In the PIAAC data all of the self-employed responders who had no other workers in their jobs indicated they did not supervise anyone. Since this item has a consistent wording in both surveys, our correction of values in STEP ensures consistency with PIAAC data.

²⁹ See www.oecd.org/skills/piaac/BQ_MASTER.HTM [accessed: 2017-05-02].

³⁰ See microdata.worldbank.org/index.php/catalog/2010 [accessed: 2017-05-04].

³¹ Some respondents in STEP indicated supervising other workers despite declaring that they worked alone. Our change corrects this in cases where respondents indicated any of the following combinations: a) being self-employed with no hired workers, b) being self-employed with no unpaid or paid workers, c) being the only paid worker at the current job or that the total number of people working at the organization equals one (the respondent). This problem is not present in CULS.

Appendix D. Comparison of task measures based on STEP and PIAAC data

Table D1. Comparison of task measures based on STEP and PIAAC data

Our measures; PIAAC and STEP		Dicarlo et al. (2016); STEP		de la Rica and Gortazar (2016); PIAAC		Marcolin et al. (2016); PIAAC; Routine Intensity Index only
Task content	Items	Task content	Items	Task content	Items	Items
Non-routine cognitive analytical	Reading news	Non-routine analytical	No. of types of documents read	Abstract	Read diagrams, maps or schematics	Planning own activities
			Length of longest documents typically read			
	Reading professional journals		Length of longest document typically written			
	Solving problems		Solving problems			
	Programming		Advanced math			
			Any of the basic mathematical tasks			
			Learning new things			
Non-routine cognitive personal	Supervising	Non-routine interpersonal	Supervising	Negotiating with people		
	Making speeches or giving presentations		Making speeches or giving presentations			
			Contact with clients			
	Collaborating					
Routine cognitive	Changing order of tasks (reversed)	Routine and Manual	Changing order of tasks (reversed)	Routine	Changing order of tasks	Changing order of tasks
	Filling forms		Repetitiveness		Change how to do work	Change how to do work
			Making speeches or giving presentations (reversed)		Operate machines or equipment	Change speed of work
	Driving				Change working hours	
	Repair electronic equipment				Learn work-related things from co-workers	
					Learning-by-doing from tasks performed	
Manual	Physical tasks	Physical tasks	Manual	Physical tasks		
Methods: uniform coding in STEP and PIAAC; standardisation (means and standard deviations); averages		Methods: standardisation (means and standard deviations); summation		Methods: principal component analysis		Methods: averages.

Source: own elaboration based on Dicarlo et al. (2016), de la Rica and Gortazar (2016) and Marcolin et al. (2016).

Appendix E. Other data sources

In order to estimate the cross-country regressions we merge the PIAAC, STEP and CULS data with three additional variables: ICT stock per worker, number of robots per worker, and the global value chain participation.

The data on ICT capital stock come from Eden and Gaggl (2015). The data are available at the country level, except seven countries in our sample: Armenia, Cyprus, Georgia, Ghana, Estonia, Laos and Macedonia. The latest year available is 2011.

The data on robots come from the International Federation of Robotics [IFR] (2017). The latest data available are from 2016 but we use the average for 2011-2016 since our survey data cover this period. The IFR data are available for ISIC 4 sectors: A, B, C, D and E (jointly), F and P. We aggregate them to three broad categories: Agriculture, Industry and Services and calculate the number of robots per worker in each country / sector cell. The IFR data are unavailable for eight countries in our sample: Armenia, Bolivia, Cyprus, Georgia, Ghana, Kenya, Laos and Macedonia.

The data on global value chain participation are sourced from the RIGVC UIBE (2016) database. We use the backward linkage-based measure, defined as the foreign value added share in production of final goods and services, and the forward-linkage measure, defined as the domestic value added from production of intermediate exports or domestic factor content in intermediate exports (Wang et al., 2017). We use the variables based on GTAP. The latest year available is 2011. We merge the RIGVC UIBE (2016) data with our data at the country-industry level. As the sector classifications are not fully compatible, we aggregate some of the ISIC 4 categories to broader groups: "E+O+P+Q+U" (water supply; sewerage, waste management and remediation activities; public administration and defence; compulsory social security; education; human health and social work activities; activities of extraterritorial organizations and bodies). In China (CULS) this group also includes category D (electricity, gas, steam and air conditioning supply). "G+I" (wholesale and retail trade, repair of motor vehicles and motorcycles; accommodation and food service activities); "L+M+N" (real estate activities; professional, scientific and technical activities; administrative and support service activities); and "R+S+T" (arts, entertainment and recreation; other service activities; activities of households as employers; undifferentiated goods- and services-producing activities of households for own use). The RIGVC UIBE (2016) data are not available for Macedonia.

Appendix F. Offshorability measures based on Blinder and Krueger (2013)

In order to distinguish between offshorable and non-offshorable occupations, we use the Blinder and Krueger (2013) classification based on professional coders' assessments in the PDII survey. We classify as offshorable the occupations that according to the PDII are "offshorable, though with some difficulties or loss of quality that can be overcome" (offshorability score 4 out of 5) or that are "easily offshorable with only minor or no difficulties or loss of quality" (offshorability score 5 out of 5). We use the official ILO crosswalk to map the SOC codes (used in PDII) into ISCO codes. When aggregating to broader occupational categories, we assume that an occupation is offshorable if at least 50% of jobs in this occupation have offshorability score of 4 or 5. As a result, 72 out of 342 occupations at the 4-digit level, 22 out of 120 occupations at the 3-digit level, 5 out of 39 occupations at the 2-digit level, and 1 out of 9 occupations at the 1-digit level are classified as offshorable. The share of offshorable jobs ranges from 6% in Canada to 26% in the Czech Republic. In the majority of countries, the routine task intensity is higher in the offshorable occupation (Table F1).

Table F1. The average routine task intensity and employment shares of offshorable and non-offshorable occupations

	Offshorable occupations		Non-offshorable occupations	
	RTI	Employment share	RTI	Employment share
Kenya	0.01	8.8	0.37	91.2
Ghana	0.08	7.3	0.63	92.7
Ukraine	1.63	9.2	1.16	85.5
Lao PDR	0.79	12.6	0.75	87.4
Bolivia	0.03	10.8	0.35	89.2
Indonesia	0.31	11.1	0.84	88.9
Armenia	0.33	15.2	0.44	84.8
Georgia	0.44	13.8	0.52	86.2
Macedonia, FYR	0.58	20.7	0.24	79.3
Colombia	0.11	13.5	0.24	86.5
China	1.13	10.7	0.72	89.3
Russian Federation	0.45	12.2	0.32	86.8
Turkey	0.50	12.3	0.37	86.2
Poland	0.46	17.2	0.21	82.7
Chile	0.03	13.7	0.20	86.2
Lithuania	0.61	16.0	0.60	81.9
Slovakia	0.38	21.4	0.22	78.0
Cyprus	0.09	12.8	0.21	86.3
Estonia	0.94	11.8	-0.04	88.2
Greece	0.25	10.6	0.41	89.4
Czech Republic	0.41	25.8	0.22	73.9
Slovenia	0.53	23.6	0.17	76.4
Spain	0.32	12.1	0.27	87.5
Korea, Rep.	0.23	19.4	0.24	80.6
Italy	0.46	20.3	0.40	77.9
France	0.24	16.0	0.14	84.0
Israel	0.08	14.6	0.25	84.2
Japan	0.09	14.9	0.06	85.0
New Zealand	-0.08	15.0	-0.04	84.7
United Kingdom	0.02	14.9	0.15	85.1
Belgium	0.27	16.1	0.19	83.9
Germany	0.23	16.5	0.06	82.7
Canada	0.87	6.3	0.05	93.7
Finland	0.44	8.2	-0.30	91.8
Austria	0.86	6.4	-0.06	93.6
Netherlands	0.09	14.1	0.14	85.9
Sweden	-0.54	9.9	-0.61	90.1
Denmark	-0.47	18.0	-0.21	81.7
Singapore	0.28	13.3	-0.03	86.7
United States	-0.15	12.6	0.02	86.7
Ireland	0.27	9.1	0.31	90.9
Norway	-0.08	8.5	-0.20	91.5

Source: own calculations based on PIAAC, STEP, CULS, and PDII data.

Appendix G. Additional regression results

Table G1. The estimated interaction terms between sector fixed effects and GDP per capita (log, demeaned), benchmark specification as in Table 4

	All workers	High-skilled occupations (ISCO 1-3)	Middle-skilled occupations (ISCO 4-5)	Low-skilled occupations (ISCO 7-9)
Agriculture [A]	0.054 (0.077)	-0.077 (0.102)	-0.133 (0.144)	0.057 (0.095)
Mining [B]	-0.083 (0.089)	-0.118 (0.075)	-0.230* (0.133)	-0.081 (0.113)
Manufacturing [C]	0.083 (0.060)	-0.010 (0.066)	-0.111 (0.073)	-0.018 (0.068)
Electricity & Water supply [D+E]	0.103* (0.062)	0.107 (0.074)	-0.110 (0.133)	0.177*** (0.064)
Construction [F]	-0.068 (0.059)	-0.201*** (0.063)	-0.129 (0.080)	-0.141** (0.068)
Transportation and storage [H]	0.224*** (0.061)	-0.038 (0.080)	0.088 (0.084)	0.131* (0.071)
Accommodation and food service [I]	0.022 (0.065)	-0.177** (0.079)	0.040 (0.083)	0.131* (0.079)
Information and communication [J]	-0.293*** (0.095)	-0.080 (0.088)	-0.115 (0.133)	0.010 (0.112)
Financial and insurance [K]	-0.060 (0.097)	0.176** (0.083)	-0.100 (0.138)	0.389*** (0.148)
Real estate & Professional [L]	-0.029 (0.075)	0.154 (0.110)	0.032 (0.100)	0.016 (0.118)
Administrative [M+N]	-0.029 (0.064)	0.002 (0.066)	0.005 (0.078)	0.220*** (0.072)
Public administration [O]	-0.061 (0.080)	0.088 (0.068)	-0.103 (0.111)	0.163* (0.094)
Education [P]	-0.262*** (0.076)	-0.066 (0.063)	-0.082 (0.117)	0.213** (0.086)
Human health [Q]	0.030 (0.069)	0.235*** (0.062)	0.042 (0.076)	0.340*** (0.091)
Arts [R]	-0.233*** (0.070)	-0.120* (0.068)	-0.023 (0.072)	-0.055 (0.113)
Other service [S]	-0.220*** (0.068)	-0.188** (0.074)	-0.259*** (0.077)	-0.011 (0.077)
Activities of household [T]	0.121 (0.111)	-0.646** (0.328)	0.036 (0.158)	0.053 (0.141)
Extraterritorial organizations [U]	-0.026 (0.134)	0.035 (0.144)	-0.322 (0.213)	0.566* (0.339)
No. of observations	148,569	62,907	47,373	38,289
R-squared	0.222	0.126	0.091	0.087

Note: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$. Standard errors in parentheses. We use standardized weights that give each country equal weight.. Source: own estimations based on PIAAC, STEP, CULS World Bank, and RIGVC UIBE (2016) data..

Table G2. The correlates of routine task intensity (RTI) at the worker level, including forward linkage-based measure of participation in global value chains (domestic value added from production of intermediate exports or domestic factor content in intermediate exports), OLS

	All workers	High-skilled occupations (ISCO 1-3)	Middle-skilled occupations (ISCO 4-5)	Low-skilled occupations (ISCO 7-9)
Computer use	-0.532*** (0.193)	-0.702*** (0.169)	-0.381 (0.307)	-0.271 (0.211)
Global Value Chain (GVC) Participation (forward linkage-based)	0.282** (0.136)	-0.038 (0.132)	0.182 (0.171)	0.827*** (0.169)
Ln(GDP per capita) – Mean(Ln(GDP per capita))	0.071 (0.070)	-0.029 (0.064)	0.025 (0.095)	0.087 (0.073)
GVC participation * [Ln(GDP pc) – mean(Ln(GDP pc))]	-0.592*** (0.167)	-0.330** (0.140)	-0.352 (0.216)	-0.638*** (0.198)
FDI / GDP	0.011* (0.006)	0.024*** (0.007)	0.011 (0.008)	-0.014** (0.007)
Education: primary	0.247*** (0.019)	0.137*** (0.029)	0.221*** (0.022)	0.139*** (0.026)
Education: tertiary	-0.486*** (0.017)	-0.267*** (0.018)	-0.201*** (0.021)	-0.143*** (0.034)
Literacy skills level: 1 or lower	0.085*** (0.018)	0.036 (0.028)	0.054** (0.026)	0.069*** (0.025)
Literacy skills level: 3	-0.139*** (0.013)	-0.087*** (0.016)	-0.062*** (0.022)	-0.049** (0.023)
Literacy skills level: 4 and 5	-0.292*** (0.019)	-0.192*** (0.020)	-0.065** (0.029)	-0.170*** (0.043)
Female	0.248*** (0.012)	0.237*** (0.013)	0.203*** (0.018)	0.345*** (0.025)
Age: 16-24	0.227*** (0.017)	0.220*** (0.027)	0.207*** (0.025)	0.147*** (0.024)
Age: 35-44	-0.055*** (0.011)	-0.063*** (0.014)	-0.022 (0.018)	-0.039* (0.021)
Age: 45-54	-0.015 (0.014)	-0.064*** (0.016)	0.014 (0.022)	0.039* (0.022)
Age: 55-65	0.016 (0.016)	-0.054*** (0.019)	0.107*** (0.026)	0.075*** (0.025)
Sector fixed effects	Yes	Yes	Yes	Yes
No. of observations	148,569	62,907	47,373	38,289
R-squared	0.222	0.126	0.091	0.087

Note: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$. Standard errors in parentheses. We use standardized weights that give each country equal weight. The reference levels are: age 25-34, secondary education, wholesale and retail trade; repair of motor vehicles and motorcycles (ISIC G), lower medium literacy skills (level 2).

Source: own estimations based on PIAAC, STEP, CULS World Bank, and RIGVC UIBE (2016) data.

Table G3. The correlates of RTI, including robots per worker and ICT stock, 31 countries with available data, OLS regressions

	Model with robots and ICT capital as controls (31 countries)				Benchmark specification for 31 countries with robot and ICT data available			
	All workers	High-skilled occupations (ISCO 1-3)	Middle- skilled occupations (ISCO 4-5)	Low-skilled occupations (ISCO 7-9)	All workers	High-skilled occupations (ISCO 1-3)	Middle- skilled occupations (ISCO 4-5)	Low-skilled occupations (ISCO 7-9)
Computer use	-0.512*** (0.172)	-0.616*** (0.163)	-0.401 (0.290)	-0.331* (0.171)	-0.509*** (0.175)	-0.635*** (0.164)	-0.389 (0.308)	-0.358** (0.166)
ICT stock per worker	-0.003 (0.018)	-0.010 (0.018)	0.016 (0.023)	-0.014 (0.026)				
Robots per worker	0.009 (0.011)	-0.008 (0.013)	-0.028* (0.016)	-0.003 (0.011)				
FVA share	0.824*** (0.234)	0.454** (0.226)	0.888*** (0.300)	1.251*** (0.243)	0.827*** (0.232)	0.462** (0.225)	0.864*** (0.300)	1.262*** (0.238)
Ln(GDP per capita) – mean(Ln(GDP per capita))	0.061 (0.099)	0.013 (0.059)	0.034 (0.122)	0.072 (0.112)	0.058 (0.094)	0.005 (0.055)	0.043 (0.118)	0.064 (0.108)
FVA share * [Ln(GDP pc) – mean(Ln(GDP pc)]	-1.361*** (0.256)	-1.422*** (0.250)	-1.308*** (0.357)	-1.082*** (0.248)	-1.363*** (0.256)	-1.415*** (0.248)	-1.283*** (0.356)	-1.078*** (0.248)
FDI / GDP	0.100*** (0.019)	0.131*** (0.017)	0.132*** (0.031)	0.060*** (0.020)	0.100*** (0.019)	0.132*** (0.017)	0.131*** (0.030)	0.062*** (0.021)
Education: primary	0.297*** (0.017)	0.160*** (0.029)	0.251*** (0.021)	0.164*** (0.020)	0.297*** (0.017)	0.159*** (0.029)	0.253*** (0.021)	0.164*** (0.020)
Education: tertiary	-0.499*** (0.016)	-0.281*** (0.018)	-0.198*** (0.021)	-0.167*** (0.035)	-0.499*** (0.016)	-0.281*** (0.018)	-0.201*** (0.021)	-0.167*** (0.036)
Literacy skills level: 1 or lower	0.136*** (0.018)	0.075** (0.033)	0.072** (0.028)	0.110*** (0.025)	0.136*** (0.018)	0.075** (0.033)	0.072** (0.028)	0.110*** (0.025)
Literacy skills level: 3	-0.131*** (0.013)	-0.091*** (0.017)	-0.046** (0.023)	-0.051** (0.025)	-0.131*** (0.013)	-0.092*** (0.017)	-0.046** (0.023)	-0.051** (0.025)
Literacy skills level: 4 and 5	-0.258*** (0.018)	-0.179*** (0.020)	-0.035 (0.030)	-0.157*** (0.041)	-0.258*** (0.018)	-0.180*** (0.021)	-0.035 (0.030)	-0.157*** (0.041)
Female	0.272*** (0.014)	0.233*** (0.014)	0.250*** (0.020)	0.381*** (0.029)	0.272*** (0.014)	0.233*** (0.014)	0.252*** (0.020)	0.382*** (0.028)
Age: 16-24	0.277*** (0.018)	0.263*** (0.029)	0.268*** (0.027)	0.161*** (0.024)	0.277*** (0.018)	0.263*** (0.029)	0.269*** (0.027)	0.160*** (0.024)
Age: 35-44	-0.085*** (0.012)	-0.079*** (0.014)	-0.051** (0.022)	-0.066*** (0.021)	-0.085*** (0.012)	-0.079*** (0.014)	-0.052** (0.022)	-0.065*** (0.021)
Age: 45-54	-0.068*** (0.014)	-0.084*** (0.017)	-0.025 (0.024)	-0.014 (0.022)	-0.068*** (0.014)	-0.085*** (0.017)	-0.025 (0.024)	-0.014 (0.022)
Age: 55-65	-0.045*** (0.016)	-0.089*** (0.019)	0.077*** (0.029)	0.013 (0.024)	-0.045*** (0.016)	-0.089*** (0.019)	0.076*** (0.029)	0.013 (0.024)

Sector fixed effects	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
No. of obs.	121,111	53,995	35,713	31,403	121 111	53 995	35 713	31 403
R-squared	0.243	0.143	0.104	0.119	0.239	0.136	0.1	0.112

Note: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$. Standard errors in parentheses. We use standardized weights that give each country equal weight. The reference levels are: age 25-34, secondary education, middle-skilled occupations (ISCO 4-5), wholesale and retail trade; repair of motor vehicles and motorcycles (ISIC G), lower medium literacy skills (level 2). The coefficients for sector fixed effects are not presented in order to save space, and are available on request. ICT stock per worker, robots per worker, FDI and the FVA share in domestic production variables are standardized in our sample.

Source: own estimations based on PIAAC, STEP, CULS, World Bank, Eden and Gaggl (2015) and IFR and RIGVC UIBE (2016) data.

Table G4. The correlates of routine task intensity (RTI) at the sector level, OLS

All workers					
Computer use	-1.302*** (0.185)	Age: 16-24	-0.208 (0.247)	Sector J	-0.378*** (0.076)
Foreign Value Added share	0.153 (0.167)	Age: 35-44	0.195 (0.203)	Sector K	-0.228*** (0.070)
Ln(GDP per capita) – mean(Ln(GDP per capita))	0.025 (0.067)	Age: 45-54	0.415 (0.302)	Sector L	-0.153*** (0.048)
FVA share * [Ln(GDP pc) –mean(Ln(GDP pc)]	-0.140 (0.113)	Age: 55-65	-0.628** (0.311)	Sector M+N	-0.151*** (0.038)
FDI / GDP	-0.008 (0.008)	Sector A	-0.046 (0.069)	Sector O	-0.203*** (0.068)
Education: primary	0.008 (0.191)	Sector B	-0.246** (0.123)	Sector P	-0.537*** (0.067)
Education: tertiary	0.089 (0.161)	Sector C	-0.081 (0.054)	Sector Q	-0.185*** (0.063)
Literacy skills level: 1 or lower	-0.422* (0.245)	Sector D+E	-0.070 (0.094)	Sector R	-0.314*** (0.048)
Literacy skills level: 3	0.107 (0.275)	Sector F	-0.260*** (0.090)	Sector S	-0.259*** (0.054)
Literacy skills level: 4 and 5	-0.940*** (0.345)	Sector H	0.070 (0.071)	Sector T	0.082 (0.130)
Female	0.192 (0.182)	Sector I	-0.017 (0.058)	Sector U	-0.213 (0.160)
No. of observations	747				
R-squared	0.775				

Note: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$. Standard errors in parentheses. All variables are calculated averages in sector s in country c . We use standardised weights that give each country equal weight. The reference levels are: age 25-34, secondary education, wholesale and retail trade; repair of motor vehicles and motorcycles (ISIC G), lower medium literacy skills (level 2). RTI, FDI and the FVA share in domestic production variables are standardized in our sample.

Source: own estimations based on PIAAC, STEP, CULS World Bank, and RIGVC UIBE (2016) data.

Table G5. The correlates of routine task intensity (RTI) among workers in offshorable and non-offshorable occupations

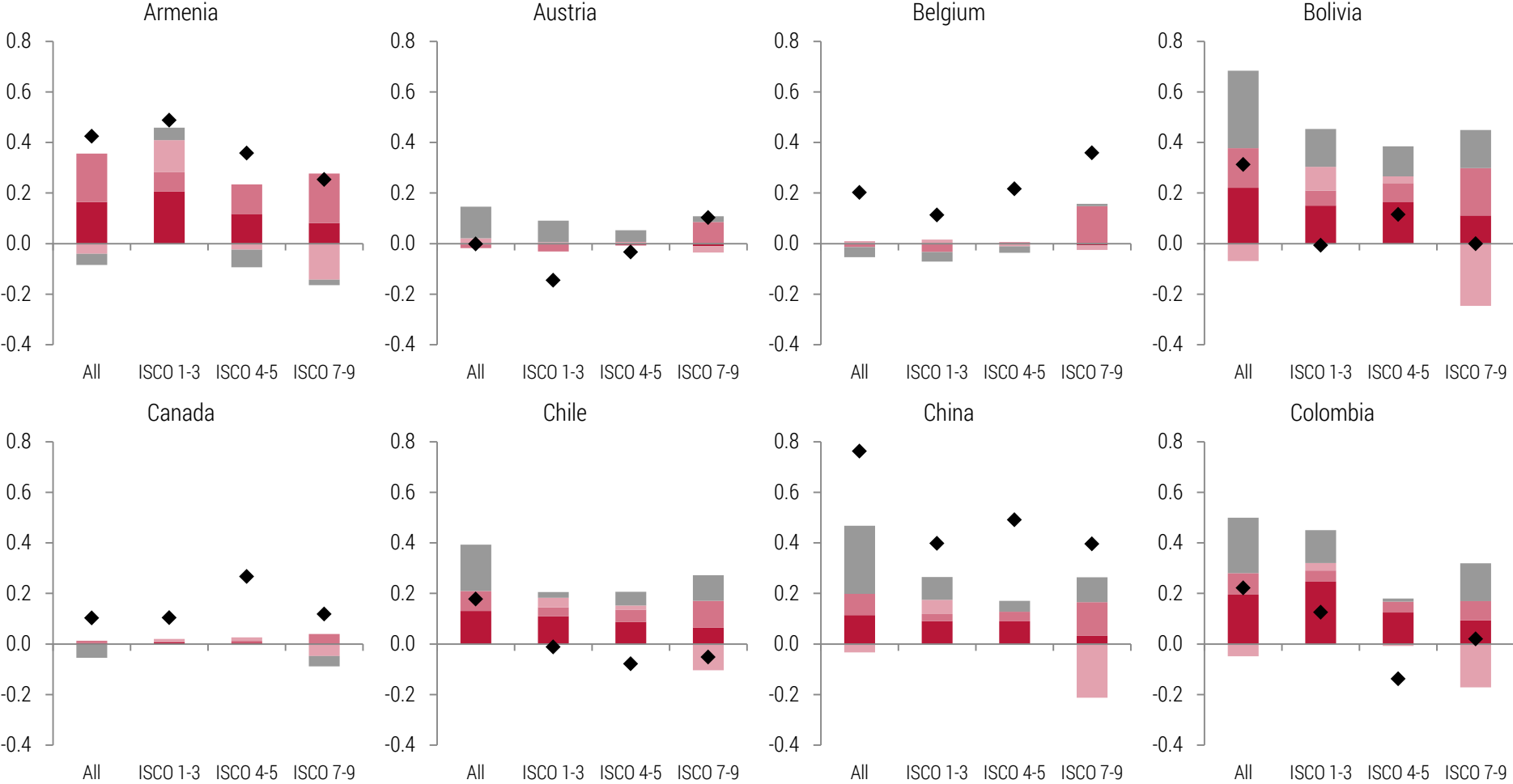
	All workers	Workers in non-offshorable occupations	Workers in offshorable occupations
Computer use	-0.508** (0.204)	-0.555*** (0.204)	-0.012 (0.309)
Foreign Value Added (FVA) share	0.269* (0.151)	0.171 (0.147)	0.762*** (0.236)
Ln(GDP per capita) – mean(Ln(GDP per capita))	0.060 (0.079)	0.062 (0.080)	0.015 (0.101)
FVA share * [Ln(GDP pc) – mean(Ln(GDP pc))]	-0.424** (0.197)	-0.396** (0.185)	-0.530* (0.300)
FDI / GDP	0.009* (0.005)	0.012** (0.005)	-0.006 (0.007)
Education: primary	0.246*** (0.020)	0.246*** (0.019)	0.227*** (0.036)
Education: tertiary	-0.486*** (0.017)	-0.488*** (0.018)	-0.396*** (0.038)
Literacy skills level: 1 or lower	0.078*** (0.018)	0.071*** (0.019)	0.117*** (0.041)
Literacy skills level: 3	-0.139*** (0.014)	-0.139*** (0.015)	-0.119*** (0.033)
Literacy skills level: 4 and 5	-0.293*** (0.019)	-0.285*** (0.021)	-0.308*** (0.044)
Female	0.250*** (0.012)	0.242*** (0.012)	0.305*** (0.027)
Age: 16-24	0.227*** (0.017)	0.231*** (0.018)	0.195*** (0.041)
Age: 35-44	-0.055*** (0.011)	-0.053*** (0.012)	-0.062** (0.027)
Age: 45-54	-0.012 (0.014)	-0.010 (0.015)	-0.017 (0.026)
Age: 55-65	0.018 (0.016)	0.012 (0.017)	0.081** (0.035)
Sector fixed effects	Yes	Yes	Yes
No. of observations	148,120	129,965	18,155
R-Squared	0.220	0.222	0.245

Note: *** p<0.01, ** p<0.05, * p<0.1. Standard errors in parentheses. We use standardized weights that give each country equal weight. The reference levels are: age 25-34, secondary education, wholesale and retail trade; repair of motor vehicles and motorcycles (ISIC G), lower medium literacy skills (level 2). RTI, FDI and the FVA share in domestic production variables are standardized in our sample.

Source: own estimations based on PIAAC, STEP, CULS World Bank, and RIGVC UIBE (2016) data.

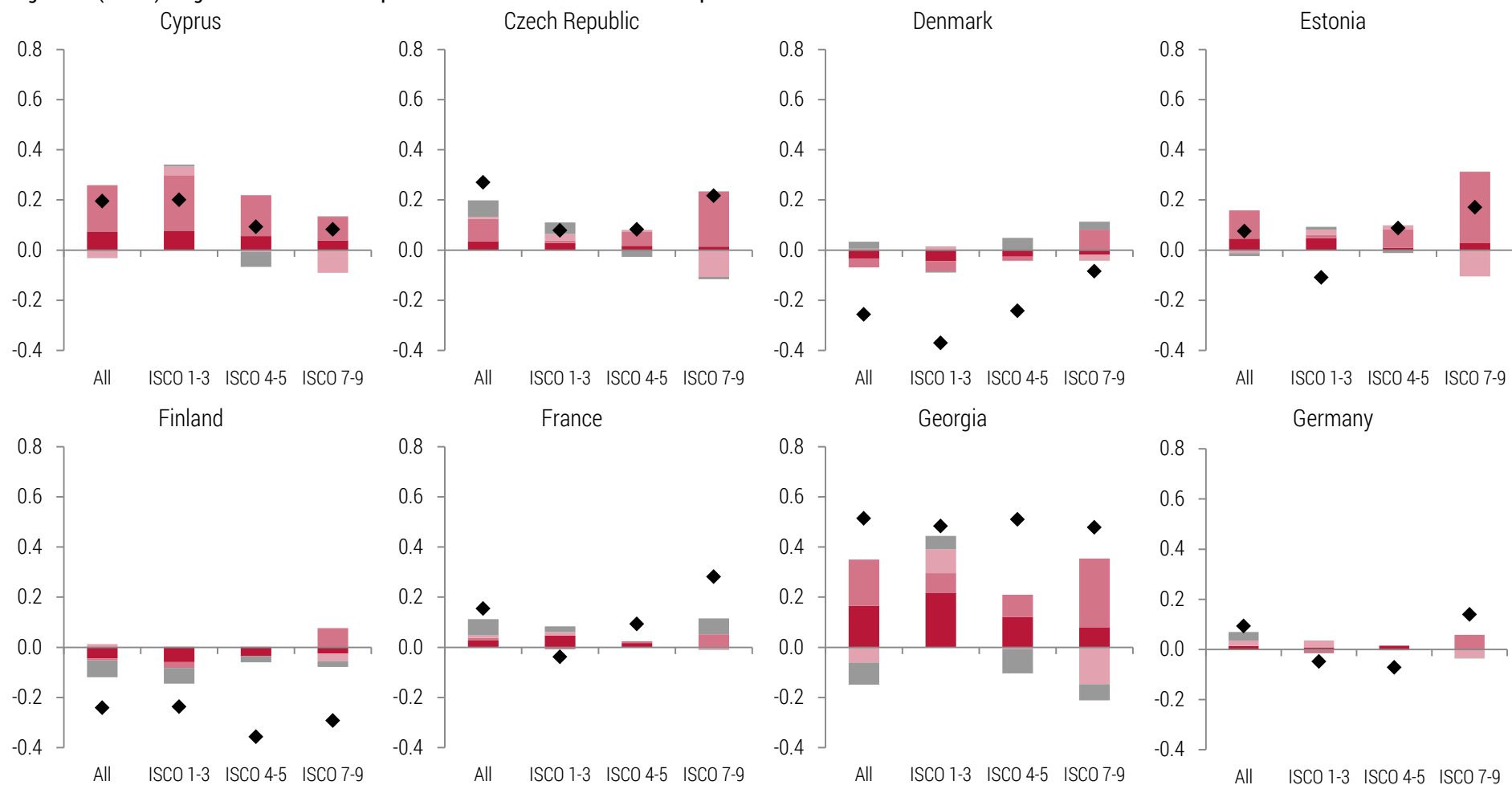
Appendix H. Decomposition results for all countries

Figure H1. Regression-based decomposition of differences in RTI between particular countries and the US.



Note: Results of decomposition (5) based on regressions presented in Table 4. 0 is set at the US average value and 1 corresponds to one standard deviation of RTI in the US. Source: own estimations based on PIAAC, STEP, CULS, World Bank and RIGVC UIBE (2016).

Figure H1 (cont'd). Regression-based decomposition of differences in RTI between particular countries and the US.



Note: Results of decomposition (5) based on regressions presented in Table 4. 0 is set at the US average value and 1 corresponds to one standard deviation of RTI in the US. Source: own estimations based on PIAAC, STEP, CULS, World Bank and RIGVC UIBE (2016).

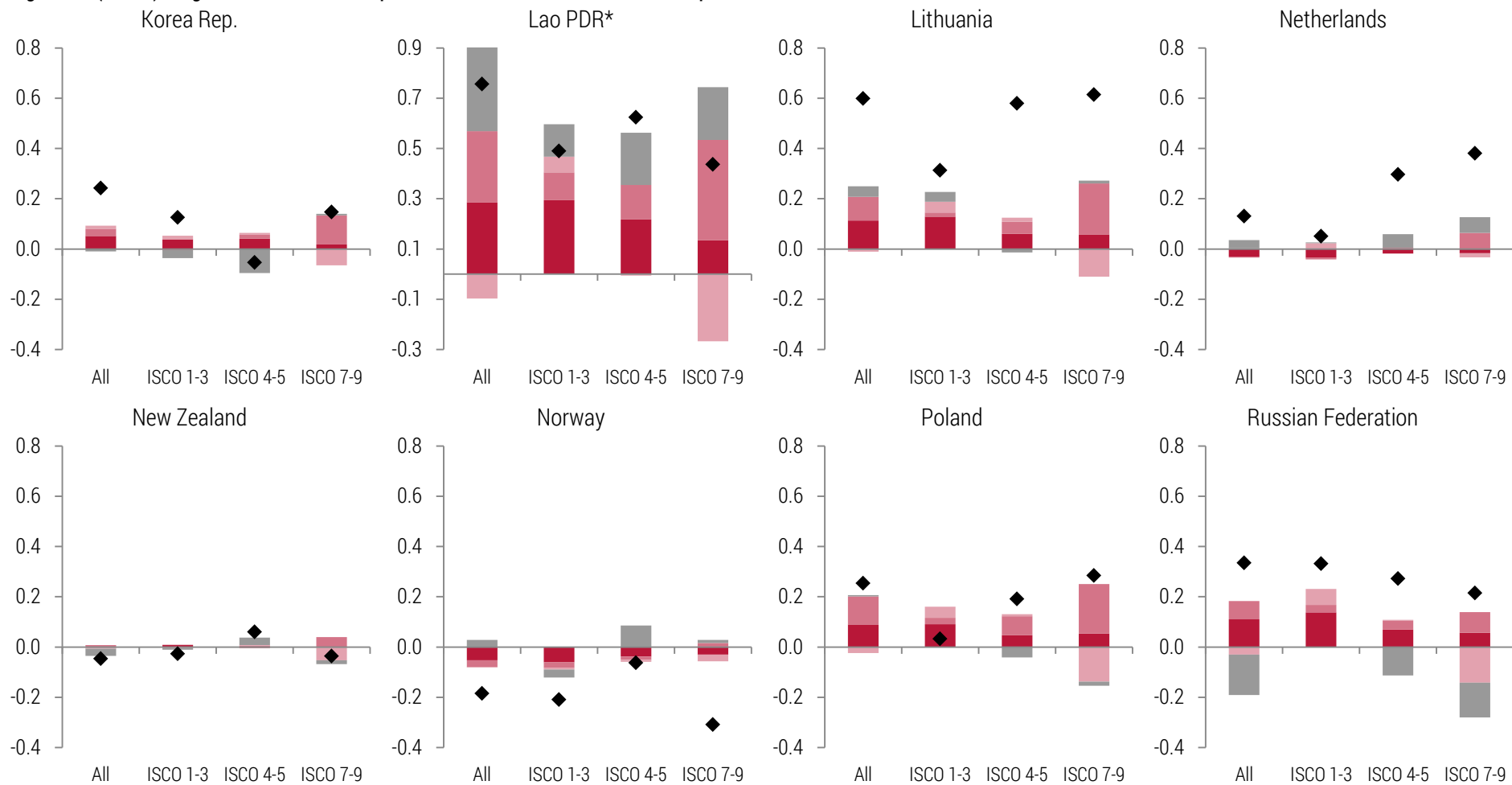
Figure H1 (cont'd). Regression-based decomposition of differences in RTI between particular countries and the US.



Note: Results of decomposition (5) based on regressions presented in Table 4. 0 is set at the US average value and 1 corresponds to one standard deviation of RTI in the US. * Figures for countries marked with asterisk have different vertical scale than most countries.

Source: own estimations based on PIAAC, STEP, CULS, World Bank and RIGVC UIBE (2016).

Figure H1 (cont'd). Regression-based decomposition of differences in RTI between particular countries and the US.



Note: Results of decomposition (5) based on regressions presented in Table 4. 0 is set at the US average value and 1 corresponds to one standard deviation of RTI in the US. * Figures for countries marked with asterisk have different vertical scale than most countries.

Source: own estimations based on PIAAC, STEP, CULS, World Bank and RIGVC UIBE (2016).

Figure H1 (cont'd). Regression-based decomposition of differences in RTI between particular countries and the US.



Note: Results of decomposition (5) based on regressions presented in Table 4. 0 is set at the US average value and 1 corresponds to one standard deviation of RTI in the US. * Figures for countries marked with asterisk have different vertical scale than most countries.

Source: own estimations based on PIAAC, STEP, CULS, World Bank and RIGVC UIBE (2016).



www.ibs.org.pl